

1 **TITLE**

2 **Cerebral microbleeds and stroke risk after ischaemic stroke or transient ischaemic**
3 **attack: a pooled analysis of individual patient data from cohort studies**

4

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1 **SUMMARY**

2 **Background** Cerebral microbleeds (CMBs) are a neuroimaging biomarker ('biological
3 marker') of stroke risk. A crucial clinical question is whether CMBs identify patients with
4 recent ischaemic stroke or TIA in whom the risk of future intracranial haemorrhage exceeds
5 that of recurrent ischaemic stroke when treated with antithrombotic drugs. We therefore aimed
6 to determine whether a large burden of CMBs, or particular anatomical patterns, can identify
7 ischaemic stroke or TIA patients at higher absolute risk of intracranial haemorrhage than
8 ischaemic stroke.

9 **Methods** In a pooled analysis of individual patient data from cohort studies in adults with
10 recent ischaemic stroke or TIA, we determined associations of baseline CMBs with stroke risks
11 using multivariable Cox regression, adjusted for pre-specified prognostic variables. We
12 registered this study (CRD42016036602).

13 **Findings** Among 20,322 patients from 38 cohorts (over 35,225 patient-years [median 1.98
14 years] follow-up), CMBs were associated with the composite risk of intracranial haemorrhage
15 and ischaemic stroke (adjusted hazard ratio[aHR], 95%CI 1.39, 1.24-1.56), intracranial
16 haemorrhage (2.56, 1.89-3.47), and ischaemic stroke (1.27, 1.10-1.45). As CMB burden
17 increased, there was a greater increase in the aHR for intracranial haemorrhage than for
18 ischaemic stroke (for ≥ 5 CMBs, aHR 4.55[3.08-6.72] vs. 1.47[1.19-1.80]; for ≥ 10 CMBs, aHR
19 5.52[3.36-6.72] vs. 1.43[1.07-1.91]; for ≥ 20 CMBs, aHR 8.61[4.69-15.81] vs. 1.86[1.23-1.82],
20 respectively). Even with high CMB burden, the absolute rate of ischaemic stroke exceeded that
21 of intracranial haemorrhage (for ≥ 10 CMBs, rate 64[48-84] per 1000 patient-years vs. 27[17-
22 41] per 1000 patient-years, respectively; for ≥ 20 CMBs, rate 73[46-108] per 1000 patient-years
23 vs. 39[21-67] per 1000 patient-years, respectively). Hazard ratios for all outcomes were similar

1 regardless of CMB anatomical distribution. Similar patterns were seen in patients taking
2 antiplatelet drugs or oral anticoagulants.

3 **Interpretation** In patients with recent ischaemic stroke or TIA, CMBs are associated with a
4 greater aHR for intracranial haemorrhage than for ischaemic stroke, but the absolute risk of
5 ischaemic stroke is consistently higher than that of intracranial haemorrhage; CMB presence,
6 burden or distribution do not identify patients at higher absolute risk of intracranial
7 haemorrhage than recurrent ischaemic stroke.

8

9 **Funding**

10 British Heart Foundation, UK Stroke Association.

11

1 **RESEARCH IN CONTEXT**

2 **Evidence before this study**

3 We searched Medline and EMBASE (search strategy: cerebral adj2 micro* OR CMB OR
4 microbleed.mp AND [stroke.mp OR stroke/ OR intracerebral h?emorr* OR intracranial
5 h?emorr* OR isch?emic stroke OR isch?emic infarct*]). An aggregate level meta-analysis
6 (n=5068) showed that CMBs were associated with both intracranial haemorrhage (risk ratio
7 [RR] 3.8, 95% CI 3.5-11.4) and ischaemic stroke (RR 1.8, 95% CI 1.4-2.5); this pooled
8 analysis, and another recent study in two cohorts (one mainly Chinese [n=1003], the other
9 mainly white Caucasian [n=1080]) reported that five or more CMBs were associated with
10 similar absolute risks of intracranial haemorrhage and ischaemic stroke. However, limited
11 sample sizes and intracranial haemorrhage outcome events in previous studies did not provide
12 enough statistical power and precision to determine whether a large CMB burden or
13 distribution pattern is associated with a higher absolute risk of intracranial haemorrhage than
14 ischaemic stroke in patients with recent ischaemic stroke or TIA treated with antithrombotic
15 drugs.

16 **Added value of this study**

17 Our pooled analysis of individual data from 20,322 patients with recent ischaemic stroke or
18 TIA confirms that CMBs are associated with a higher adjusted hazard ratio (aHR) for
19 intracranial haemorrhage than ischaemic stroke; with increasing CMB burden the aHR for
20 intracranial haemorrhage rises more steeply than that of ischaemic stroke. Regardless of the
21 number or distribution of CMBs, or the type of antithrombotic treatment received (antiplatelet
22 drugs or oral anticoagulants), the absolute rate of ischaemic stroke substantially exceeded that
23 of intracranial haemorrhage.

24

1 **Implications of all the available evidence**

2 Although CMBs can inform hazard for intracranial haemorrhage in patients with recent
3 ischaemic stroke or TIA treated with antithrombotic drugs, the absolute risk of ischaemic stroke
4 is consistently much higher than that of intracranial haemorrhage, regardless of CMB presence,
5 burden, or pattern.

1 **INTRODUCTION**

2

3 A central challenge in stroke prevention after ischaemic stroke or TIA is to predict the risk of
4 intracranial haemorrhage, and to differentiate this from the risk of recurrent ischaemic stroke,
5 in patients treated with antithrombotic therapy - usually antiplatelet agents or, in patients with
6 atrial fibrillation, oral anticoagulants (OAC)¹. Cerebral microbleeds (CMBs) are a radiological
7 finding of small (<10mm), hypointense (black) ovoid or rounded regions on T2*-weighted
8 gradient-recalled echo (GRE) or susceptibility-weighted imaging (SWI)². CMBs mostly
9 correspond pathologically to haemosiderin-laden macrophages close to arterioles affected by
10 small vessel diseases^{3,4}; strictly lobar CMBs suggest cerebral amyloid angiopathy (CAA),
11 while deep or mixed patterns probably indicate arteriolosclerosis or mixed pathologies,
12 respectively⁵⁻⁸. It is hypothesised that CMBs result from red blood cell leakage from arterioles
13 and capillaries, raising clinical concerns that they herald an increased risk of potentially
14 devastating intracranial haemorrhage, particularly in patients treated with antithrombotic
15 drugs⁹. However, CMBs signal small vessel diseases that can also cause ischaemic stroke, and
16 might result from non-haemorrhagic mechanisms¹⁰⁻¹³. In ischaemic stroke cohorts, CMBs are
17 associated with the risks of both subsequent intracranial haemorrhage and recurrent ischemic
18 stroke¹⁴⁻²⁸. As the number of CMBs increases, the risk of intracranial haemorrhage seems to
19 rise more steeply than that of ischaemic stroke, and ≥ 5 CMBs have been reported to be
20 associated with similar absolute risks of intracranial haemorrhage and ischaemic stroke^{28,29}.

21 Because previous studies had limited sample sizes and few intracranial haemorrhage
22 outcome events, they could not reliably answer the crucial clinical question of whether many
23 CMBs, or particular CMB patterns, predict a higher risk of intracranial haemorrhage than
24 recurrent ischaemic stroke. We established the Microbleeds International Collaborative
25 Network³⁰ to undertake large-scaled pooled analyses of prospective observational cohort

1 studies. We tested the hypothesis that a large burden of CMBs, or particular anatomical
2 patterns, can identify ischaemic stroke or TIA patients at higher absolute risk of intracranial
3 haemorrhage than ischaemic stroke.

4 5 **METHODS**

6
7 The Microbleeds International Collaborative Network protocol and statistical analysis plan
8 were registered with PROSPERO on 5th April 2016 (36602)
9 (https://www.crd.york.ac.uk/PROSPERO/display_record.php?RecordID=36602).

10 11 **Study design and participants**

12
13 We identified cohorts by searching Medline and EMBASE (search strategy: cerebral adj2
14 micro* OR CMB OR microbleed.mp AND stroke.mp OR stroke/ OR intracerebral h?emorr*
15 OR intracranial h?emorr* OR isch?emic stroke OR isch?emic infarct*), clinical trial databases
16 (clinicaltrials.gov, strokecenter.org), and scientific meeting abstracts. We invited members of
17 METACOHORTS³¹; an international database of over 90 studies in small vessel disease,
18 including 660,000 patients. Two authors (DW and DJW) independently did the search and
19 reviewed all titles and abstracts; DW and DJW also did an independent risk of bias assessment
20 for all included studies. Cohorts were eligible for inclusion if they: prospectively recruited adult
21 participants with ischaemic stroke or TIA; included at least 50 participants; collected data on
22 stroke events over at least three months follow up; used an appropriate MRI sequence sensitive
23 to magnetic susceptibility (GRE or SWI); and documented the number and distribution of
24 CMBs reliably using consensus criteria and validated scales. Each patient was only included
25 in one cohort.

1 **Risk of bias and study quality**

2

3 We assessed all studies for risk of bias (including selection bias) and quality using the Cochrane
4 Collaboration tool³².

5

6 **Ethical approval**

7 All cohorts obtained ethical approval as required by local regulations to allow data sharing. All
8 data reviewed by the co-ordinating centre was fully anonymised. The project was approved by
9 the Health Research Authority United Kingdom (REC reference: 8/HRA/0188).

10

11 **Outcome events**

12

13 Our pre-specified primary outcome events were: a composite of any symptomatic intracranial
14 haemorrhage (confirmed radiologically, including subdural, extradural and subarachnoid
15 haemorrhage, and excluding intracranial haemorrhages attributed to intravenous thrombolysis
16 or trauma) or ischaemic stroke (acute or subacute neurological symptoms lasting >24 h, and
17 attributed to cerebral ischaemia, diagnosed clinically, with or without radiological
18 confirmation); symptomatic intracranial haemorrhage; and symptomatic ischaemic stroke.
19 Secondary outcome events were death (all cause) and vascular death. All events were
20 adjudicated according to individual cohort protocols.

21

22 **Statistical analysis**

23

24 As per our prespecified protocol, a single dataset was created by combining individual
25 participant data from the 38 cohorts. We compared baseline demographic and risk factor

1 profiles between patients with and without CMBs, and between patients with and without
2 outcome events, using the Mann-Whitney test if not normally distributed, or the t-test if
3 normally distributed; we compared categorical variables between groups with the Chi-squared
4 test or Fisher's exact test. We censored patients at the last available follow-up (truncated to
5 five years) or at the time of the pre-specified outcome. In instances where patient had multiple
6 events of the same type, we censored follow up at the first event. We calculated absolute event
7 rates per 1000 patient-years for primary outcomes in patients with and without CMBs. We
8 assessed the proportional hazards assumption through visual inspection of (log-log) plots of
9 log cumulative hazard against time and tested for a non-zero slope in a regression of scaled
10 Schoenfeld residuals against time. We calculated univariate Kaplan-Meier survival
11 probabilities in patients with and without CMBs to estimate event rates, and used the log-rank
12 test to compare groups. We performed multivariable Cox regression adjusting for the following
13 prognostic and confounding variables (selected by consensus, based upon availability,
14 biological plausibility and known associations with CMBs and outcomes): age; sex;
15 presentation with TIA or ischaemic stroke; history of hypertension; previous stroke; known
16 atrial fibrillation; antithrombotic use after index event; and type of MRI sequence used to detect
17 CMBs [T2*-weighted GRE, or SWI]). We investigated the effect of predefined CMB burden
18 categories (1, 2-4, ≥ 5 , ≥ 10 and ≥ 20). When investigating CMB distribution, we adjusted for
19 total CMB number. We added a shared frailty term³³ to account for patients being nested in
20 individual studies (thus potentially having correlated data). We performed sub-analyses for
21 patients treated with OAC and antiplatelet drugs and added interaction terms between
22 antithrombotic therapy and CMB presence. We categorised ethnicity (when available) as White
23 Caucasian or Asian (Japanese, Chinese, Malays, Indian, Pakistani, Korean) to investigate the
24 interaction between ethnicity and CMB presence. We performed two pre-specified sensitivity
25 analyses: first, exploring time-varying risks within the Cox model to investigate later events

1 (beyond the first year) accounting for death as a competing risk (using the Fine-Gray
2 subdistribution hazard model), calculating subdistribution hazard ratios (SHR); and second, a
3 two-stage individual patient meta-analysis to quantify between-study heterogeneity using the
4 inverse-variance method (which fits a separate survival model for each cohort, then pools and
5 displays estimates in a forest plot). We performed three post-hoc analysis: first adding white
6 matter hyperintensities (WMH; another common marker of cerebral small vessel disease, rated
7 using the Fazekas scale³⁴ and considered severe if ≥ 2 in the periventricular or deep white
8 matter) into our multivariable model; second, including only intracerebral haemorrhage,
9 convexity subarachnoid haemorrhage and subdural haemorrhage, as these bleeding events are
10 the most likely to be associated with CMBs; and third, investigating the interaction between
11 CMB and age, dichotomised as less than or greater than 80 years. If a variable was not
12 sufficiently available across all cohorts for the final multivariable model, it was excluded. In
13 sensitivity analyses, if a variable of interest was not sufficiently available in a cohort, the cohort
14 was excluded. We undertook all statistical analysis using STATA version 15 (StataCorp LP,
15 TX).

16

17 **Role of the funding source**

18 The funder of the study had no role in the study design, data collection, data analysis or data
19 interpretation, or writing of the report. The corresponding author had access to all the data in
20 the study and had final responsibility for the decision to submit for publication.

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1 **RESULTS**

2

3 We included 20,322 participants from 38 cohorts including 23 published studies and 15
4 unpublished studies (study flow chart shown in figure 1; demographics, risk factors and
5 outcome events for each cohort are shown in table 1). Although more than one-half of
6 participants and outcome events came from the six largest cohorts, there was no major risk of
7 bias for any included cohort (appendix table 1). The mean age was 70 years (SD 13); 8593
8 (42%) were female. CMBs were present in 5649 (28%) of patients (characteristics shown in
9 appendix table 2), including 2,145 (12%) with 1 CMB, 1,990 (10%) with 2-4 CMB, and 1,244
10 (6%) with ≥ 5 CMBs. Over 35,225 patient-years of follow-up (median 1.34 years, IQR 0.19 to
11 2.44), there were 1,474 composite events: 189 intracranial haemorrhages; 1,113 ischaemic
12 strokes; and 172 composite events of unknown type from one cohort (total n=3,355
13 participants) that did not subclassify composite outcomes as intracranial haemorrhage or
14 ischaemic stroke. Characteristics between patients with and without events are presented in the
15 appendix tables 2 to 4. Visual assessment of the log-log plots and the results of testing the
16 Schoenfeld residuals suggest that the proportional hazards assumption was not violated in any
17 of the following analyses.

18

19 **Survival estimates according to baseline CMBs**

20

21 The composite of any intracranial haemorrhage or ischaemic stroke ($p < 0.0001$; log-rank test),
22 symptomatic intracranial haemorrhage ($p < 0.0001$) and symptomatic ischaemic stroke
23 ($p < 0.0001$) were more frequent in patients with CMBs compared to those without. Kaplan-
24 Meier curves for CMB presence and burden categories are shown in figure 2, and for different
25 CMB distributions in the appendix figure 1.

1 **Multivariable Cox regression for outcome events according to CMB presence, burden**
2 **and distribution**

3

4 **Composite of ischaemic stroke and intracranial haemorrhage**

5

6 1,461 of 20,332 patients had 1,474 composite outcome events during 35,225 patient-years of
7 follow up. The incidence of all composite events in patients with any CMB was 59 per 1000
8 patient-years (95% CI 54-64) compared to 35 per 1000 patient-years (95% CI 33-38) in those
9 without CMBs, an absolute increased incidence of 24 outcomes per 1000 patient-years. Table
10 2 shows the incidence of a composite event for each CMB burden and distribution category.
11 The risk of a composite event was 1.35 (95% CI 1.20-1.50) times higher for patients with
12 CMBs than those without and rose with increasing CMB burden (table 2), p trend <0.0001).
13 Mixed, deep, lobar, and strictly lobar CMBs were associated with similar hazards (table 2).

14

15 **Intracranial haemorrhage**

16

17 189 patients had a symptomatic intracranial haemorrhage during 32,847 patient-years of follow
18 up (151 intracerebral haemorrhages, 31 subdural haemorrhages, 8 subarachnoid haemorrhages
19 [4 cortical], and 3 extradural haemorrhages; 4 patients had more than 1 type of intracranial
20 haemorrhage). The incidence of intracranial haemorrhage was 12 per 1000 patient-years (95%
21 CI 10-14) in those with CMBs compared to 4 per 1000 patient-years (95% CI 3-5) in those
22 without CMBs, an absolute increased incidence of 8 intracranial haemorrhages per 1000
23 patient-years. Table 2 shows the incidence of intracranial haemorrhage for each CMB burden
24 and distribution category. The risk of symptomatic intracranial haemorrhage was 2.45 (95% CI
25 1.82-3.29) times higher for patients with CMBs than patients without CMBs and rose with

1 increasing CMB burden (table 2), p trend <0.0001. The presence of mixed, deep, and lobar
2 CMBs were associated with similar hazards (table 2). Patients with multiple strictly lobar
3 CMBs (fulfilling the Boston criteria⁵ for probable CAA) did not have a significantly higher
4 risk for symptomatic intracranial haemorrhage than those without (HR 1.29, 95% CI 0.60-2.77)
5 (Table 2). There was no interaction between CMBs and antiplatelet medication (p-
6 interaction=0.358), OAC (p-interaction=0.717) or combined OAC and antiplatelet medication
7 (p-interaction=0.163) for intracranial haemorrhage risk.

8

9 **Ischaemic stroke**

10

11 1113 patients had a symptomatic ischaemic stroke during 32,293 patient-years of follow up.
12 The incidence of symptomatic ischaemic stroke in patients with CMBs was 46 per 1000
13 patient-years (95% CI 42-51) compared to 30 per 1000 patient-years (95% CI 28-33) in those
14 without CMBs, an absolute increased incidence of 16 strokes per 1000 patient-years. Table 2
15 shows the incidence of ischaemic stroke for each CMB burden and distribution category. The
16 risk of symptomatic ischaemic stroke was 1.23 (95% CI 1.08-1.40) times higher for patients
17 with CMBs than patients without CMBs, and rose with increasing CMB burden (table 2), p
18 trend =0.0053. CMB distribution had little effect on ischaemic stroke risk (table 2). There was
19 no interaction between CMBs and antiplatelet medication (p-interaction=0.943) or OAC (p-
20 interaction=0.408) for ischaemic stroke risk.

21

22

23

24

25

1 **Secondary outcomes: death and vascular death**

2

3 There were 2418 deaths, 484 of which were vascular. In multivariable analyses, CMB presence
4 was not associated with all-cause death (aHR 1.03, 95% CI 0.94-1.12) or vascular death (aHR
5 0.97, 95% CI 0.79-1.19).

6

7 **Ethnicity**

8

9 There was no interaction between CMB and ethnicity (n=15123; 6743 white, 8380 Asian) for
10 the risks of: the composite outcome of intracranial haemorrhage or ischaemic stroke (p-
11 interaction=0.707); intracranial haemorrhage (p-interaction=0.537); or ischaemic stroke (p-
12 interaction=0.654).

13

14 **Age**

15

16 There was no interaction between CMB and older age (4376 patients aged >80 years) for the
17 risk of the composite outcome of intracranial haemorrhage or ischaemic stroke (p-
18 interaction=0.538); intracranial haemorrhage (p-interaction=0.219); or ischaemic stroke (p-
19 interaction=0.286).

20

21 **Sensitivity analysis to quantify heterogeneity**

22

23 Using a two-stage meta-analysis, the estimated risks associated with CMB presence were
24 consistent with our main model for: the composite of intracranial haemorrhage or ischaemic

1 stroke (heterogeneity ($I^2=31.7\%$); intracranial haemorrhage ($I^2=0\%$); and ischaemic stroke
2 ($I^2=24.2\%$). The forest plots are shown in the appendix (appendix Figure 2)

3

4 **Sensitivity analysis including white matter hyperintensities**

5

6 23 cohorts, including 10,235 patients, provided ratings for WMH, which were moderate to
7 severe (Fazekas grade ≥ 2) in 3,105 (30%). Including WMH in multivariable models did not
8 substantially change the aHR associated with the presence of CMBs for the composite outcome
9 (aHR 1.30, 95% CI 1.12-1.52), intracranial haemorrhage (aHR 2.44, 95% CI 1.68-3.53) or for
10 ischaemic stroke (aHR 1.16, 95% CI 0.98-1.37).

11

12 **Sensitivity analysis including only intracerebral, convexity subarachnoid and subdural** 13 **intracranial haemorrhage**

14

15 183 patients had a symptomatic intracranial haemorrhage during 32,847 patient-years of follow
16 up. The risk of symptomatic intracranial haemorrhage was 2.59 (95% CI 1.91-3.50) times
17 higher for patients with CMBs than patients without CMBs and rose with increasing CMB
18 burden. Compared to no CMBs, aHRs were: 1.92 (95% CI 1.25-2.94) for 1 CMB; 2.02 (95%
19 CI 1.30-3.16) for 2-4 CMBs; 4.88 (95% CI 3.29-7.25) for ≥ 5 CMBs; 5.87 (95% CI 3.56-9.66)
20 for ≥ 10 CMBs; and 9.32 (95% CI 5.06-17.16) for ≥ 20 CMBs. These results are consistent with
21 our primary findings.

22

23

24

25

1 **Incidence of outcome events beyond the first year and change in risks over time**

2

3 There were 102 intracranial haemorrhage events over 12,794 patient-years of follow up within
4 the first year, and 87 symptomatic intracranial haemorrhages over 31,059 patient-years of
5 follow-up after the first year. In patients with CMBs, the incidence of intracranial haemorrhage
6 (per 1000 patients-years) was 18 (95% CI 14-23) within the first year, and 5 (95% CI 3-6) after
7 the first year.

8

9 There were 696 ischaemic strokes over 12,873 patient -years of follow up within the first year,
10 and 417 symptomatic ischaemic strokes during 30447 patient-years of follow up after the first
11 year. In patients with CMBs the incidence of symptomatic ischaemic stroke within the first
12 year was 70 (95% CI 62-80), and 18 (95% CI 15-21) after the first year. Accounting for death
13 as a competing risk, we found no evidence for a change in risk over time associated with CMB
14 presence for intracranial haemorrhage (subdistribution Hazard Ratio [SHR] 4·96, 95% CI 3·18-
15 7·74 at day 0 vs. 4·81, 95% CI 3·15-7·35 after 1 year) or ischaemic stroke. (SHR 1·46, 95% CI
16 1·23-1·73 at day 0 vs. 1·49, 95% CI 1·27-1·75 after 1 year).

17

18 **Patients treated with oral anticoagulants**

19

20 In those treated with OAC after their index ischaemic stroke or TIA (n=7737; VKA=5253,
21 NOAC=2484) there were 91 intracranial haemorrhages during 13,942 patient-years of follow
22 up, and 384 ischaemic strokes during 13,737 patient-years of follow up. For those with CMBs,
23 the incidence of intracranial haemorrhage was 12 per 1000 patient-years (95% CI 9-16); the
24 incidence of ischaemic stroke was 32 per 1000 patient-years (95% CI 26-39). The incidence of
25 ischaemic stroke was much higher than that of intracranial haemorrhage for all CMB burden

1 and distribution categories (table 2); the risk of intracranial haemorrhage associated with CMBs
2 (vs. those without) rose more steeply with increasing CMB burden (table 2). Mixed and deep
3 CMB distributions had similar aHRs for intracranial haemorrhage, but patients with lobar
4 CMBs had a lower risk of intracranial haemorrhage (table 2). CMBs were not statistically
5 associated with ischaemic stroke risk. We found no evidence of an interaction between OAC
6 type (vitamin K antagonist vs. direct oral anticoagulant) and CMB presence for intracranial
7 haemorrhage (p-interaction=0.4) or ischaemic stroke (p-interaction=0.61).

8

9 **Patients treated with antiplatelet drugs**

10

11 In patients treated with antiplatelet drugs only (n=11520) there were 93 intracranial
12 haemorrhages during 18059 patient-years of follow up and 664 ischaemic strokes during 17731
13 patient-years of follow up. The incidence of ischaemic stroke remained higher than the
14 incidence of intracranial haemorrhage for all CMB burden and distribution categories (table 2).
15 The adjusted risks of intracranial haemorrhage and ischaemic stroke were similar to those in
16 the full cohort, with little variation according to CMB distribution (table 2).

17

18 **Patients not treated with antithrombotic medication**

19

20 Compared to patients who received antithrombotic treatment (OAC or antiplatelets), those not
21 treated with antithrombotic drugs (n=1065) were older (median age 72 vs. 70), and more likely
22 to be female (46% vs. 42%), to have presented with ischaemic stroke (91% vs. 83%), to have
23 had a previous intracranial haemorrhage (6% vs. 2%), or to have atrial fibrillation (44% vs.
24 37%). There was no difference in the prevalence of CMBs in those not receiving antithrombotic
25 treatment (29% vs. 28%). In those not treated with any antithrombotic drugs there were 5

1 intracranial haemorrhages over 846 patient-years, and 65 ischaemic strokes over 825 patient-
2 years. The aHRs associated with CMB presence were 1.10 (95% CI 0.17-7.34) for intracranial
3 haemorrhage and 1.51 (95% CI 0.87-2.65) for ischaemic stroke.

4

5 **DISCUSSION**

6

7 Our large-scale pooled analysis of individual patient data confirms that, in patients with recent
8 ischaemic stroke or TIA treated with antithrombotic drugs, CMBs are associated with the
9 subsequent risks of symptomatic intracranial haemorrhage and ischaemic stroke; as CMB
10 burden increases, the relative risk (aHR) of intracranial haemorrhage rises more steeply than
11 that of ischaemic stroke. Our most important new finding is that, regardless of CMB burden
12 and distribution (i.e. mixed, deep, or lobar), or the type of antithrombotic treatment received
13 (OAC or antiplatelet therapy), the absolute risk of ischaemic stroke is consistently substantially
14 higher than that of intracranial haemorrhage.

15

16 As well as confirming the association with CMBs and both recurrent ischaemic stroke and
17 symptomatic intracranial haemorrhage found in smaller cohorts of patients with ischaemic
18 stroke and TIA treated with antiplatelet agents²⁸ or OACs,^{27,35,36} the large number of
19 participants has improved the precision of our estimates of stroke recurrence rates and risks,
20 while allowing adjustment for potential confounding factors. Our study also adds new data on
21 important subgroups of patients with many (e.g. ≥ 20) CMBs, which cause most clinical
22 concern and could not be addressed by any of the previously published meta-analyses.
23 Moreover, only our study has undertaken individual patient data pooled analyses to allow
24 adjustment for important potential confounding factors. The association of CMBs with a
25 consistently higher absolute risk of ischaemic stroke than intracranial haemorrhage suggests

1 that CMBs are a marker for cerebral small vessel diseases that can cause not only intracranial
2 haemorrhage, but also ischaemic stroke. Although it has been inferred that CMBs are a marker
3 of direct extravasation of red blood cells from arterioles and capillaries damaged by “bleeding-
4 prone arteriopathies”, alternative non-haemorrhagic mechanisms include ischemia-mediated
5 iron store release by oligodendrocytes¹⁰ or phagocytosis of red cell micro-emboli into the
6 perivascular space¹¹. A recent report of haemorrhagic transformation of small acute
7 “microinfarcts” into CMBs provides direct evidence that CMBs can result from ischaemic
8 mechanisms¹³. These varied mechanisms underlying CMBs might explain why even patients
9 at highest risk of intracranial haemorrhage still have a higher absolute risk of ischaemic stroke.
10 Indeed, patients with CMBs often have multiple vascular risk factors, so are at risk of not only
11 small vessel ischemic stroke but also other ischaemic stroke subtypes³⁷. Patients with CMBs
12 usually also have WMH - themselves associated with the risk of recurrent stroke, death and
13 poor functional outcome after ischaemic stroke³⁸ - which might also contribute to the increased
14 risk of ischaemic stroke associated with CMBs.

15
16 We found no evidence that a strictly lobar pattern of CMBs (fulfilling the Boston criteria for
17 probable CAA⁵, causing particular clinical concern for intracranial bleeding risk³⁶) is
18 associated with the risk of intracranial haemorrhage or ischaemic stroke. These findings might
19 reflect limited diagnostic accuracy when using CMBs for diagnosis of CAA in patients without
20 intracerebral haemorrhage or dementia³⁹, rather than a true lack of any association of CAA
21 with intracranial haemorrhage. Furthermore, the HRs for intracranial haemorrhage associated
22 with lobar CMBs (when compared to patients without lobar CMBs [including none]) were
23 closer to those associated with deep or mixed CMBs (when compared to patients without deep
24 or mixed CMBs [including none]).

25

1 Our results differ from some previous observations in smaller cohorts. First, in contrast to a
2 smaller two-centre study²⁹, we did not find that the risk of intracranial haemorrhage approached
3 the risk of ischaemic stroke after one year. Rather, we found the risk associated with CMB for
4 both ischaemic stroke and intracranial haemorrhage remained stable over time. Second, our
5 data indicate a smaller increase in the relative risk of intracranial haemorrhage for patients with
6 ≥ 5 CMBs than reported in a previous smaller meta-analysis²⁸, but our much larger individual
7 participant sample size allowed us to investigate high CMB burdens (≥ 5 , ≥ 10 and ≥ 20) with
8 adjustment for confounders and greater statistical precision and power.

9
10 The comparatively low frequency of symptomatic intracranial haemorrhage after ischaemic
11 stroke or TIA, and the consistently higher risk of recurrent ischaemic stroke, make randomised
12 controlled trials of antithrombotic treatment (themselves proven in large randomised trials)
13 guided by CMBs challenging. However, ongoing and future randomised controlled trials
14 should provide further insights. The MRI sub-study in the RESTART trial⁴⁰ of antiplatelet
15 therapy after intracerebral haemorrhage identified reassuring effects in patients with CMB, but
16 also illustrates how very large sample sizes are likely to be required to identify statistically
17 significant interactions in CMB sub-groups in current (e.g. the MRI sub-study of NAVIGATE
18 ESUS [NCT02313909]), and future, randomised controlled trials. Nevertheless, our large
19 collaborative pooled analysis provides the best available evidence on the associations of CMBs
20 with intracranial haemorrhage and ischaemic stroke after ischaemic stroke or TIA.

21
22 Our study has strengths. We included data from a worldwide collaborative network, making
23 our results globally generalisable. The large individual patient dataset provides high statistical
24 power and precision for risk estimates, allowing us to explore associations with several
25 clinically important primary outcomes, while adjusting for important prognostic variables to

1 minimise confounding. The large number of events allow us to add new data on important
2 subgroups of patients with a large number of CMBs (e.g. ≥ 20 CMBs) which are those patients
3 who cause most clinical concern; this issue could not be addressed by any of the previously
4 published meta-analyses. Included cohorts used validating rating instruments for CMBs, and
5 in our analyses we accounted for the use of different MRI sequences (T2* GRE or SWI) to
6 detect CMBs which accounts for the higher sensitivity of SWI for detecting CMBs when
7 compared with T2* GRE⁴¹. Finally, we followed a published statistical analysis plan and
8 confirmed our findings in a two-stage meta-analysis indicating the robustness of our results.

9

10 We also acknowledge limitations. Our observational study design has potential for selection
11 bias and confounding of antithrombotic therapy by indication or unmeasured physician factors;
12 thus, the hazard ratios for intracranial haemorrhage and ischaemic stroke must be interpreted
13 with caution. To definitively determine whether CMBs modify the net clinical benefit of
14 antithrombotic drugs would require a randomised controlled trial. Many of the included studies
15 did not undertake formal adjudication of events. The requirement for MRI suitable patients
16 probably led to the inclusion of less severe strokes than an unselected ischaemic stroke
17 population. Even with a large number of individual patients, we could not precisely estimate
18 risks associated with an extremely large number of CMBs (e.g. ≥ 50), but such patients are very
19 rare in clinical practice. Although we adjusted for known prognostic variables, it is possible
20 that residual confounding secondary to unknown or uncontrolled factors such as stroke
21 mechanism could still have affected our results. Furthermore, we were unable to include some
22 candidate variables in our multivariable models because they were not sufficiently widely
23 available across all participating cohorts (e.g. white matter hyperintensities, MRI field strength,
24 diabetes, ischaemic heart disease, renal function, statin use on discharge). Our analyses did not

1 formally assess net clinical benefit, accounting for the greater severity of intracranial
2 haemorrhage compared to recurrent ischaemic stroke.

3

4 In summary, our large-scale pooled analysis in patients with recent ischaemic stroke or TIA
5 found that the absolute risk of ischaemic stroke is consistently higher than that of intracranial
6 haemorrhage, regardless of the number or anatomical distribution of CMBs. However, CMBs
7 are associated with a greater adjusted hazard ratio for intracranial haemorrhage than ischaemic
8 stroke; further studies are needed to establish the value of neuroimaging biomarkers, including
9 CMBs, in improving risk prediction scores for intracranial haemorrhage and ischaemic stroke.

10

1 **CONTRIBUTIONS**

2

3 DJW, DW, GA and JM-F drafted the initial protocol, which was reviewed with critical
4 revisions and approval by all authors. DW and GA undertook the statistical analysis. DW, DJW
5 and GA wrote the first draft of the manuscript. All authors contributed to data acquisition,
6 management, and brain imaging analyses. All authors contributed to critical revision of the
7 manuscript and approved the final manuscript for submission.

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2

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Table 1: Demographics, risk factors and events for each included cohort

Cohort	Total n	OAC n	TIA (%)	Age mean, years (SD)	Sex female (%)	HTN (%)	AF (%)	Prior stroke (%)	IHD (%)	Any CMB (%)	SWI (%)	Median follow up, days (IQR)	Patients with composite events (n)	Patients with ICH Events (n)	Patients with ischaemic stroke events (n)
CROMIS-2 ²⁷	1490	1436	16	76 (10)	42	63	100	10	16	21	0	774 (705-974)	70	14	56
HBS	660	114	9	69 (15)	44	75	29	24	21	15	0	90 (90-90)	4	0	4
Bern ⁴²	392	74	0	68 (14)	43	64	46	15	18	23	100	93 (5-106)	16	0	16
CU-STRIDE ⁴³	536	24	15	67 (11)	42	70	7	15	7	23	44	524 (472-557)	17	2	15
TABASCO ⁴⁴	436	33	28	67 (9)	43	59	8	0	14	15	0	1825 (1164-1825)	57	0	57
Graz	460	78	10	67 (13)	39	78	25	22	21	19	0	117 (87-973)	65	13	54
PERFORM-MRI ⁴⁵	1056	0	12	68 (8)	35	84	2	11	7	36	0	774 (701-1042)	104	10	94
PARISK ⁴⁶	228	0	56	71 (9)	29	68	0	29	22	27	0	786 (757-819)	10	0	10
SAMURAI NVA ⁴⁷	1103	1039	4	78 (10)	44	93	100	22	9	24	74	723 (758-818)	82	10	72
RUNDMC ⁴⁸	179	19	50	65 (9)	35	81	10	26	17	20	0	1825 (1825-1825)	25	2	23
Wuerzburg	358	122	22	71 (13)	44	80	37	25	11	24	45	95 (89-103)	22	1	21
Monash Stroke ⁴⁹	359	356	15	75 (11)	48	79	100	28	34	43	94	530 (280-898)	14	7	9
Basel TIA ¹⁸	192	33	100	69 (13)	38	71	14	7	20	11	0	90 (90-90)	26	0	26
Yonsei ⁵⁰	504	487	6	70 (11)	57	78	100	20	22	31	0	849 (393-1398)	56	7	49

SNUBH Stroke Cohort ^{51,52}	3355	625	11	67 (13)	40	69	19	15	8	35	0	355 (340-365)	172	N/A	N/A
BIOSTROKE/TIA ⁵³	260	73	62	68 (13)	37	59	31	8	22	9	0	90 (90-365)	14	0	14
Kushiro City ⁵⁴	784	63	0	72 (11)	42	64	13	18	11	41	0	1008 (105-1825)	139	22	119
Soo ⁵⁵	81	81	20	72 (9)	49	69	100	31	10	30	88	737 (641-794)	8	3	5
CASPER ⁵⁶	135	18	0	66 (11)	29	71	10	7	21	59	100	453 (444-465)	3	0	3
HERO	937	933	13	78 (7)	52	74	92	27	16	26	0	737 (641-794)	49	18	32
HAGAKURE	426	157	8	74 (13)	41	76	32	18	11	37	9	748 (350-1040)	34	9	25
Leuven ¹⁴	487	133	27	72 (9)	39	64	21	13	23	26	0	804 (686-968)	36	4	32
NOACISP	306	286	10	73 (19)	45	79	100	20	27	28	98	735 (417-836)	28	10	19
Min Lou ⁵⁷	126	14	0	65 (13)	37	75	20	14	3	33	100	92 (87-218)	2	0	2
MICRO ²¹	397	40	91	65 (12)	42	55	8	9	6	18	0	1212 (579-1825)	30	11	21
Orken ⁵⁸	454	454	4	72 (12)	51	79	65	27	32	30	55	575 (228-1825)	11	3	8
CATCH ⁵⁹	416	67	42	67 (14)	39	54	6	0	N/A	16	0	88 (80-100)	14	1	13
MSS2 ⁶⁰	263	26	0	67 (12)	41	72	10	12	20	16	95	368 (253-403)	31	0	31
Sainte-Anne	385	302	0	80 (11)	53	72	100	16	19	26	0	440 (163-733)	25	5	23
STROKDEM	181	48	0	64 (13)	38	55	7	11	9	13	0	1150 (420-1820)	17	0	17
Singapore	45	15	0	67 (10)	29	76	24	7	9	56	100	1057 (703-1199)	6	0	6
FUTURE Study	19	0	37	44 (6)	53	42	0	0	0	5	100	164 (131-242)	4	0	4
Heidelberg ⁶¹	650	119	16	64 (14)	37	76	18	17	N/A	24	100	1534	34	4	30

												(1271-1825)			
NNI	184	32	0	58 (11)	31	78	15	15	N/A	27	0	251 (86-477)	0	0	0
OXVASC²⁹	1080	118	52	68 (14)	48	55	15	19	13	15	0	1271 (681-1825)	90	11	79
HKU²⁹	1003	104	0	69 (12)	40	66	13	19	9	45	100	1005 (599-1549)	112	20	92
IPAAC Warfarin³⁶	182	173	15	73 (9)	46	87	98	20	9	37	100	738 (191-812)	7	1	6
SIGNaL	213	18	10	72 (14)	41	61	32	28	23	44	68	225 (202-249)	27	1	26
TOTAL	20322	7737	761	70 (13)	42	71	38	16	14	28	39	534 (243-928)	1461	189	1113

Footnote: OAC -oral anticoagulant; TIA -transient ischaemic attack; HTN -hypertension; AF- Atrial fibrillation; IHD- ischaemic heart disease; SWI – susceptibility-weighted imaging; ICH (intracranial haemorrhage).

Table 2. Incidence and risk of outcome events according to baseline CMBs

ALL PATIENTS (n=20322)												
	Composite of intracranial haemorrhage and ischaemic stroke (n=19816 for multivariable model)				Symptomatic Intracranial haemorrhage (n=16447 for multivariable model)				Symptomatic ischaemic stroke (n=16464 for multivariable model)			
CMB category	n/time at risk (years)	Rate/1000 patient years (95% CI)	Absolute rate increase/1000 patient years (95% CI)	Adjusted Hazard ratio (95% CI)	n/time at risk (years)	Rate/1000 patient years (95% CI)	Absolute rate increase/1000 patient years (95% CI)	Adjusted Hazard ratio (95% CI)	n/time at risk (years)	Rate/1000 patient years (95% CI)	Absolute rate increase/1000 patient years (95% CI)	Adjusted Hazard ratio (95% CI)
No CMB	919/25997	35 (33 to 38)	Reference	Reference	91/24496	4 (3 to 5)	Reference	Reference	733/24094	30 (28 to 33)	Reference	Reference
CMB present	542/9228	59 (54 to 64)	24 (21 to 26)	1.35 (1.20 to 1.50)	98/8351	12 (10 to 14)	8 (7 to 9)	2.45 (1.82 to 3.29)	380/8200	46 (42 to 51)	16 (14 to 18)	1.23 (1.08 to 1.40)
1 CMB	190/4095	46 (40 to 53)	11 (7 to 15)	1.21 (1.03 to 1.42)	30/3691	8 (5 to 12)	4 (2 to 7)	1.87 (1.23 to 2.84)	134/3635	37 (31 to 44)	7 (3 to 11)	1.14 (0.94 to 1.37)
2-4 CMBs	184/3166	58 (50 to 67)	23 (17 to 29)	1.25 (1.06 to 1.47)	27/2853	9 (6 to 14)	5 (3 to 9)	1.89 (1.22 to 2.93)	133/2797	48 (40 to 56)	18 (12 to 23)	1.17 (0.97 to 1.42)
≥5 CMBs*	168/1967	85 (73 to 99)	50 (40 to 61)	1.74 (1.46 to 2.06)	41/1807	23 (16 to 31)	19 (13 to 26)	4.55 (3.08 to 6.72)	113/1767	64 (53 to 77)	34 (25 to 43)	1.47 (1.19 to 1.80)
≥10 CMBs*	82/898	91 (73 to 113)	56 (40 to 75)	1.82 (1.44 to 2.29)	22/818	27 (17 to 41)	23 (14 to 36)	5.52 (3.36 to 9.05)	52/809	64 (48 to 84)	34 (20 to 51)	1.43 (1.07 to 1.91)
≥20 CMBs*	42/363	118 (86 to 160)	83 (53 to 122)	2.61 (1.90 to 3.57)	13/332	39 (21 to 67)	35 (18 to 62)	8.61 (4.69 to 15.81)	24/331	73 (46 to 108)	43 (18 to 75)	1.86 (1.23 to 1.82)
Mixed CMBs	160/1991	80 (68 to 94)	45 (35 to 56)	1.28 (1.06 to 1.54)	35/1721	20 (14 to 28)	16 (11 to 23)	2.38 (1.55 to 3.65)	101/1677	60 (49 to 73)	30 (21 to 40)	1.12 (0.88 to 1.41)
Deep CMBs	307/4202	73 (65 to 82)	38 (32 to 44)	1.29 (1.12 to 1.48)	59/3468	17 (13 to 22)	13 (10 to 17)	2.57 (1.78 to 3.70)	194/3398	57 (49 to 66)	27 (21 to 33)	1.14 (0.96 to 1.36)
Lobar CMBs	278/4532	60 (53 to 67)	25 (20 to 29)	1.22 (1.06 to 1.41)	51/4071	13 (9 to 16)	9 (6 to 9)	1.87 (1.29 to 2.71)	193/3983	48 (42 to 56)	18 (14 to 23)	1.17 (0.99 to 1.40)
Probable CAA	45/821	55 (40 to 73)	20 (7 to 35)	1.21 (0.90 to 1.64)	7/789	9 (4 to 18)	5 (1 to 13)	1.29 (0.60 to 2.77)	37/767	48 (34 to 66)	18 (6 to 33)	1.31 (0.94 to 1.83)

PATIENTS TREATED WITH ORAL ANTICOAGULANTS (+/- antiplatelet drugs) (n=7737)												
	Composite of intracranial haemorrhage and ischaemic stroke (n=7582 for multivariable model)				Symptomatic Intracranial haemorrhage (n=6942 for multivariable model)				Symptomatic ischaemic stroke (n=6958 in multivariable models)			
	n/time at risk (years)	Rate/1000 patient years (95% CI)	Absolute rate increase/1000 patient years (95% CI)	Adjusted Hazard ratio (95% CI)	n/time at risk (years)	Rate/1000 patient years (95% CI)	Absolute rate increase/1000 patient years (95% CI)	Adjusted Hazard ratio (95% CI)	n/time at risk (years)	Rate/1000 patient years (95% CI)	Absolute rate increase/1000 patient years (95% CI)	Adjusted Hazard ratio (95% CI)
No CMB	332/10591	31 (28 to 35)	<i>Reference</i>	<i>Reference</i>	47/10383	5 (3 to 6)	<i>Reference</i>	<i>Reference</i>	271/10221	27 (23 to 30)	<i>Reference</i>	<i>Reference</i>
CMB present	168/3671	46 (39 to 53)	15 (11 to 18)	1.30 (1.07 to 1.57)	44/3559	12 (9 to 16)	7 (6 to 10)	2.49 (1.64 to 3.79)	113/3515	32 (26 to 39)	5 (3 to 9)	1.07 (0.86 to 1.35)
1 CMB	65/1689	38 (30 to 49)	7 (2 to 14)	1.19 (0.91 to 1.56)	17/1641	10 (6 to 17)	5 (3 to 11)	2.15 (1.23 to 3.75)	42/1624	26 (19 to 35)	-1 (-4 to 5)	0.96 (0.69 to 1.33)
2-4 CMBs	61/1307	47 (36 to 60)	16 (8 to 25)	1.23 (0.93 to 1.62)	14/1267	11 (6 to 19)	6 (3 to 13)	2.22 (1.21 to 4.06)	45/1245	36 (26 to 48)	11 (3 to 18)	1.10 (0.80 to 1.52)
≥5 CMBs*	42/675	62 (45 to 84)	31 (17 to 49)	1.69 (1.22 to 2.35)	13/651	20 (11 to 34)	15 (8 to 28)	3.91 (2.08 to 7.34)	26/650	40 (26 to 59)	13 (3 to 29)	1.27 (0.84 to 1.91)
≥10 CMBs*	20/266	75 (46 to 116)	44 (18 to 81)	2.15 (1.35 to 3.43)	6/261	23 (8 to 50)	18 (5 to 44)	4.63 (1.92 to 11.22)	12/260	46 (24 to 81)	19 (1 to 51)	1.52 (0.84 to 2.67)
Mixed CMBs	46/793	58 (42 to 77)	27 (14 to 42)	1.43 (1.02 to 2.00)	11/751	15 (7 to 26)	10 (4 to 20)	2.21 (1.09 to 4.47)	31/737	42 (29 to 60)	15 (6 to 30)	1.28 (0.85 to 1.94)
Deep CMBs	91/1763	52 (42 to 63)	21 (14 to 28)	1.43 (1.11 to 1.84)	24/1665	14 (9 to 21)	9 (6 to 15)	2.71 (1.61 to 4.59)	58/1661	35 (27 to 46)	8 (4 to 16)	1.16 (0.85 to 1.59)
Lobar CMBs	80/1969	41 (32 to 51)	10 (4 to 16)	1.13 (0.87 to 1.47)	20/1905	10 (6 to 16)	5 (3 to 10)	1.63 (0.94 to 2.83)	54/1877	29 (22 to 38)	2 (-1 to 8)	1.00 (0.73 to 1.38)
Probable CAA	11/415	27 (13 to 47)	-4 (-15 to 12)	0.76 (0.41 to 1.39)	4/413	10 (3 to 25)	5 (0 to 19)	1.29 (0.47 to 3.57)	7/407	17 (7 to 35)	-10 (-16 to 5)	0.64 (0.30 to 1.37)

PATIENTS TREATED WITH ANTIPLATELET DRUGS ONLY (n=11520)												
Composite of intracranial haemorrhage and ischaemic stroke (n=11312 in multivariable models)					Symptomatic intracranial haemorrhage (n=8670 in multivariable models)				Symptomatic ischaemic stroke (n=8670 in multivariable models)			
	n/time at risk (years)	Rate/1000 patient years (95% CI)	Absolute rate increase/ 1000 patient years (95% CI)	Adjusted Hazard ratio (95% CI)	n/time at risk (years)	Rate/1000 patient years (95% CI)	Absolute rate increase/ 1000 patient years (95% CI)	Adjusted Hazard ratio (95% CI)	n/time at risk (years)	Rate/1000 patient years (95% CI)	Absolute rate increase/ 1000 patient years (95% CI)	Adjusted Hazard ratio (95% CI)
No CMB	544/14798	37 (34 to 40)	<i>Reference</i>	<i>Reference</i>	41/13518	3 (2 to 4)	<i>Reference</i>	<i>Reference</i>	424/13290	32 (29 to 35)	<i>Reference</i>	<i>Reference</i>
CMB present	343/5301	65 (58 to 72)	28 (24 to 32)	1.38 (1.19 to 1.59)	52/4541	11 (9 to 15)	8 (7 to 11)	2.59 (1.68 to 4.00)	240/4442	54 (47 to 61)	22 (18 to 26)	1.32 (1.11 to 1.56)
1 CMB	114/2340	49 (40 to 59)	12 (6 to 19)	1.20 (0.97 to 1.47)	12/1984	6 (3 to 11)	3 (1 to 7)	1.60 (0.83 to 3.08)	82/1951	42 (33 to 52)	10 (4 to 17)	1.22 (0.96 to 1.55)
2-4 CMBs	111/1759	63 (52 to 76)	26 (18 to 36)	1.25 (1.02 to 1.55)	13/1487	9 (5 to 15)	6 (3 to 11)	1.82 (0.95 to 3.47)	78/1457	54 (42 to 67)	22 (13 to 32)	1.20 (0.94 to 1.55)
≥5 CMBs*	118/1203	98 (81 to 117)	61 (47 to 77)	1.85 (1.50 to 2.28)	27/1069	25 (17 to 37)	22 (15 to 33)	5.69 (3.36 to 9.65)	80/1034	77 (61 to 96)	45 (32 to 61)	1.63 (1.27 to 2.09)
≥10 CMBs*	59/593	99 (76 to 128)	62 (42 to 88)	1.82 (1.38 to 2.40)	16/519	31 (18 to 50)	28 (16 to 46)	6.81 (3.67 to 12.63)	37/511	72 (51 to 100)	40 (22 to 65)	1.47 (1.04 to 2.07)
≥20 CMBs*	30/221	136 (92 to 194)	99 (58 to 154)	2.86 (1.97 to 4.96)	11/194	57 (28 to 101)	54 (26 to 97)	15.71 (7.69 to 32.11)	16/192	83 (48 to 135)	51 (19 to 100)	2.00 (1.20 to 3.32)
Mixed CMBs	102/1121	91 (74 to 110)	54 (40 to 70)	1.21 (0.96 to 1.52)	24/896	27 (17 to 40)	24 (15 to 36)	2.90 (1.64 to 5.13)	60/867	69 (53 to 89)	37 (21 to 54)	1.03 (0.76 to 1.39)
Deep CMBs	193/2320	83 (72 to 96)	46 (38 to 56)	1.22 (1.02 to 1.47)	34/1692	20 (14 to 28)	17 (12 to 24)	2.63 (1.54 to 4.47)	116/1647	70 (58 to 84)	38 (29 to 49)	1.11 (0.88 to 1.40)
Lobar CMBs	180/2423	74 (64 to 86)	37 (30 to 46)	1.26 (1.05 to 1.52)	31/2029	15 (10 to 22)	12 (8 to 18)	2.44 (1.44 to 4.15)	123/1973	62 (52 to 74)	30 (23 to 39)	1.26 (1.01 to 1.57)
Probable CAA	32/372	86 (59 to 121)	49 (25 to 81)	1.52 (1.07 to 2.18)	3/340	9 (2 to 26)	6 (0 to 22)	1.43 (0.44 to 4.59)	28/326	86 (57 to 124)	54 (28 to 89)	1.79 (1.21 to 2.63)

Footnote: CMB location HRs are compared to patients without CMBs in each location and are adjusted for CMB number in addition to our pre-specified variables, *overlapping categories

Figure 1 legend

Kaplan-Meier estimates for the primary outcomes: (A) All patients (n=20,322); (B) Patients treated with oral anticoagulants (with or without antiplatelet drugs) (n=7,737); and (C) Patients treated with antiplatelet drugs only (n=11,520).