

VALIDATION OF LUTEOLYSIS MONITORING TOOL

1 **INTERPRETATIVE SUMMARY. Validation of luteolysis monitoring tool for dairy cows.**

2 **Adriaens.**

3 In this study, the performance of two monitoring algorithms to detect luteolysis using milk
4 progesterone measurements was validated on a simulated dataset of realistic milk progesterone
5 profiles. The synergistic control-based algorithm, PMASC, was able to identify luteolysis almost
6 simultaneously with its occurrence. It was found to be more robust against missing samples and less
7 dependent on the absolute milk progesterone values compared to a multiprocess Kalman filter
8 combined with a fixed threshold. This research showed that implementation of PMASC could
9 improve progesterone-based fertility monitoring on farm.

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12 **Validation of a novel milk progesterone-based tool to monitor luteolysis in dairy cows. Timing**
13 **of the alerts and robustness against missing values**

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26 **ABSTRACT**

27 Automated monitoring of fertility in dairy cows using milk progesterone is based on the accurate and
28 timely identification of luteolysis. In this way, well-adapted insemination advice can be provided to
29 the farmer to further optimize the fertility management. To properly evaluate and compare the
30 performance of new and existing data-processing algorithms, a test dataset of progesterone time-
31 series that fully covers the desired variability in progesterone profiles is needed. Further, the data
32 should be measured with a high frequency to allow rapid onset events, such as luteolysis, to be
33 precisely determined. Collecting this type of data would require a lot of time, effort and budget. In
34 the absence of such data, an alternative was developed using simulated progesterone profiles for
35 multiple cows and lactations, in which the different fertility statuses were represented. To these,
36 relevant variability in terms of cycle characteristics and measurement error was added, resulting in a
37 large cost-efficient dataset of well-controlled but highly variable and farm-representative profiles.
38 Besides the progesterone profiles, information on (the timing of) luteolysis was extracted from the
39 modelling approach and used as a reference for the evaluation and comparison of the algorithms. In
40 this study, two progesterone monitoring tools were compared: a multiprocess Kalman filter combined
41 with a fixed threshold on the smoothed progesterone values to detect luteolysis, and a progesterone
42 monitoring algorithm using synergistic control ‘PMASC’, which uses a mathematical model based
43 on the luteal dynamics and a statistical control chart to detect luteolysis. The timing of the alerts and
44 the robustness against missing values of both algorithms were investigated using two different
45 sampling schemes: one sample per cow every eight hours versus one sample per day. The alerts for
46 luteolysis of the PMASC algorithm were on average 20 hours earlier compared to the ones of the
47 multiprocess Kalman filter, and their timing was less sensitive to missing values. This was shown by
48 the fact that, when one sample per day was used, the Kalman filter gave its alerts on average 24 hours
49 later, and the variability in timing of the alerts compared to simulated luteolysis increased with 22%.

50 Accordingly, we postulate that implementation of the PMASC system could improve the consistency
51 of luteolysis detection on farm and lower the analysis costs compared to the current state of the art.

52 **Key words.** milk progesterone, fertility, dairy cow, simulation, monitoring tool

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INTRODUCTION

55 Monitoring of milk progesterone (**P4**) in dairy cows allows identification of a cows' reproduction
56 status. Because P4 is fat-soluble and transfers from the blood into the milk, the concentration of P4
57 in milk is 4 to 5 times higher than in blood. High P4 concentrations, produced by an active corpus
58 luteum (**CL**) on the ovaries are associated with the luteal phase of the cycle or pregnancy, while low
59 P4 concentrations are known to occur during the follicular phase of the P4 cycle and in the postpartum
60 anestrus phase after calving. Luteolysis, under influence of the uterine $\text{PGF}_{2\alpha}$ signal and defined as
61 the regression of the CL, is accompanied with a steep and fast decrease in P4, seen as a drop in milk
62 P4 from over 15 ng/mL to below 5 ng/mL in approximately 12 to 24 hours. This drop in P4 is
63 necessary to allow for a LH surge that induces rupture of a pre-ovulatory follicle (ovulation). Estrus
64 detection based on milk P4 dynamics therefore relies on the accurate and timely identification of
65 luteolysis preceding ovulation. Since recently, it is possible to automatically measure milk P4 on
66 farm, in which regular milk analyses clearly show the P4 dynamics during an estrous cycle (Adriaens
67 et al., 2017; Bruinjé et al., 2017). The current state-of-the-art in P4-based fertility monitoring is to
68 smooth the raw measured values with a multi-process Kalman filter (**MPKF**), after which a fixed
69 threshold (**T**) to these smoothed values is applied to detect luteal activity and luteolysis (Friggens and
70 Chagunda, 2005; Friggens et al., 2008). The MPKF hereby ensures that no alerts are triggered for a
71 single low measurement. The set threshold's value might depend on the P4 measurement technique
72 or the calibration method, but is generally taken between 4 and 6 ng/mL (Friggens et al., 2008; Bruinjé
73 et al., 2017). In contrast, the P4 monitoring algorithm using synergistic control (**PMASC**) enables the
74 identification of fertility events on farm using the underlying physiological basis of the related P4

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75 dynamics (Adriaens et al., 2017, 2018a). It employs a combination of mathematical functions to
76 describe the development and regression of the CL and a statistical control chart for detection of
77 luteolysis. Until now, this system was designed, optimized and evaluated on high-frequent P4
78 measurements in which milk P4 was analyzed *post-hoc* via ELISA-testing in the lab. This did not yet
79 represent on-farm measured data for which the measurement error is representative, the time series
80 are sufficiently long and in which all the variability in P4 profiles on farm is included. Before the
81 PMASC algorithm can be used on farm, it should therefore also be validated as such.

82 In the ideal scenario, this validation would be performed on a large dataset representative for on-
83 farm measurements and containing numerous milk P4 profiles with as much variability as possible,
84 not only in fertility and profile characteristics (e.g. including follicular and luteal cysts, early and late
85 pregnancy and embryonic losses within a lactation) but also in cycle shapes (e.g. height, baseline and
86 slopes of each P4 cycle). The frequency of measurement should be as high as possible (e.g. once per
87 milking) in order to be able to vary sampling schemes and test all possible scenarios. Moreover, to
88 validate a luteolysis monitoring algorithm, the actual moment of luteolysis should ideally be known
89 in order to use this as the ‘gold standard’. To our knowledge, and especially with respect to time of
90 luteolysis measures, a dataset does not exist that meets all these criteria.

91 Alternatively, a convenient and more efficient way to obtain an appropriate dataset is through
92 simulations, which allow generation of extensive datasets while avoiding analysis costs both in terms
93 of measurements and time. Recently, a systemic white box model representing a virtual cow
94 (GARUNS, Martin and Sauvant, 2010) coupled to a model describing the reproductive functioning
95 (reproduction function model, **RFM**, Martin et al., 2018) was developed. This model allows
96 simulation of virtual cows with diverse fertility characteristics, and provides scaled P4 profiles
97 corresponding to the reproductive functioning of the simulated cows. The RFM outputs scaled P4
98 profiles, meaning that the dynamics are representative for the fertility, but need to be adjusted to
99 represent the targeted measurement technique and substrate (milk vs. blood), from which we know

100 the absolute values can vary. In this way, large datasets representing cows with sufficiently variable
 101 fertility characteristics, for which the number of estruses and timing of luteolyses is known, can be
 102 obtained.

103 In a previous validation study, PMASC was shown to be successful in correctly identifying
 104 luteolysis using milk P4 data measured on-farm with a lateral flow immunoassay (LFIA) (Adriaens
 105 et al., 2019). Nevertheless, as those P4 data originated from a commercial sensor system, the sampling
 106 frequency was set by the system software and the effect of the measurement frequency and timing,
 107 as well as the effect of missing values on the performance of PMASC could not be evaluated.
 108 Additionally, the timing of the alerts generated by PMASC and MPKF+T could not be verified as no
 109 reference information on the exact moment of luteolysis was available on those farms. Therefore, the
 110 objective of the current study was to compare the detection performance, robustness and consistency
 111 of alerts from PMASC and the MPKF+T method based on simulated realistic P4 profiles.
 112 Furthermore, this approach allowed for the evaluation of the effect of missing samples or a reduced
 113 sampling scheme during luteolysis. A good understanding of the performance of the P4 monitoring
 114 algorithm under variable sampling conditions and schemes is important to quantify the uncertainty in
 115 the prediction of the optimal insemination window.

116

117

MATERIALS AND METHODS

Simulating Progesterone Profiles

119 In a first step, 100 dairy cows were simulated using a systemic model for describing lifetime
 120 performance in dairy cows (GARUNS), developed by Martin et al. (Martin and Sauvant, 2010; Martin
 121 et al., 2013, 2018; Gaillard et al., 2016). The fertility characteristics of these cows were defined by a
 122 recently developed reproduction module coupled to GARUNS, the RFM, for which a schematic
 123 overview is shown in Figure A1. The general idea of the RFM is that a cow shifts continuously
 124 between 11 different fertility compartments, namely ‘prepubertal’, ‘anestrous’, ‘anovulatory’, ‘pre-

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125 ovulating', 'ovulating', 'post-ovulating', 'luteinizing', 'luteal', 'cystic', 'dysfunctional' and
126 'gestating'. The dynamics of these shifts are influenced by the lifetime performance model GARUNS,
127 for instance by the rate of hormonal clearance and the energy balance. In turn, the RFM' outputs
128 'conception' and 'embryonic/fetal death' manipulate the course of GARUNS, and thus the life of a
129 cow, to trigger body weight changes, dry-off, initiation of a new lactation and so on. The competence
130 stages 'cystic' and 'dysfunctional' are modifications of the original published work (Martin et al.,
131 2018) that allow for interruption of cyclicity in absence of P4 and a P4-producing (luteal) cyst-like
132 structure, respectively. The first can be either sudden anestrus (no activity on the ovaries) or a
133 follicular cyst (large, fluid filled follicular structure on the ovaries) (Ranasinghe et al., 2011; Jeengar
134 et al., 2014). The latter results in intermediate P4 concentrations during the luteal phase as described
135 by Braw-Tal et al., (2009), Peter et al., (2009) and Rosenberg, (2010). The parameters of the model
136 (Table A1) were chosen to obtain cows with a large range of fertility characteristics, both in terms of
137 length of the postpartum anestrus period, number and length of the cycles, occurrence of interrupted
138 cyclicity due to follicular and luteal cysts and the interval to successful pregnancy.

139 Each simulated cow had 6 to 7 lactations with different fertility features reflected in the scaled P4
140 profiles (example in Figure 1, red). A subset of lactations was selected based on the variability in P4
141 profile characteristics. For this, the profiles were successively sorted by postpartum anestrus length,
142 number of cycles and incidence of interrupted cyclicity, after which each time the 50 most variable
143 lactations were selected. This resulted in a dataset of 150 profiles (i.e. the consecutive dynamics over
144 1 lactation) containing in total 731 scaled estrous cycles. The dataset characteristics are summarized
145 in Table 1.

146 The outputs of the GARUNS/RFM simulations are scaled P4 profiles (i.e. representing
147 reproduction performance over a lactation) of cows reflecting realistic fertility characteristics and
148 with a measurement frequency of 1 sample per 2 hours (Martin et al., 2018). However, as these
149 profiles are scaled, they do not represent the variability at cycle level (de-scaling), nor do they contain

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150 measurement noise associated with the P4 measurement in milk. Accordingly, two additional steps
151 were required. The first step was to allow variability in the length and the relative P4 concentrations
152 of the follicular and luteal phases. To this end, and to ensure maximal variability, the following
153 properties of each cycle were adjusted according to a value randomly sampled from a uniform
154 distribution, chosen according to the characteristics described by Meier et al. (2009a), Gorzecka et
155 al. (2011) and Blavy et al. (2016):

156 (1) The rate of decrease in P4 during luteolysis. In this step, the duration of the drop in P4 from
157 maximum to minimum concentration was adjusted to last between 0.5 and 3 days (Gorzecka
158 et al., 2011; Bruinjé et al., 2017). The uniform distribution used was thus $U[0.5;3]$.

159 (2) Adjustment of cycle height and shape. Meier and colleagues reported three different types of
160 serum P4 profiles: ‘peaked’ profiles without a clear platform, ‘flat-top’ profiles with a
161 distinguishable platform of constant high P4 production, and ‘structured’ profiles in which
162 the P4 seems to rise in two phases (Meier et al., 2009b). To simulate these different types, the
163 maximum P4 value in cycle was adjusted with a certain percentage varying from 40 to 100%
164 (distribution $U[0.4;1]$) by either cutting off the data higher than this percentage or by
165 multiplying the whole cycle with this percentage. The first procedure results in a ‘flat top’
166 shaped cycle, the second in a ‘peaked’ cycle. A structured shape was not considered, because
167 in contrast to P4 in serum, this shape was not yet reported in milk.

168 (3) The last step was to adjust the length of the baseline of each cycle from 3 to 8 days (Friggens
169 and Chagunda, 2005; Blavy et al., 2016).

170 This procedure was repeated for each of the 731 cycles in the simulated dataset. An example of a (de-
171)scaled milk P4 profile is shown in black in the upper panel of Figure 1. For each simulated cycle,
172 the reference moment of luteolysis was defined as the time (days in milk) at which the P4 level
173 decreased below 70% of the difference between maximum and minimum P4 concentration within
174 that cycle, further referred to as REF_{LUT} . This moment of REF_{LUT} was chosen based on the

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175 assumption that the P4 concentration has to be sufficiently low before the dominant follicle can
176 become LH sensitive. However occasionally, high-P4 estruses exist and thus full clearance of the P4
177 is not required (Friggens et al., 2008).

178 The second step entailed the addition of measurement noise corresponding to the on-line LFIA
179 technique (resulting profile shown in lower panel of Figure 1). The characteristics of this
180 measurement noise were defined from an available dataset containing 10,958 on-line measured P4
181 measurements collected at the Hooibeekhoeve in Geel, Belgium, which was described in (Adriaens
182 et al., 2019). The P4 data in this LFIA dataset was smoothed using a second order Savitzky-Golay
183 filter with a span of 7 measurements. Subsequently, the smoothed curve was subtracted from the data
184 and the residuals were sorted based on their smoothed value, in which 28 classes of 1 ng/mL were
185 created. For example, all measurements with a smoothed level between 0 and 1 ng/mL were assigned
186 to the first class. Next, the standard deviation of the data in each class was calculated, and a second
187 order polynomial was fitted to these standard deviation data, shown in Figure 2. This figure shows
188 that the on-line measured P4 data is heteroscedastic, with less variability at the extremes than at
189 intermediate values. The level of each simulated P4 measurement was obtained by multiplying the
190 scaled simulated value with 28 ng/mL, while the second order polynomial fitted on the residuals was
191 used to determine the standard deviation corresponding to this level. Next, a measurement was
192 sampled from a normal distribution $\sim N(\text{level}, \text{standard deviation})$. To introduce outliers in the dataset
193 caused by e.g. a sampling error, one outlier on average each 8 days of measurements was sampled
194 from a normal distribution with the maximum variance (i.e. a standard deviation of 5 ng/mL). This
195 outlier represents e.g. milkings in which the sampling did not represent the whole milking (e.g. failed
196 milkings or problems with the sampling unit). Accordingly, there was a chance of about 4% [= 1/
197 (8*3)] for each milking that the actual milk P4 value was replaced by an outlier sampled from a
198 normal distribution with the maximum variance. A fixed milking interval of 8 hours (i.e. 3 samples
199 per day) was chosen because the additional variability introduced by using realistic milking intervals

200 (6 – 20 hours) would not have different results (own unpublished data). An overview of the cycle
201 characteristics of the resulting dataset is given in the right part of Table 1.

202

203 *Progesterone-based monitoring algorithm using synergistic control*

204 The first algorithm tested in this paper, PMASC, consists of a mathematical model describing
205 the luteal dynamics (Adriaens et al., 2017), and a statistical process control chart to detect luteolysis
206 (Adriaens et al., 2018a). The mathematical model consists of two sigmoidal functions, a symmetrical
207 Hill function to characterize the increase in P4 during luteal development, and a Gompertz function
208 to describe the decrease during luteolysis. The control chart detects strong negative residuals from
209 the predicted luteal P4 concentration, while taking the variability in P4 measurements into account.
210 The decreasing function is only added after the detection of luteolysis, and can be used to calculate
211 model-derived indicators that can be employed to inform the farmer on relevant actions to be taken
212 (e.g. inseminations). It was shown before that the estimation of timing of luteolysis precedes the
213 preovulatory LH surge by about 55 to 65 hours (Adriaens et al., 2018b). In this study, two outputs of
214 PMASC were tested for their accuracy to relate to luteolysis. The first is the timing of the first
215 measurement detected to be out-of-control (**OO**C), i.e. the exact time in hours that a first drop in P4
216 lower than expected is detected, followed by the confirmation of luteolysis in the two successive
217 measurements. The second output '**TB85**' is an indicator calculated from the model as follows: the
218 moment that the model describing the drop in P4 during luteolysis (i.e. the Gompertz function)
219 decreases below 85% of the difference between maximum P4 model concentration minus the
220 baseline; both calculated from the current P4 cycle characteristics (Adriaens et al., 2018b).

221

222 *Multi-process Kalman filter*

223 To benchmark PMASC against the current state-of-the-art for on-line P4-based fertility
224 monitoring, the simulated P4 values were smoothed using a MPKF as described in Friggens and

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225 Chagunda, (2005); Friggens et al., (2008); Løvendahl and Chagunda, (2010). More specifically,
226 posterior mean estimates for the raw P4 values (i.e. the smoothed values) were calculated based on a
227 ‘mixture’ of 4 local linear dynamic models. These local linear models represent the 4 possible ‘states’
228 in which the P4 time-series can be: steady state, encountering a slope or a level change, or an outlier.
229 The mixture is calculated using the likelihood to be in a certain state taken from a predefined prior,
230 and the one-step-back and two-steps-back posterior probabilities. This means that the reaction of the
231 MPKF on a slope or level change increases with the extent of evidence for this state. For example,
232 when the P4 values rapidly decrease from luteal to follicular concentrations, a first low measurement
233 will be seen as very unlikely (‘outlier’), and the smoothed value will only decrease by a small amount.
234 However, when the next sample is low again, there is more ‘evidence’ that this is a slope change, and
235 an increased probability will be given to the ‘slope’ or ‘level’ change models, resulting in a larger
236 decrease in the smoothed value.

237 The framework for the MPKF was set up based on the information provided in Smith and West,
238 (1983), West and Harrison, (1997), Korsgaard and Løvendahl, (2002) and Friggens et al., (2007). The
239 parameters were estimated based on the raw and smoothed P4 values of the same P4 dataset described
240 before and in (Adriaens et al., 2019). The mean squared difference between our implementation of
241 the MPKF and the smoothed values obtained from the on-line device was 0.094 ng/mL which was
242 considered to be sufficiently low to compare results and derive relevant conclusions.

243 To provide a decision on when luteolysis has happened, the smoothed P4 values are combined
244 with a fixed threshold to extract information from the time series. The threshold currently considered
245 reasonable for estrus detection alerts based on milk P4 lies in between 4 and 6 ng/mL, and therefore
246 in this study was taken 5 ng/mL (Friggens et al., 2008; Bruinjé et al., 2017). The MPKF+T algorithm
247 gave an estrus alert when the smoothed value undercut the threshold value for the first time having
248 previously exceeded this threshold value. To avoid multiple alerts within the same follicular period,
249 a minimum time-interval of 5 days between two alerts was applied.

250 Once initiated and trained, this method does not need adjustments nor tuning to detect luteolysis.
 251 However, this user-friendly approach is at the cost of not including between-cow variation in
 252 responsiveness to P4. Moreover, the MPKF results in a time-lag between actual and detected
 253 luteolysis which is strongly dependent on luteolysis speed and the absolute P4 level measured
 254 (Friggens and Chagunda, 2005). A solution for this can be the implementation of a smart sampling
 255 scheme in which more samples are taken during the expected moment of luteolysis (e.g. from 16 days
 256 after the previous luteolysis on) and the calibration of the device to favor discrimination between high
 257 and low values.

258

259 ***Detection performance***

260 The simulated milk P4 dataset was analyzed using both algorithms (PMASC and MPKF+T) using
 261 2 different sampling schemes. In the first sampling scheme, a milk P4 measurement was taken at each
 262 simulated milking (i.e. 3 times a day, sampling scheme ‘ALL’). For the second, only 1 sample per
 263 day (a i.e. 24h interval between samples) was provided to the algorithms (sampling scheme ‘1D’),
 264 which corresponded to a sample or data reduction of 66%. The latter mimics the effect of missing
 265 samples during luteolysis or a sampling scheme in which only 1 sample is taken per day during
 266 luteolysis to minimize the analysis costs. This is still a very high sampling rate, especially during the
 267 growth phase of the CL which is less of interest. Nevertheless, it gives an indication of what the
 268 performance can be with a sampling rate of only 1 sample per day in the period of expected luteolysis,
 269 while previously described studies sample once per milking in that period (Bruinje et al., 2017). Next,
 270 the number and timing of simulated luteolyses was compared with the number of alerts given by each
 271 algorithm, and based hereon, the sensitivity, precision, and false negative rate (FNR) were calculated
 272 as follows (Eq. 1 to 3):

273
$$\text{Sensitivity} = \text{true positive rate} = \frac{TP}{TP + FN} \quad (\text{Eq. 1})$$

274 $Precision = positive\ predictive\ value = \frac{TP}{TP + FP}$ (Eq. 2)

275 $False\ Negative\ rate\ (FNR) = \frac{FN}{TP + FN}$ (Eq. 3)

276 With **TP** being the true positives, i.e. the times PMASC or MPKF+T gave a luteolysis alert within a
 277 window of 2 days before (i.e. 6 samples for the ALL, and 2 samples for the 1D sampling scheme) to
 278 4 days after (i.e. 12 samples for ALL and 4 samples for the 1D sampling scheme) a luteolysis was
 279 simulated (REF_{LUT}). This window was chosen based on its practical relevance: detection later than 4
 280 days would imply that ovulation had occurred by the time of detection, and the alert would be of no
 281 relevance for insemination (Roelofs, 2005; Adriaens et al., 2018b). False positives (**FP**) were alerts
 282 given by an algorithm which were not associated with a REF_{LUT}. The false negatives (**FN**) were
 283 defined as cases where no alert was given at all or it was given later than 4 days after REF_{LUT}. For
 284 this study, we chose not to include the specificity for a combination of reasons: (1) the specificity of
 285 the algorithms in this study is not fixed, unambiguously defined, but dependent on both the sampling
 286 rate and the window in which alerts are considered ‘true’ or ‘false’; (2) the specificity is not a measure
 287 where the farmer can do something with, as it is generally accepted that sensor systems should output
 288 as much as possible ‘farmers’ actions’, rather than raw data. Both concepts are further clarified in
 289 Hogeveen et al., (2010).

290

291 ***Timing of the Alerts and Consistency and Robustness against Missing Samples***

292 Besides sensitivity and accuracy measures, also the timing and variability in timing of the
 293 luteolysis alerts (i.e. P4 decrease) are of interest for an on-farm monitoring system. The first aspect,
 294 timing of the alert, is important because early alerts allow for correct planning of the insemination
 295 moment in order to achieve the highest chance of successful conception (Roelofs, 2005). The second
 296 aspect, variability in timing, is mainly of importance when a fixed insemination advice is coupled to
 297 the alerts. To this end, both the timing of the alerts and their variability compared to REF_{LUT} were

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298 evaluated for both sampling schemes. For PMASC, the first out-of-control measurement as well as
299 the model-based indicator TB85 were included. The latter is calculated as the moment the decreasing
300 Gompertz function reaches 85% of the difference between the between maximum and baseline P4
301 value of the model for that cycle, and thus takes the absolute milk P4 values of a cycle into account.
302 This was shown previously to be the most consistent model-based decision criterion in relation to the
303 LH surge (Adriaens et al., 2018b). Accordingly, 6 different groups were obtained, 3 for each sampling
304 scheme: (1) the difference between REF_{LUT} and the first out-of-control measurement detected by
305 PMASC, further referred to as OO_{ALL} and OO_{ID} respectively; (2) the difference between REF_{LUT}
306 and TB85, referred to as $TB85_{ALL}$ and $TB85_{ID}$; and (3) the difference between REF_{LUT} and the timing
307 of the milking that MPKF+T generated an alert, further indicated as $MPKF+T_{ALL}$ and $MPKF+T_{ID}$ for
308 each of the sampling schemes, respectively.

309 Because of the unequal variances between the 6 different groups, we could not perform normal
310 analysis of variance to compare their means. Therefore, the Wilcoxon Signed Rank Test was used to
311 assess differences in median timing of the alerts, and the Brown-Forsythe test was used to investigate
312 differences in variability. The latter tests the mean absolute difference from the median, which makes
313 it robust for non-normality. However, because multiple comparison tests are not available for these
314 statistical tests, it was decided to test each group against each other group in single pairwise
315 comparisons. To avoid capitalization of chance due to multiple tests, a Bonferroni-correction on the
316 significance level $\alpha=0.05$ was applied by dividing it by the number of tests run, being 15 (each of the
317 6 groups compared to the 5 other groups). Differences were thus considered significant if the p -value
318 was below $0.05/15 = 0.0033$.

319

320

RESULTS AND DISCUSSION

321 *Simulated data*

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322 The simulated dataset contained 150 milk P4 profiles and 731 individual cycles. Eighty-two
323 cycles had a prolonged follicular phase (≥ 9 days), 53 cycles had a prolonged luteal phase (≥ 17 days)
324 and 17 cycles had both. The average cycle length was 23.6 ± 7.6 days (mean \pm SD) when all cycles
325 were included, and 20.6 ± 1.9 days when cycles with prolonged phases were excluded. The average
326 length of the follicular phase was 4.5 ± 1.7 days and the average duration of the P4 drop during
327 luteolysis was 1.8 ± 0.7 days. The average baseline and maximum P4 concentrations were 2.1 ± 0.9
328 ng/mL and 18.8 ± 5.2 ng/mL, respectively. The characteristics of these profiles correspond to those
329 obtained in a real on-farm setting and are in line with reported characteristics in the literature (Blavy
330 et al., 2016). As such, the described methodology is a valuable way to generate large datasets while
331 avoiding measurement costs and controlling both fertility and cycle characteristics while having a
332 more precise reference for luteolysis. This study included in total 731 simulated luteolyses.
333 Technically, simulating many more (e.g. a million) cycles was possible, but because of computational
334 limitations, this current size was considered large enough number to show all possible difference
335 between the groups (MPKF/PMASC and the 2 sampling schemes).

336

337

338 ***Detection performance***

339 In Table 2, the detection performance statistics for PMASC and MPKF+T are summarized for
340 the different sampling schemes. When one measurement per milking was considered, PMASC and
341 MPKF+T both had high sensitivities of 99.2% and 98.6%, respectively. For PMASC, also the
342 precision of the alerts was high (95.3%), with only 4.7% of the 766 alerts being FP. With 23.0% FP
343 alerts, the MPKF+T algorithm was more sensitive to variations in the data. More specifically, these
344 false positive alerts can be classified as (1) outliers in the follicular phase (respectively 40 and 34%
345 of the FP for 1D and ALL); (2) coincidentally successive measurements below the threshold,
346 triggering the MPKF to low levels (respectively 43 and 49% of the FP for 1D and ALL); (3) cycles

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347 with intermediate maximal P4 concentration varying close to the threshold (17% of the FP for both
348 1D and ALL), which is often the case during e.g. luteal cysts (Yimer et al., 2018). To solve the first
349 problem, an additional requirement in the MPKF+T model was that P4 had been above 15 ng/mL
350 since the preceding estrus. The latter two situations are associated with the dependency of MPKF+T
351 on the absolute measured P4 values, and are therefore not easily solved from a detection perspective.
352 The MPKF+T of Friggens and Chagunda, (2005) gives an indication of the ‘goodness’ of the shape
353 of the preceding cycle as information to aid the farmer in deciding whether to inseminate, which
354 effectively flags these FPs (Friggins et al., 2008).

355 Reducing the sampling frequency to one sample per day decreases the sensitivity for estrus
356 detection of PMASC with 0.3% to 98.9%, and of MPKF+T with 2.2% to 96.4% (Table 2). The
357 number of false detections were halved to 2.6% and 12.7%, respectively for PMASC and MPKF+T
358 compared to the maximal sampling scheme, mainly because the number of outliers and coincidentally
359 successive low values decreased. Correspondingly, the precision of MPKF+T increased with 10.3%.
360 This shows the sensitivity of MPKF+T to the actual entered data and their absolute values. For
361 PMASC, the FNR was very similar to that for the ‘ALL’ sampling scheme (1.1% and 0.8%
362 respectively for ‘1D’ and ‘ALL’), which shows that the algorithm can work with less samples, as also
363 presented in Adriaens et al., (2019). The MPKF+T seems to be more sensitive to this as the FNR
364 increased from 1.4 to 3.6% when using less samples. The FN cases can be attributed to high-P4
365 estruses, in which samples close to the threshold were selected by coincidence. As a result, the
366 smoothed values do not cross the threshold, or do not cross it in time (within a window of 4 days after
367 simulated luteolysis). This phenomenon also shows the dependency of the MPKF+T on the actual
368 absolute values. For example, when the P4 concentration drops significantly, e.g. from 25 ng/mL to
369 4.9 ng/mL, the MPKF tends to be conservative and will not drop to a value lower than the thresholds.
370 Accordingly, at least 1 additional sample with a low P4 concentration has to be taken to trigger an
371 alert. This time lag thereby ensures that real outliers do not immediately trigger an alert, which is

372 important to guard the farmers' trust. The advantage of PMASC is that its control limits are
373 independent of the exact raw P4 values measured, and that the average value during the luteal phase
374 is indirectly taken into account to monitor the drop during luteolysis. A similar observation was made
375 by Friggens et al., (2008) when evaluating the luteolysis detection based on the model and algorithm
376 described in Friggens and Chagunda, (2005). Although the authors started with a fixed threshold of
377 4 ng/mL, they had to implement another threshold of 6 ng/mL in order to detect high P4 estruses
378 (Friggens et al., 2008). Because of the quite large measurement error for determining P4 in milk, it is
379 not yet known whether the differences in P4 concentrations during estrus (and during the luteal
380 phases) are due to inaccurate measurement and calibration methods, or due to real elevated P4
381 concentrations, e.g. due to increased P4 production by the adrenal cortex. When an improved P4
382 detection method becomes available, the inclusion of other information (e.g. health status, parity, fat
383 content of the milk) might become possible and might improve the detection algorithms, but to date
384 this is not yet available.

385 We did not test detection performance with other, more intelligent sampling schemes, because
386 this was considered outside the scope of this study. First of all, other regular sampling schemes (e.g.
387 1 sample per 2 days) are irrelevant as detection would likely be too late for timely insemination, even
388 if a follicular phase was detected. More specifically, PMASC was designed not just to 'detect', but
389 also to 'timely detect' estruses in order to increase the chance for successful insemination.

390

391 ***Timing of the Alerts: Variability and Robustness against Missing Samples***

392 In Figure 3, the results of the timing of the alerts using all samples and one sample per day
393 compared to REF_{LUT} are presented. The Wilcoxon signed rank test showed that not all medians are
394 equal ($p < 0.001$), which can also be seen in Figure 3. Using all samples, the median difference
395 between REF_{LUT} and the alerts generated by PMASC (TB85 and OOC) was close to 0, and their
396 individual comparison test had a p -value larger than 0.44. More concretely, a median of zero for the

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397 OOC group means that the first indication of luteolysis (i.e. first sample out-of-control) is obtained
398 almost simultaneously with the actual luteolysis REF_{LUT} . This ‘early’ first indication allows to
399 estimate the moment of luteolysis precisely, and has two large advantages: (1) it allows to organize
400 the insemination at the best moment; and (2) it allows to wait for additional proof of real estrus or
401 luteolysis in case of doubt, either by taking additional P4 samples, or by checking for external estrous
402 symptoms. The latter can be of added value for example when farmers are skeptical about the
403 reliability of new technologies. As expected, the most consistent insemination advice will be derived
404 from a model-based indicator as can be calculated from PMASC, which allows determination of the
405 moment of luteolysis independently of the sampling rate (when comparing 3 versus 1 sample per 24
406 hours). This paves the way to provide more consistent information to the farmer.

407 For the OOC_{ID} group (first out-of-control measurement detected by PMASC for the one-sample-
408 per day scheme), the first indication of estrus is obtained on average 10 hours later (significantly
409 different from all other groups, $p < 0.001$). However, using $TB85_{ID}$, (i.e. the model-based indicator
410 using the maximum and minimum model P4 value) estimating the actual moment of luteolysis in a
411 consistent way close to REF_{LUT} remains possible. The alerts of $MPKF+T_{ALL}$ and $MPKF+T_{ID}$ came
412 respectively on average 18 and 48 hours after REF_{LUT} , which also were significantly different from
413 all other medians ($p < 0.001$). Accordingly, the $MPKF+T$ algorithm has a time lag in detection of
414 about 2 milkings when no samples are missed during luteolysis, while this lag increases to
415 approximately 2 days when only one sample per day is taken. As a result, it becomes difficult to
416 provide the farmer with a reliable estimation of optimal insemination time, which is not only
417 dependent on the number of samples missed, but also on how fast P4 decreased (and thus, the $MPKF$
418 reacts), the moment of luteolysis compared to the timing of the milkings and milking intervals.

419 Table 3 shows that the Brown-Forsythe test for differences in variability in the interval alert to
420 REF_{LUT} was highly significant. The variability for $TB85_{ALL}$ was the smallest, with a standard
421 deviation respectively 35.3% and 15.8% smaller than OOC_{ALL} and $MPKF+T_{ALL}$. The OOC_{ALL} and

422 MPKF+T_{ALL} groups had an equal variability (p -value = 0.017 > 0.0033). Furthermore, individual
 423 comparisons pointed out that the variability does not differ between the OOC and TB85 group when
 424 one sample per day was considered (p -value = 0.047 > 0.0033). In contrast, we see that missing
 425 samples during luteolysis have a larger effect on the consistency of alerts given by the MPKF+T
 426 algorithm than on those obtained from OOC and TB85, resulting in a significantly larger variability
 427 than all other groups (p -value < 0.0033). We can therefore conclude that OOC and TB85 are less
 428 sensitive to missing samples both in terms of median timing of alerts and in terms of variability,
 429 making it a more robust algorithm for missing samples during luteolysis compared to MPKF+T.

430 Part of the variability within the groups is caused by the timing of the milkings relative to the P4
 431 profiles. For example, in this study a sample could have been taken in a window of 0 to 8 hours after
 432 REF_{LUT}. Accordingly, the effect of the timing of milking relative to REF_{LUT} seems to overrule the
 433 effect of the algorithm for the scheme in which milk P4 is measured every 8 hours. Although taking
 434 only one sample per day might not seem a ‘random’ way to mimic missing samples during luteolysis,
 435 the variability in luteolysis length and the independency of the P4 profiles to the simulated time of
 436 REF_{LUT} ensures that the timing of missed samples compared to luteolysis was variable. In a real on-
 437 farm setting, it is more probable that not in all cows samples are skipped, which would make the
 438 variability in alerts compared to real luteolysis for OOC and MPKF+T even larger (see also Adriaens
 439 et al., 2019). Furthermore, based on the results of this study, we can assume that the TB85 indicator
 440 remains consistent in its estimation of the timing luteolysis, independent of the sampling scheme or
 441 interval, which supports its use for monitoring purposes. A robust indicator is important when a fixed
 442 rule is used for decision making.

443 **CONCLUSIONS**

444 The detection performance of PMASC and the MPKF+T in terms of sensitivity, precision, false
 445 negative rate, false detection rate, robustness for sampling frequency and missing samples during
 446 luteolysis were studied on a simulated dataset of 150 highly variable P4 profiles containing 731

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447 luteolyses preceding estrus. Both PMASC and MPKF+T had a high luteolysis detection rate, but
448 MPKF+T was more sensitive to the absolute values of the P4 data, shown by its higher false detection
449 rate. This illustrates the value and limitations of both algorithms for on-line fertility monitoring. Using
450 a PMASC-based model indicator taking into account the luteal and follicular P4 concentrations, a
451 more robust estimation of the timing of luteolysis was obtained, which was less sensitive to missing
452 values compared to the current state-of-the-art MPKF+T both in terms of detection rate and
453 variability. Accordingly, PMASC has shown its potential to improve consistency and robustness of
454 progesterone-based, cost-effective detection of luteolysis on farm.

455

456

ACKNOWLEDGEMENTS

457 This work was supported by the Institute for the Promotion of Innovation through Science and
458 Technology in Flanders, Belgium (IWT) [IWT-LA project 110770]. Ines Adriaens and Ben Aernouts
459 are supported by the Fund for Scientific Research (FWO) Flanders, respectively grant number
460 11ZG916N and 12K3916N. Ines Adriaens obtained additional funding to perform a research stay at
461 INRA, grant V410318N. The RFM model was developed within the ‘Project Pluridisciplinary study
462 for a RObust and sustainabLe Improvement of Fertility In Cows’, PROLIFIC, EU grant no.
463 FP7.KBBE.2012.6, grant agreement no. 311776.

464

465

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550

551 **Figure 1** Upper panel: example of a simulated progesterone (P4) profile before (RFM, red) and after
 552 adding variability in profile characteristics (RFM_C, black). Lower panel: resulting simulated P4
 553 profile after descaling and adding measurement noise.

554

555 **Figure 2** Distribution of measurement noise for when progesterone (P4) is determined via an on-farm
 556 lateral flow immune assay (LFIA) device. This device is optimized to discriminate between high and
 557 low P4 values and accordingly, has a low standard deviation at the extremes. The blue bars represent
 558 the standard deviation per smoothed progesterone concentration bin (e.g. 1-2 ng/mL), calculated from
 559 a real on-farm measured dataset and smoothed using a second order Savitzky-Golay filter with a span
 560 of 7 measurements. The red thick line is the fitted second order polynomial fitted on these standard
 561 deviations, and used to determine the simulated measurement noise for each progesterone
 562 measurement.

563

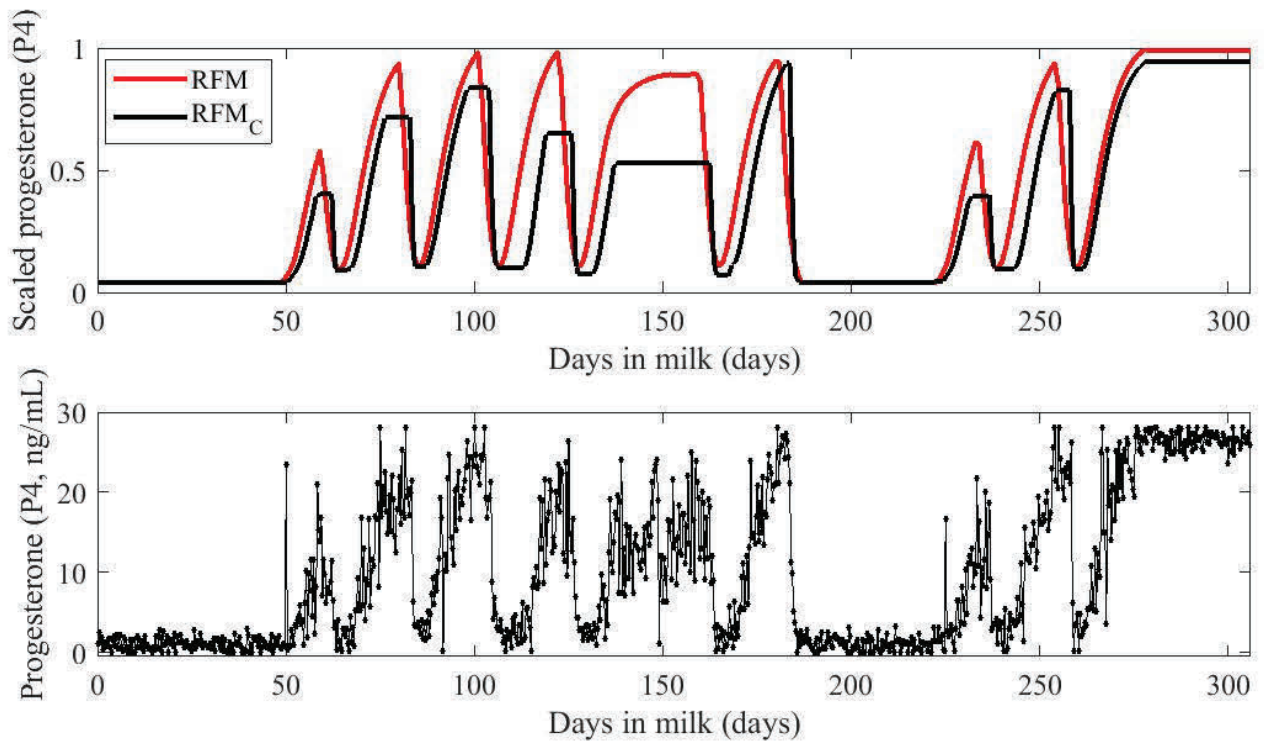
564 **Figure 3** Boxplots for the differences between alerts given by the 2 algorithms included in this study
 565 (progesterone monitoring algorithm using synergistic control, PMASC and a multiprocess Kalman
 566 filter plus threshold, MPKF+T) and the simulated luteolysis (reference timing of luteolysis, REF_{LUT}).
 567 A difference of zero means that luteolysis is detected on the moment it is simulated, which is seen for
 568 the first out-of-control sample of PMASC using all samples (first out-of-control, milk progesterone
 569 simulated 3 times a day, OOC_{ALL}) and for alerts based on the model-derived indicator TB85, even
 570 when samples during luteolysis are missed (TB85_{ALL} and TB85_{1D}). ‘1D’ hereby means that only 1
 571 sample per day was taken, also during the period in which luteolysis occurred.

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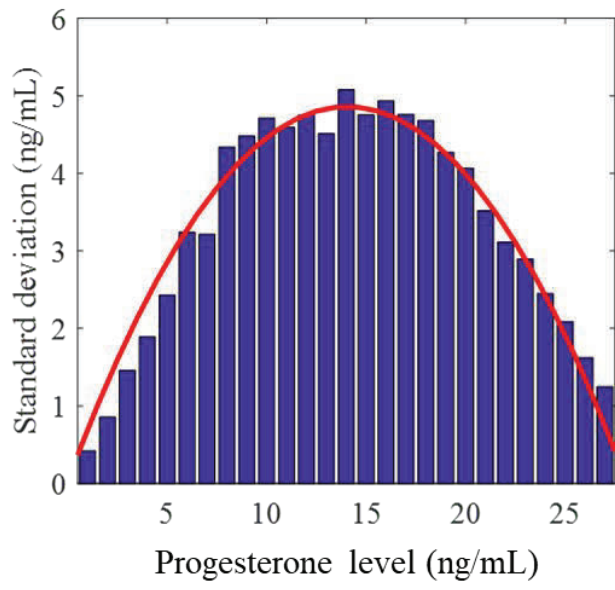
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575 Adriaens, Figure 1



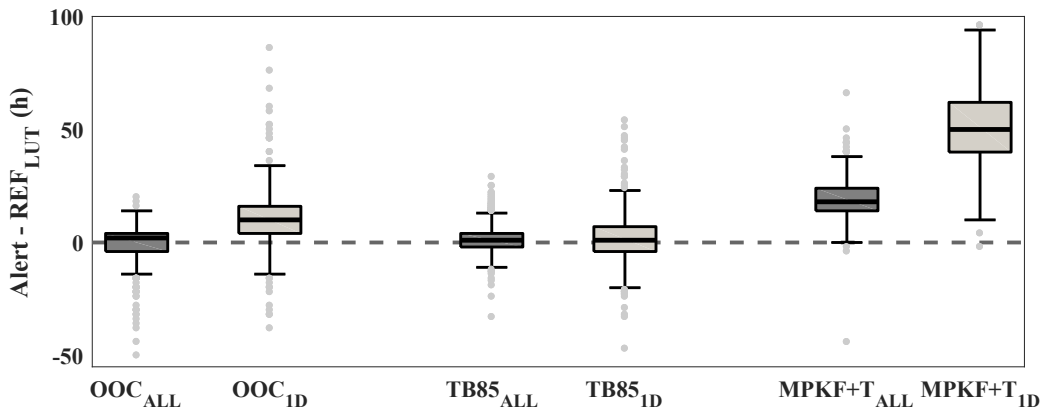
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577 Adriaens, Figure 2



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579 Adriaens, Figure 3



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581 Adriaens, Table 1. Progesterone (P4) profile and cycle characteristics of the simulated cows

	Profile characteristics				Total	Cycle characteristics			
	Min ¹	Max ²	Mean	SD ³		Min ¹	Max ²	Mean	SD ³
Length postpartum anestrus	14.9	59.9	29.9	6.2					
No. of cycles	2	9	4.9	1.3	731				
No. of P4⁴ measurements	281	843	518	106	77707				
No. of prolonged phases⁵	0	4			169				
Cycle length⁶ (days)						15	56	20.6	1.9
Baseline length (days)						2	8	4.5	1.7
Luteolysis length (days)						0.5	3	1.9	0.7
P4⁴ concentration (ng/mL)						0.0	28.0	12.9	9.7

582 ¹ minimum

583 ² maximum

584 ³ standard deviation

585 ⁴ progesterone (P4)

586 ⁵ a prolonged cycle phase represents the occurrence of a luteal or follicular cyst, in which respectively the luteal or follicular phase of

587 cycle is prolonged

588 ⁶ cycles with prolonged phases excluded

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589 Adriaens, Table 2. Number of estruses detected by PMASC and MPKF+T, and the resulting
 590 sensitivity, precision and false negative rate (FNR) when the progesterone time series consisted of 3
 591 samples per day (ALL) or 1 sample per day ('1D')

	PMASC ⁸		MPKF+T ⁹	
	ALL ⁶	1D ⁷	ALL	1D
Simulated	731	731	731	731
P¹	761	742	937	808
TP²	725	723	721	704
FP³	36	19	216	104
FN⁴	6	8	10	27
Sensitivity (%)	99.2	98.9	98.6	96.3
Precision (%)	95.3	97.4	76.9	87.1
FNR⁵ (%)	0.8	1.1	1.4	3.7

592 ¹ P, positives. This is the total number of alerts for estrus
 593 ² TP, true positives. These are the alerts within the window of 2 days before and 4 days after a simulated luteolysis.
 594 ³ false positives, FP. These are the alerts not associated with a simulated luteolysis
 595 ⁴ False negatives, FN. These are simulated luteolyses not associated with an alert within the window of 2 days before and 4 days after
 596 a simulated luteolysis
 597 ⁵ FNR, false negative rate
 598 ⁶ ALL: dataset all contains all samples, corresponding to 1 measurement each 8 hours
 599 ⁷ dataset 1D contains one sample per day
 600 ⁸ PMASC: Progesterone Monitoring Algorithm using Synergistic Control
 601 ⁹ MPKF+T: algorithm based on smoothed progesterone values using a multiprocess Kalman filter going below a fixed threshold of 5
 602 ng/mL
 603

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604 Adriaens, Table 3. Overview of mean, standard deviation and the outcome of the Brown-Forsythe
 605 test for equal absolute deviations of the median for all comparisons for the different algorithms (the
 606 progesterone monitoring algorithm using synergistic control, PMSAC and the multiprocess Kalman
 607 filter + a fixed threshold, MPKF+T) and sampling schemes (ALL: three times a day sampling vs. 1D:
 608 once a day sampling).

	No. of luteolysis	Mean (h)	SD ¹ (h)
REF_{LUT}²-OOC_{ALL}³	686	-1.94	13.14
REF_{LUT}-OOC_{ID}⁴	686	9.41	15.78
REF_{LUT}-TB85_{ALL}⁵	686	-0.26	8.50
REF_{LUT}-TB85_{ID}⁶	686	-1.44	14.61
REF_{LUT}-MPKF+T_{ALL}⁷	686	17.20	10.10
REF_{LUT}-MPKF+T_{ID}⁸	686	48.32	18.03
Pooled	4116	11.90	13.75
Brown-Forsythe statistic	64.08		
Degrees of freedom⁹	5, 4110		
p-value	<0.001		

609 ¹ Standard deviation
 610 ² REF_{LUT}: the timing of the simulated moment of luteolysis for each cycle, taken as the reference for the comparisons
 611 ³ REF_{LUT}-OOC_{ALL}: the difference between REF_{LUT} and the timing of the first out-of-control sample detected by the progesterone
 612 monitoring algorithm using synergistic control, (PMSAC) when the ALL sampling scheme was used (1 sample per 8 hours, i.e. 3
 613 samples per day)
 614 ⁴ REF_{LUT}-OOC_{ID}: the difference between REF_{LUT} and the timing of the first out-of-control sample detected by PMSAC when the 1D
 615 sampling scheme was used (1 sample per day)
 616 ⁵ REF_{LUT}-TB85_{ALL}: the difference between REF_{LUT} and the timing of the moment when the progesterone model during the decreasing
 617 in progesterone during luteolysis as modeled by PMSAC goes below 85% of the difference between the progesterone baseline and
 618 maximal model value, in the case of the ALL sampling scheme (3 samples per day)
 619 ⁶ REF_{LUT}-TB85_{ID}: the difference between REF_{LUT} and the timing of the moment when the progesterone model during the decreasing
 620 in progesterone during luteolysis as modeled by PMSAC goes below 85% of the difference between the progesterone baseline and
 621 maximal model value, when using the 1D sampling scheme (1 sample per day)
 622 ⁷ REF_{LUT}-MPKF+T_{ALL}: the difference between REF_{LUT} and the timing when the smoothed progesterone value calculated using the
 623 multiprocess Kalman filter (MPKF) crosses the threshold value of 5 ng/mL when using the ALL sampling scheme
 624 ⁸ REF_{LUT}-MPKF+T_{ID}: the difference between REF_{LUT} and the timing when the smoothed progesterone value calculated using the
 625 multiprocess Kalman filter (MPKF) crosses the threshold value of 5 ng/mL when using the 1D sampling scheme
 626 ⁹ the Brown-Forsythe test statistic has an F-distribution with ‘the number of groups minus 1’ numerator degrees of freedom, and ‘pooled
 627 number of samples minus number of groups’ denominator degrees of freedom
 628

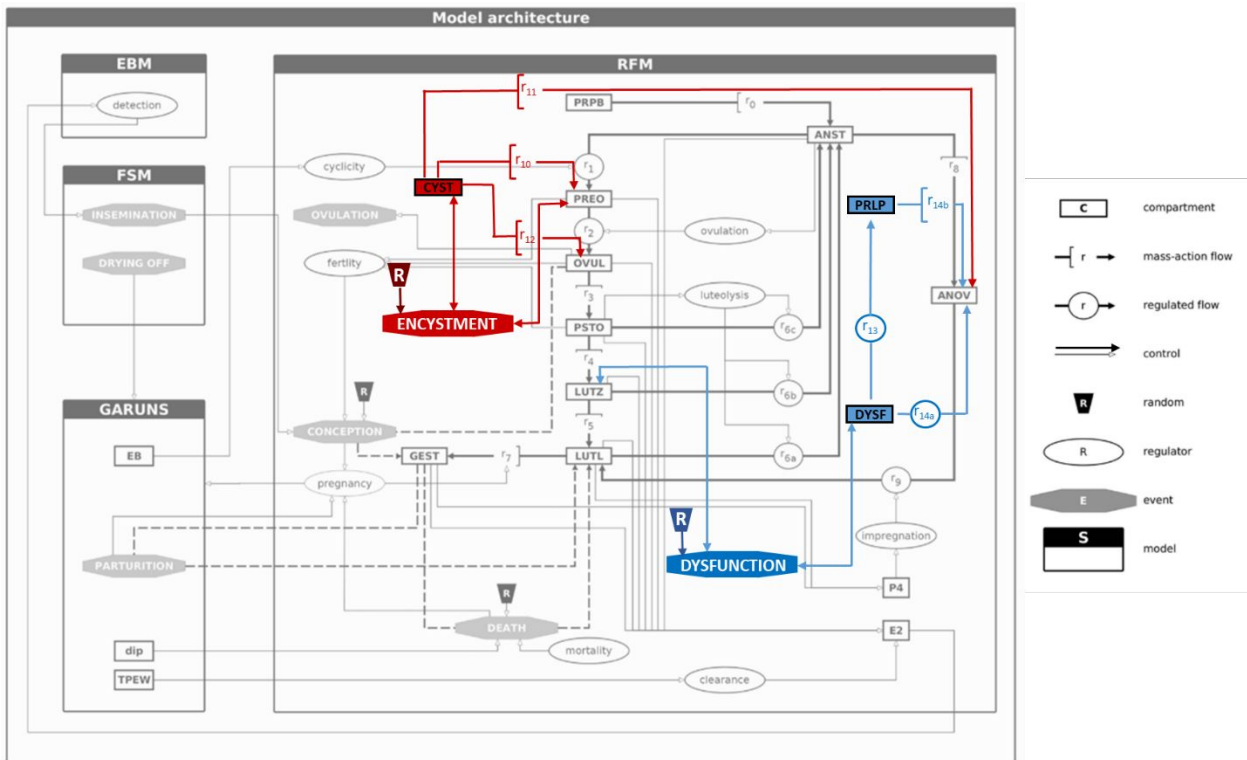
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APPENDIX

630

631 **Figure A1** Model structure of the Reproduction Function Model and the adaptations made to allow
632 for simulation of interrupted cyclicity by prolonged luteal (DYSF and PRLP, indicated in blue) and
633 follicular phases (CYST, indicated in red), adapted from Martin et al., (2018).

634 Adriaens, Figure A1



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636 Adriaens, Table A1. Parameters for the reproduction function model (RFM) model as described in
 637 Martin et al., (2018). The variability in genetic scaling parameters allows to alter individual
 638 performance of cows at the level of GARUNS. To allow for additional variability in fertility
 639 performance of the simulated cows, two sets of parameters (set 1 and set 2) were entered in the model.
 640 Using this table, new progesterone profiles can be simulated in a similar way as done in this study.

641

Parameter name RFM ¹	Set 1	Set 2	Effect
h_CYCL ²	-7.0	-15.0	Adjust postpartum anestrus length
Risk_CYST ³	1.6	2.0	Incidence of follicular cysts
Risk_DYSF ⁴	1.6	2.0	Incidence of luteal dysfunction
EMO ⁵	2.1	2.5	Incidence early embryonic mortality (conception failure)
GSP_cv ⁶	0.05	0.05	Variability between cows & characteristics (genetic scaling parameters)
RFM_h_impregnation[4] ⁷	0.10	0.10	Ensure restart of cyclicity after R_CYST
RFM_k09 ⁸	0.10	0.10	Ensure restart of cyclicity after R_CYST
RFM_k10 ⁹	0.12	0.12	Duration follicular cyst / ovarian anestrus
RFM_k11 ¹⁰	0.12	0.12	Duration follicular cyst / ovarian anestrus
RFM_k13 ¹¹	0.3	0.3	Duration luteal dysfunction
RFM_EMO[2] ¹²	4.6	4.6	Avoid occurrence late embryonic mortality
RFM_EMO[3] ¹³	8.0	8.0	Avoid occurrence fetal death

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¹ RFM: reproduction function model, the model used for simulating progesterone profiles in this study

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² Parameter determining the length of the postpartum anestrus phase. One of the 4 parameters varied for both simulations (set 1 and set 2)

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³ Parameter determining the incidence of follicular cysts. One of the 4 parameters varied for both simulations (set 1 and set 2)

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⁴ Parameter determining the incidence of luteal dysfunction. One of the 4 parameters varied for both simulations (set 1 and set 2)

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⁵ Parameter determining the incidence of early embryonic mortality. One of the 4 parameters varied for both simulations (set 1 and set 2)

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⁶ GSP_cv: coefficient of variation of the genetic scaling parameters. This parameter determines how much difference there is between simulated cows, and is indicative for the absolute values of some of the other outcomes (milk production, body weight) of the GARUNS model

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⁷ One of the 2 parameters that ensure cyclicity restarts in the model after an interruption due to a prolonged follicular period (low progesterone)

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⁸ One of the 2 parameters that ensure cyclicity restarts in the model after an interruption due to a prolonged follicular period (low progesterone)

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⁹ One of the 2 parameters of the RFM model influencing the duration of follicular problems, reflected in prolonged follicular phases of the progesterone cycle (low progesterone)

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¹⁰ One of the 2 parameters of the RFM model influencing the duration of follicular problems, reflected in prolonged follicular phases of the progesterone cycle (low progesterone)

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¹¹ Parameter influencing the duration of luteal dysfunction, reflected in prolonged luteal phases of the progesterone cycle (high progesterone)

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¹² Parameter determining when and how often late embryonic mortality occurs

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¹³ Parameter determining when and how often fetal death occurs

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