Risk measurement with equivalent utility principles

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Received: December 15, 2005; Accepted: March 16, 2006

Summary: Risk measures have been studied for several decades in the actuarial literature, where they appeared under the guise of premium calculation principles. Risk measures and properties that risk measures should satisfy have recently received considerable attention in the financial mathematics literature. Mathematically, a risk measure is a mapping from a class of random variables to the real line. Economically, a risk measure should capture the preferences of the decision-maker.

This paper complements the study initiated in Denuit, Dhaene & Van Wouwe (1999) and considers several theories for decision under uncertainty: the classical expected utility paradigm, Yaari's dual approach, maximin expected utility theory, Choquet expected utility theory and Quiggin's rank-dependent utility theory. Building on the actuarial equivalent utility pricing principle, broad classes of risk measures are generated, of which most classical risk measures appear to be particular cases. This approach shows that most risk measures studied recently in the financial mathematics literature disregard the utility concept (i.e., correspond to linear utilities), restricting their applicability. Some alternatives proposed in the literature are discussed.

1 Introduction and motivation

1.1 The need for risk measures

One of the major needs for risk measures is related to pricing in incomplete markets. In complete financial markets, hedging and arbitrage-free pricing are two sides of the same problem: the arbitrage-free price of a contingent claim equals the price of the hedging (in this case: replicating) portfolio. When the market is incomplete, however, there may be no replicating portfolio and the hedging strategy could involve a risky position. Moreover, there exists an infinite number of martingale measures and each of them generates a different price for the contingent claim. This leads to a set of possible prices. Now, since the hedging of the contingent claim involves risk, the price of the hedging portfolio depends explicitly on the agent's attitude towards risk.

AMS 2000 subject classification: Primary: 91B06, 91B30; Secondary: 62P05

Key words and phrases: risk measures, theories for decision under uncertainty, axiomatic characterization, equivalent utility, risk aversion

Note that the problem of market incompleteness is particularly relevant in insurance. This is due to several reasons: the stochastic nature of insurance processes (with jumps), the fact that many insurance risks are not actively traded in the (financial) market and the fact that securitized insurance risks often have an underlying index rather than an underlying traded asset.

Recent years have witnessed the emergence in the financial mathematics literature of a sophisticated theory of risk measures. However, risk measures have been extensively studied in the actuarial literature for more than 30 years, under the guise of premium calculation principles; see, e.g., Bühlmann (1970) and Goovaerts, De Vijlder & Haezendonck (1984). A risk measure is defined as a mapping from a class of random variables defined on a measurable space (representing the risks at hand) to the real numbers.

1.2 How to define a risk measure?

There are basically two ways to define a particular risk measure. In both cases, a set of axioms is given, but the axioms either describe how the risk is measured or describe how individuals make their decisions under risk or uncertainty. More specifically, this means that we can either axiomatize a risk measure directly by imposing axioms (e.g., with respect to ordering and aggregation, such as those discussed in Section 1.3 below) on random variables, or we can axiomatize a functional representing preferences (which in fact is also a risk measure) and use economic indifference arguments to obtain risk measures. The axioms in (actuarial and financial) mathematics are just phrased differently from the axioms in economics. In (actuarial and financial) mathematics we impose axioms on random variables (risks) whereas in economics we impose axioms on acts. That the two approaches coincide is explained nicely by Föllmer & Schied (2004). Let us now briefly explain these two approaches.

First, a set of axioms for risk measurement can be given that is regarded as reasonable in the situation under consideration. Then, the form of the risk measure is deduced from the agreed axioms. Risk measures are appropriate if, and only if, their characterizing axioms are. Axiomatizations can then be used to justify a risk measure but also to criticize it.

Second, a paradigm for decision under uncertainty can be selected, for instance the Von Neumann–Morgenstern expected utility theory, Yaari's dual utility theory or other generalized utility principles, for instance. Note that this selection also amounts to opting for a set of axioms, but this set explains how decision-makers choose between uncertain prospects; see, e.g., Wakker (2004) for more details. The risk measure is then obtained by an equivalence principle: the decision-maker is indifferent between the cash-flow $\varrho[X]-X$ (corresponding to the case where X is covered) and 0 (no coverage).

Both approaches for measuring risk will be considered in this paper. It will be seen that many classical risk measures can be obtained in both ways, which allows for a better understanding of their intrinsic properties. This will also enlighten the fact that most risk measures studied recently in the financial mathematics literature disregard utilities (considering only linear utility functions). This restriction explains some problems, like insensitivity to liquidity risk, for instance.

1.3 The axiomatic approach to risk measurement

Axioms to characterize a risk measure can generally be divided into three classes: rationality axioms, additivity axioms and technical axioms. First there are the basic rationality axioms such as monotonicity, which are satisfied by most risk measures. Additivity axioms deal with sums of risks. They describe the sensitivity of the risk measure with respect to risk aggregation. Let us mention the additivity for independent random variables, the additivity for co-monotonic random variables, the subadditivity or the superadditivity. These additivity requirements often are the most characteristic for the corresponding risk measure. Finally, there are technical requirements, mostly continuity conditions. They are usually necessary for obtaining mathematical proofs and are typically difficult to validate or to explain economically.

Let us now detail some of the properties that risk measures may or may not satisfy:

Objectivity The risk measure ϱ does not depend on the risk itself but only on its underlying distribution, i.e. $\varrho[X] = \varrho[F_X]$ where F_X is the distribution function of X.

The objectivity requirement implies that risk measurement can be entirely based on distribution functions. This is usually the case in actuarial science. This condition ensures that F_X contains all the information needed to measure the riskiness of X. This property is sometimes called "law invariance", and phrased as $X =_d Y \Rightarrow \varrho[X] = \varrho[Y]$ where $=_d$ denotes the equality in distribution.

Two monotonicity requirements can then be envisaged: either we impose that the risk measure ϱ agrees with almost sure comparisons (weak monotonicity) or with stochastic dominance \leq_{st} (strong monotonicity).

Weak monotonicity For any *X* and *Y*, $P[X \le Y] = 1$ implies $\varrho[X] \le \varrho[Y]$.

Strong monotonicity For any X and Y such that $X \leq_{\text{st}} Y$, i.e. $P[X > t] \leq P[Y > t]$ for all $t \in \mathbb{R}$, implies $\varrho[X] \leq \varrho[Y]$.

Co-monotonicity is a useful notion in decision theory; see, for example, the review papers by Dhaene, Denuit, Goovaerts, Kaas & Vyncke (2002a,b) and the references therein. Recall that the random variables X and Y are said to be co-monotonic if they satisfy

$$P[X \le x, Y \le y] = \min \{P[X \le x], P[Y \le y]\} \text{ for all } x, y \in \mathbb{R}.$$

For two co-monotonic random variables X and Y, we have that $X \leq_{\text{st}} Y \Leftrightarrow P[X \leq Y] = 1$. In the absence of co-monotonicity, weak monotonicity will not imply strong monotonicity. However, if ϱ is objective, then both requirements have the same force. We refer the reader to Bäuerle & Müller (2006) for a careful examination of objectivity and monotonicity properties.

Positive homogeneity For any a > 0, $\varrho[aX] = a\varrho[X]$.

In the literature it is often stated that positive homogeneity is equivalent with currency independency. This is however a wrong interpretation. Take as an example $\varrho[X] = \mathbb{E}[(X-d)_+]$ with $\xi_+ = \max\{\xi, 0\}$, which is clearly currency independent (changing X)

from dollars to euros only makes sense if also d is changed in that way), but not positively homogeneous. Hence, positive homogeneity is about multiplying all euro amounts of the claims by a constant factor. Let us also mention that future currencies are random so that the constant a involved in the definition of positive homogeneity cannot be used to convert X from euros to dollars, say. It is rather related to the "volume" of the risk and has a close relation with additivity for co-monotonic risks (see below).

Translation equivariance For any $b \in \mathbb{R}$, $\varrho[X + b] = \varrho[X] + b$.

Translation equivariance might be considered desirable for a risk measure ϱ used to calculate provisions or premiums. Indeed it seems to be reasonable when the risk contains a certain amount b that $\rho[b+X] = b + \rho[X]$ in both cases. This is of course not the case for risk measures used for the calculation of solvency (regulatory or economic) capital. These quantities constitute amounts of money for safety, in addition to the provisions and premiums. Such risk measures are of the form $\widetilde{\varrho}[X] = \varrho'[(X - \varrho[X])_+]$ for some appropriate risk measures ϱ' and ϱ . Consequently, if ϱ is translation equivariant then $\tilde{\varrho}$ satisfies $\widetilde{\rho}[X+b] = \widetilde{\rho}[X]$, so that we have translation invariance in this case. It is of course not forbidden to take translation equivariance as an axiom for the construction of risk measures in this situation, but it cannot be interpreted in the economic concept of solvency. The prescribed rules in the framework of Solvency 2 for the calculation of solvency capital (the difference between the Value-at-Risk at level 99.5% and the Value-at-Risk at level 75%) recognizes this reality. Related of course is the requirement $\rho[b] = b$ which can be defended in the case of insurance pricing (to avoid free lunches and/or the no-ripoff condition) but in the definition of a risk measure $\rho[b]$ might be equal to a function u(b), the utility.

Combining positive homogeneity with translation equivariance guarantees that ϱ is linear, that is, whatever a > 0, $\varrho[aX + b] = a\varrho[X] + b$.

Subadditivity $\varrho[X+Y] \leq \varrho[X] + \varrho[Y]$ whatever the dependency structure of X and Y.

The rationale behind subadditivity can be summarized as "a merger does not create extra risk". Subadditivity reflects the idea that risk can be reduced by diversification. According to Rootzen & Klüppelberg (1999), subadditivity is a convenient mathematical property that does not hold in reality. The behavior of a risk measure with respect to the aggregation of risks is manifested by the award of diversification discounts and the imposition of penalties.

When equality holds, we speak of additivity. In this case, the dependence structure between X and Y is often specified: for example, additivity for independent risks or additivity for co-monotonic risks (see below). We refer the interested reader to Dhaene, Laeven, Vanduffel, Darkiewicz & Goovaerts (2004) for a detailed discussion on the desirability of the subadditivity axiom when setting solvency capital requirements.

In Artzner, Delbaen, Eber & Heath (1999), a risk measure that satisfies the properties of monotonicity, positive homogeneity, translation equivariance and subadditivity is called a "coherent" risk measure. As explained below, coherent risk measures coincide with

upper expectations introduced by Huber (1981) in robust statistics. Coherent risk measures satisfying the objectivity requirement have been studied by Kusuoka (2001); see also Inoue (2003) and Bäuerle & Müller (2006) for an extension to unbounded random variables. The terminology "coherent" is somewhat flawed in the sense that it may suggest that any risk measure that is not "coherent" is always inadequate. It is worth to mention that coherency is defined with respect to a set of axioms, and no set is universally accepted. Modifying the set of axioms regarded as desirable leads to other risk measures that may be called "coherent" with just as much reason.

Convexity For any $\lambda \in [0, 1]$, $\varrho[\lambda X + (1 - \lambda)Y] \le \lambda \varrho[X] + (1 - \lambda)\varrho[Y]$ whatever the dependency structure of X and Y.

Föllmer & Schied (2002), and independently Frittelli & Rosazza Gianin (2002), introduced the concept of convex risk measures, which satisfy the properties of monotonicity, translation equivariance and convexity. It has been further developed by Krätschmer (2005). The class of coherent risk measures can be characterized as the class of convex risk measures that satisfy the positive homogeneity property. As the class of convex risk measures is an extension of the class of coherent risk measures, it is sometimes called the class of weakly coherent risk measures. It is worth to mention that the work by Deprez & Gerber (1985) already contains many nice results on convex risk measures.

Additivity for co-monotonic risks $\varrho[X+Y] = \varrho[X] + \varrho[Y]$ for co-monotonic random variables X and Y.

Co-monotonic risks are bets on the same event and in this sense neither of them provides a hedge against the other. Because of the no-hedge condition, no reduction in the risk-load is awarded for a combined policy, resulting in co-monotonic additivity.

Additivity for independent risks $\varrho[X + Y] = \varrho[X] + \varrho[Y]$ for independent random variables *X* and *Y*.

Additivity for independent risks may be a reasonable requirement for a premium calculation principle, since it ensures that the policyholder has no interest in splitting the risk in independent components asking for coverage to several insurers. It is also justified by premium computation from top to down; see, e.g., Borch (1962) or Bühlmann (1985). The most general representation of risk measures that are additive for independent random variables is due to Gerber & Goovaerts (1981) and is known as the mixed Esscher principle. Despite the numerous appealing features of the Esscher principle, it does not satisfy the strong monotonicity requirement. Counterexamples have been provided by Van Heerwaarden, Kaas & Goovaerts (1989). This is why Goovaerts, Kaas, Laeven & Tang (2004) provided a new axiomatic characterization of risk measures that are additive for independent random variables, involving an axiom that guarantees monotonicity. The obtained risk measure is a restricted version of the mixed Esscher principle that can be regarded as an ordinary mixture of exponential premiums.

1.4 Risk measures and economic decision under uncertainty

Goovaerts, Kaas, Dhaene & Tang (2004) gave several examples where evidently the properties that risk measures should have are determined by the realities of the actuarial applications (namely, insurance-reinsurance, premium calculation, premium calculation from top to down, capital allocation, solvency margin and setting provisions). See also Dhaene, Vanduffel, Tang, Goovaerts, Kaas & Vyncke (2004).

Instead of stating a set of axioms that risk measures should satisfy and then determining mathematically the corresponding functional representation, one could determine the functional form of the risk measure via economic indifference arguments. For instance, the equivalent utility premium principle defined by Bühlmann (1970) is obtained by considering an insurer whose preferences are characterized by expected utility (EU, in short). The distortion premium principle proposed by Denneberg (1990) and Wang (1996) can be obtained via indifference arguments based on distortion utility (DU, in short). The distortion-exponential premium principle has been similarly derived by Tsanakas & Desli (2003) and Heilpern (2003) using rank-dependent expected utility (RDEU, in short) theory. See also Luan (2001). The present paper follows Denuit, Dhaene & Van Wouwe (1999) where only EU and DU were considered (with a brief excursion to RDEU).

Utility theories are typically presented in an axiomatic form. The axioms are intended as reasonable consistency requirements: the concept of rationality is embodied in the axioms, with the possible addition of the postulate of risk aversion. The difference is that here, axioms impose rational behavior in decision under uncertainty. As it was the case for the sets of axioms defining risk measures, the preference foundations typically use intuitive axioms such as transitivity, which refer to directly meaningful empirical properties of preferences, and technical axioms such as continuity, which describe the structural richness of the model and serve to simplify the mathematical analysis.

Fishburn (1982) and Yaari (1987) considered the following four axioms on the preference relation ≤:

Objectivity If X and Y have the same distribution function then the decision-maker is indifferent between X and Y, that is, $X \leq Y$ and $Y \leq X$ both hold.

Weak order \leq is complete, transitive and reflexive.

Strong monotonicity If $X \leq_{\text{st}} Y$ then $X \leq Y$.

Continuity The set $\{Y|Y \succeq X\}$ is closed for every X, the closure being often meant with respect to the topology of weak convergence, associated with the L^1 norm.

Then, a fifth axiom is added that specifies the particular form of the indicator of preferences. Whereas there is a rather large consensus about the four axioms listed above, the last one continues to be subject to debate. The EU axiomatization hinges on the independence axiom, which requires preferences to be separable across mutually exclusive events. Alternative theories proposed by Quiggin (1982), Yaari (1987) and Schmeidler (1989) replaced this condition by a type of co-monotonic independence axiom, an intuitive and appealing condition requiring that the usual independence axiom only holds when hedging effects are absent.

In this paper, we will consider in addition to the classical EU and DU theories several alternative paradigms for decision under uncertainty: the Choquet expected utility (CEU, in short) theory proposed by Schmeidler (1989), the maximin expected utility (MEU, in short) theory proposed by Gilboa & Schmeidler (1989) and the rank-dependent expected utility (RDEU, in short) theory proposed by Quiggin (1982). The application of the actuarial equivalent utility pricing principle in these paradigms will generate most classes of risk measures considered so far in the financial mathematics literature.

1.5 Risk and uncertainty

Decision-making under risk and decision-making under uncertainty are differentiated according to the condition that in the former case the probabilities of events are given while in the latter case they are not. With regard to decision-making under uncertainty, models have been introduced which associate some kind of probabilities to events.

In the financial mathematics literature, and the actuarial papers in the realm of financial mathematics, the probability distribution of a risk does not need to be fixed (a risk is just a random variable defined on a measurable space, and the probability measure can be left unspecified). The majority of actuarial papers uses the framework of probability theory and statistics, in which a random variable is defined on a particular and fixed underlying probability space, not on just a measurable space with many possible probability measures. The probability measure may be unknown, but it is fixed and unique. This does not prevent that many weight functions operating on events can be used to calculate the premiums. For instance, the actuarial practice in life insurance consists in replacing the distribution of the remaining lifetime by a "more dangerous one" in order to generate a safety loading. We will also see below that in economics terminology, for example, a coherent risk measure is not a measure of risk but a measure of uncertainty.

Uncertainty is often represented with the help of capacities. Specifically, we assume that the uncertainty a decision maker faces can be described by a non-empty set of states, denoted as Ω . This set may be finite or infinite. Associated with Ω is the set of events, taken to be a sigma-algebra of subsets of Ω , denoted by \mathcal{B} . Capacities are real-valued functions defined on \mathcal{B} that generalize the notion of probability distributions. Formally, a capacity is a normalized monotonic set function. More precisely, $C:\mathcal{B}\to [0,1]$ is a capacity if $C[\emptyset]=0$, $C[\Omega]=1$ and for any $A,B\in\mathcal{B},A\subseteq B\Rightarrow C[A]\leq C[B]$. Otherwise, the capacity has little structure. In particular, the sum of the capacities of two disjoint subsets may exceed 1. Or the sum of the capacities of two subsets may be strictly less than the capacity of the union of these sets. The capacity C is called coherent if there exists a non-empty set \mathcal{M} of probability measures $P:\mathcal{B}\to [0,1]$ such that

$$C[A] = \sup_{P \in \mathcal{M}} P[A]$$
 for all $A \in \mathcal{B}$.

As we will see, coherent capacities are closely related to coherent risk measures.

To each capacity C we can associate a dual capacity. The dual capacity of C, denoted as \overline{C} , is defined by

$$\overline{C}[A] = 1 - C[\Omega \setminus A]$$
 for all $A \in \mathcal{B}$.

The dual capacity is useful when losses are considered instead of gains. Given a coherent capacity C, a natural and widely used measure of the degree of uncertainty (or imprecision, vagueness, ambiguity) associated with an event $A \in \mathcal{B}$ is given by the interval

$$\left[\overline{C}[A], C[A]\right] = \left[\inf_{P \in \mathcal{M}} P[A], \sup_{P \in \mathcal{M}} P[A]\right]. \tag{1.1}$$

The interval (1.1) captures uncertainty in the sense of Knight (1921). If $\overline{C}[A] = C[A]$, the event A does not involve any vagueness and all priors $P \in \mathcal{M}$ agree on A.

Decisions under risk are usually performed in the EU or RDEU paradigms. In situations of uncertainty (i.e., in situations where there does not exist a given objective probability distribution available to the decision-maker), these paradigms are replaced by the subjective utility theory and the Choquet expected utility (CEU, in short), respectively.

1.6 Aversion to risk and uncertainty

There are many notions of risk aversion. Chateauneuf & Cohen (2000) distinguish five kinds of risk aversion: weak, monotonic, left monotonic, right monotonic and strong. In this paper, we restrict ourselves to the following two standard concepts of risk aversion. The first one defines an individual to be risk averse if the sure gain $\mathbb{E}[X]$ is always preferred to the random prospect X itself. An alternative definition of risk aversion requires a risk averse individual to behave according to the concave order, denoted as \leq_{cv} , and defined by $X \leq_{cv} Y \Leftrightarrow \mathbb{E}[\phi(X)] \leq \mathbb{E}[\phi(Y)]$ for all concave functions $\phi : \mathbb{R} \to \mathbb{R}$, provided the expectations exist. This order extends the concept of mean preserving increase in risk.

We thus have a concept of weak risk aversion ($\mathbb{E}[X]$ is preferred over X) and a concept of strong risk aversion ($X \leq_{cv} Y \Rightarrow X$ is preferred over Y). It is interesting to note that strong risk aversion is defined with the help of a stochastic order relation. Relations like the dispersive order have been used to define alternative attitudes towards risk; see, e.g., Chateauneuf & Cohen (2000) and Chateauneuf, Cohen & Meilijson (2004).

1.7 Outline of the paper

This paper aims to demonstrate that, as pointed out by Geman (1999), in the fundamental subject of financial risk analysis, some valuable lessons may be drawn from insurance. Actuaries have used indifference arguments based on economic theories for decision under uncertainty for decades, after Bühlmann (1970) introduced the equivalent utility principle in the context of setting insurance premiums. Considering this simple and appealing approach to risk measurement, most risk measures can be obtained in a way that enlightens their intrinsic properties. Unfortunately, apart from some authors including Föllmer & Schied (2004), the close connections between risk measurement and economic decision theory have been disregarded by financial mathematicians, who independently developed their own axiomatic theories. This paper tries to reconcile the two approaches.

2 Equivalent utility risk measures in EU theory

2.1 EU theory

The classical expected utility (EU, in short) theory is identified by the independence axiom. Every EU decision-maker possesses some utility function $u(\cdot)$ and acts in order to maximize the expected utility. The prospect X is then evaluated by

$$\mathbb{U}_{EU}[X|u] = \int u(X) dP = \mathbb{E}[u(X)]$$

and X is preferred over Y if $\mathbb{U}_{EU}[X|u] \geq \mathbb{U}_{EU}[Y|u]$. A classical reference for EU is Von Neumann & Morgenstern (1947). The subjective utility theory of Savage (1954) for decision under uncertainty differs from EU only by its interpretation of the underlying probability distribution P.

If the decision-maker, with initial wealth w and utility function $u(\cdot)$, has to choose between random losses X and Y, then he compares $\mathbb{U}_{EU}[w-X|u]$ with $\mathbb{U}_{EU}[w-Y|u]$ and chooses the loss which gives rise to the highest expected utility.

2.2 Risk aversion

The two definitions of risk aversion are equivalent under EU theory, where they both coincide with the concavity of the utility function: for an EU decision-maker, weak risk aversion implies automatically strong risk aversion.

Note that in the EU framework, the agent's attitude towards risk and the agent's attitude towards wealth are forever bonded together (since they are both derived from the characteristics of u): risk aversion and diminishing marginal utility of wealth are synonymous. Nevertheless, risk aversion expresses an attitude towards risk while decreasing marginal utility expresses an attitude towards wealth. Thus, in EU theory, the concept of risk aversion being encapsulated in the utility function is a property of attitudes towards wealth rather than of attitudes towards risk.

2.3 Equivalent EU risk measures

Consider a decision-maker with initial wealth w and with a utility function $u(\cdot)$. Facing a random loss X, he sets its price for coverage $\varrho[X]$ as the solution of the equation

$$\mathbb{U}_{EU}[w + \varrho[X] - X|u] = \mathbb{U}_{EU}[w|u] = u(w). \tag{2.1}$$

Condition (2.1) expresses that $\varrho[X]$ is fair in terms of utility: the right-hand side of (2.1) represents the utility of not not covering X; the left-hand side of (2.1) represents the expected utility of the decision-maker assuming the random financial loss X. Therefore (2.1) means that, provided an amount of $\varrho[X]$ is obtained when X is covered, the expected utility of wealth with X is equal to the utility without X. Putting w=0 and normalizing the utility function u, we get the so-called "equivalent utility principle": $\varrho[X]$ calculated according to this principle is the root of the equation

$$\mathbb{U}_{EU}[\varrho[X] - X|u] = 0 \tag{2.2}$$

which can be interpreted as an equality between the expected utility of the income $\varrho[X] - X$ and the utility of not accepting the risk.

If we assume that the moment generating function of X exists and that the utility function of the insurance company is of the form

$$u(x) = \frac{1 - \exp(-cx)}{c}, \quad x \ge 0,$$
 (2.3)

for some positive constant c, then (2.2) admits an explicit solution and $\varrho[X]$ can be expressed as

$$\varrho[X] = \frac{1}{c} \ln \mathbb{E}[\exp(cX)]. \tag{2.4}$$

Note that the exponential principle (2.4) also possesses an appealing interpretation in terms of ruin theory; see, e.g., Kaas, Goovaerts, Dhaene & Denuit (2001, Section 5.2) or Tsanakas & Desli (2003).

3 Equivalent utility risk measures in DU theory

3.1 DU theory

Under the same set of axioms but with a modified independence axiom, Yaari (1987) has shown that there must exist a "distortion function" f such that a prospect X is valued at its "distorted expectation" $\mathbb{U}_{DU}[X|f]$ defined as

$$\mathbb{U}_{DU}[X|f] = -\int_{x=-\infty}^{0} \left(1 - f(\overline{F}_X(x))\right) dx + \int_{x=0}^{+\infty} f(\overline{F}_X(x)) dx. \tag{3.1}$$

Here the distortion function $f:[0,1] \to [0,1]$ is non-decreasing with f(0)=0 and f(1)=1. Instead of using the tail probabilities $\overline{F}_X(x)$, the decision-maker uses the distorted tail probabilities $f(\overline{F}_X(x))$. Yaari's model has been further developed by Roëll (1987), and an alternative axiomatization is proposed in Guriev (2001). We also refer to Carlier & Dana (2003) for interesting representations of $\mathbb{U}_{DU}[X|f]$.

3.2 Risk aversion

Contrary to the EU case, the two concepts of risk aversion are no more equivalent in DU. For the weak risk aversion, it suffices that $f(p) \le p$, $p \in [0, 1]$, as we immediately find that

$$\mathbb{U}_{DU}[X|f] \le \mathbb{E}[X] = \mathbb{U}_{DU}[\mathbb{E}[X]|f]. \tag{3.2}$$

For the strong risk aversion, the distortion function f has to be convex.

3.3 Equivalent DU risk measures

Let us introduce the dual distortion function \overline{f} for each distortion function f:

$$\overline{f}(x) = 1 - f(1 - x), \qquad x \in [0, 1].$$
 (3.3)

The dual distortion function is again a distortion function. It is clear that $\overline{\overline{f}} \equiv f$ and that f convex $\Leftrightarrow \overline{f}$ concave. Furthermore, we have that $\mathbb{U}_{DU}[-X|f] = -\mathbb{U}_{DU}[X|\overline{f}]$.

The equivalent DU risks measures are obtained as the solution of the indifference equation

$$\mathbb{U}_{DU}[0|f] = \mathbb{U}_{DU}[\varrho[X] - X|f]. \tag{3.4}$$

Solving (3.4) gives

$$\varrho[X] = \mathbb{U}_{DU}[X|\overline{f}]$$

$$= -\int_{-\infty}^{0} \left(1 - \overline{f}(\overline{F}_X(x))\right) dx + \int_{0}^{\infty} \overline{f}(\overline{F}_X(x)) dx. \tag{3.5}$$

One immediately finds that $\overline{f}(\overline{F}_X(x))$ is a non-increasing function of x with values in the interval [0, 1]. The risk measure (3.5) can be interpreted as a "distorted expectation" of X, evaluated with a "distorted probability measure" in the sense of a Choquet-integral, see Denneberg (1994). Risk measures (3.5) are often called distortion risk measures. In the actuarial literature, distortion risk measures have been introduced by Denneberg (1990) and Wang (1996). Axiomatic characterizations of insurance prices leading to distorted premiums have been proposed in Goovaerts & Dhaene (1998) and Wang, Young & Panjer (1997).

Many risk measures can be cast into the form (3.5). Let us give some prominent examples.

Example 3.1 (Value-at-Risk) For any p in (0, 1), the p-quantile risk measure (or Value-at-Risk) for a random variable X with given distribution function F_X is defined by $F_X^{-1}(p) = \inf\{x \in \mathbb{R} \mid F_X(x) \geq p\}, p \in (0, 1)$. It can be shown that $F_X^{-1}(p)$ corresponds to (3.5) with distortion function

$$\overline{f}(q) = \begin{cases} 1 & \text{if } q > 1 - p, \\ 0 & \text{otherwise}, \end{cases} \quad \text{for } 0 \le q \le 1.$$

Example 3.2 (Tail-VaR) The Tail-VaR at level p, denoted by $TVaR_p[X]$, is defined by

$$TVaR_p[X] = \frac{1}{1-p} \int_p^1 F_X^{-1}(q) \, dq, \qquad p \in (0,1).$$
 (3.6)

It is the arithmetic average of the quantiles of X, from p on, and corresponds to (3.5) with distortion function

$$\overline{f}(x) = \min\left(\frac{x}{1-p}, 1\right), \qquad 0 \le x \le 1. \tag{3.7}$$

VaR's are the building blocks of the distortion measures (3.5). To see this, it is enough to replace $\overline{f}(\overline{F}_X(x))$ by $\int_0^{\overline{F}_X(x)} \mathrm{d}\overline{f}(q)$ in (3.5) and to revert the order of the integrations. One then finds that any distortion risk measure can be written as

$$\varrho[X] = \int_0^1 F_X^{-1} (1 - q) d\overline{f}(q). \tag{3.8}$$

Note that when the distortion function \overline{f} is differentiable, (3.8) can be rewritten as

$$\varrho[X] = E \left[F_X^{-1} (1 - U) \overline{f}'(U) \right]. \tag{3.9}$$

Formulas (3.8)–(3.9) show that the spectral risk measures of Acerbi (2002) are particular cases of distortion risk measures. It can be further shown that coherent distortion risk measures can be represented as mixtures of Tail-VaR's, as pointed out by Inui & Kijima (2005); see also Bassett, Koenker & Kordas (2004) for portfolio allocation with mixtures of Tail-VaR's, and Pflug (2002).

Distortion risk measures (3.5) with \overline{f} concave are subadditive. The VaR therefore fails to be subadditive. Danielsson, Jorgensen, Samorodnitsky, Sarma & de Vries (2005) proved that for most practical applications VaR is subadditive. More precisely, they demonstrated that VaR is subadditive for the relevant tails of all fat tailed distributions, provided the tails are not super fat (i.e., distributions with so fat tails that the first moment is not defined, such as the Cauchy law, with tail index strictly less than 1). Inui, Kijima & Kitano (2005) demonstrated the considerable bias of the VaR when used for a portfolio with fat-tailed distributions.

Note that the class of distortion risk measures (3.5) with concave \overline{f} is a strict subclass of the class of coherent risk measures, as is shown by the following example, taken from Dhaene, Vanduffel, Tang, Goovaerts, Kaas & Vyncke (2004).

Example 3.3 (Dutch risk measure) For any random variable X, consider the risk measure

$$\varrho[X] = \mathbb{E}[X] + \mathbb{E}[(X - \mathbb{E}[X])_{+}]. \tag{3.10}$$

This is a particular case of the "Dutch premium principle", defined for non-negative random variables by Kaas, Van Heerwaarden & Goovaerts (1994).

The Dutch risk measure (3.10) is in general not additive for co-monotonic risks. Hence, the Dutch risk measure (3.10) is an example of a risk measure that is coherent, although it is not a distortion risk measure. The example also illustrates the fact that coherent risk measures are not necessarily additive for co-monotonic risks.

4 Equivalent utility risk measures in CEU theory

4.1 CEU theory

A possible device to construct a capacity is to take a probability measure P on \mathcal{B} together with a distortion function f, and then define $C = f \circ P$. Such a capacity C is then called

a distorted probability measure. Conversely, for a given capacity C, one may wonder whether there exists a probability measure P such that $C = f \circ P$ for some distortion function f. The answer is in general negative: even if Ω is finite, there exist capacities that cannot be obtained as non-decreasing transformations of probability measures. See, e.g., Laeven (2005) for a simple counterexample.

Schmeidler (1989) modified the independence axiom into co-monotonic independence. This gave rise to the Choquet expected utility (CEU, in short) theory. In CEU theory, the prospect X is evaluated by

$$\mathbb{U}_{CEU}[X|u, C] = \int u(X)dC$$

$$= \int_{-\infty}^{0} \left(C[u(X) > t] - 1 \right) dt + \int_{0}^{+\infty} C[u(X) > t] dt$$

where u is a utility of wealth function, C is a capacity and integration is in the sense of Choquet; see Choquet (1953). The functional \mathbb{U}_{CEU} representing the preference relation in the CEU model is thus the Choquet integral of u(X) with respect to the capacity C. Note that the integrals on the right-hand side are Riemann integrals and they are well defined since C[u(X) > t] is a monotonic function in t. Moreover, with C defined directly on events, there is no necessity for any intermediate evaluation of probabilities.

4.2 Uncertainty aversion

Risk refers to situations where the perceived likelihoods of events of interest can be represented by probability distributions, whereas uncertainty refers to situations where the information available to the decision-maker is too imprecise to be summarized by a probability measure. Schmeidler's (1989) CEU model is set in the framework of uncertainty as defined by Anscombe & Aumann (1963).

Due to the absence of an objective probability distribution, the definition of uncertainty aversion has to be based on the random variables themselves. Schmeidler (1989) suggests the idea that a convex combination of non-co-monotonic random variables smooths the outcomes and makes an uncertainty-averse decision-maker better off. Specifically, a decision-maker is uncertainty-averse if, for every pair of random variables X and Y, if X is preferred over Y then $\alpha X + (1 - \alpha)Y$ is preferred over Y for any $\alpha \in [0, 1]$. The corresponding preferences are thus convex. Chateauneuf, Dana & Tallon (2000) established that the decision-maker is uncertainty-averse if, and only if, the capacity C is convex, that is,

$$C[A \cup B] + C[A \cap B] \ge C[A] + C[B] \text{ for all } A, B \in \mathcal{B}, \tag{4.1}$$

and the utility function u is concave. See also Leitner (2005) for related results.

If the reverse inequality holds in (4.1) then the capacity is called concave. Note that if the dual capacity \overline{C} is convex then C is concave. Furthermore, if C is convex then \overline{C} is concave and satisfies $C[A] \leq \overline{C}[A]$ for any $A \in \mathcal{B}$. The decision weights used in the computation of the Choquet integral will overweight large outcomes if C is concave and will overweight low outcomes if C is convex. Additivity is a special case that characterizes

uncertainty neutrality. Probability distributions are special cases of capacities which are both concave and convex. In the case of a distorted probability C[A] = f(P[A]), $A \in \mathcal{B}$, if f is convex then C is convex, and if f is concave then C is concave.

The core of the capacity C is defined as

$$core(C) = \{ probability measures P | P[A] \ge C[A] \text{ for all } A \in \mathcal{B} \}.$$

Since Schmeidler (1986), it is well-known that when the capacity is convex, its core is non-empty and the Choquet integral is given by

$$\int u(X)dC = \inf_{P \in core(C)} \int u(X)dP.$$

This relates the CEU model to the MEU one, presented in the next section.

There are thus two properties of special interest concerning the uncertainty aversion in CEU theory, namely the non-emptiness of the core and the convexity of the capacity. Chateauneuf & Tallon (2002) considered the preference for diversification in CEU. The agent is said to exhibit preference for diversification if, for any random variables X_1, \ldots, X_n that are all considered as equivalent by the agent (i.e., $\mathbb{U}_{CEU}[X_1|u,C] = \cdots = \mathbb{U}_{CEU}[X_n|u,C]$ in CEU), $\sum_{i=1}^n \alpha_i X_i$ is preferred over X_1 for any $\alpha_1 \geq 0, \ldots, \alpha_n \geq 0$ with $\sum_{i=1}^n \alpha_i = 1$. This concept is obviously related to the convexity for risk measures. They established that preference for portfolio diversification is equivalent to the agent having a convex capacity and a concave utility function.

4.3 Equivalent CEU risk measures

Most often in the financial mathematics literature, u is taken to be the identity function and C is assumed to be convex. We then get for the preference indicator simply the Choquet integral of X with respect to C, that is,

$$\mathbb{U}_{CEU}[X|\text{identity}, C] = \int X dC = \int_{-\infty}^{0} \left(C[X > t] - 1 \right) dt + \int_{0}^{+\infty} C[X > t] dt.$$

This integral is positively homogeneous, monotonic and translation equivariant.

The risk measure ϱ is then deduced from the indifference equation

$$0 = \mathbb{U}_{CEU}[\varrho[X] - X | \text{identity}, C] \Rightarrow \varrho[X] = -\int -X dC.$$

Note that, in general, $\int X dC \le -\int -X dC$. Equality holds for all random variables if, and only if, C is additive. This inequality can be used to explain the bid-ask spread observed in some financial markets, according to Chateauneuf, Kast & Lapied (1996). The integrals $\int X dC$ and $-\int -X dC$ are sometimes called lower and upper Choquet integrals, respectively. More precisely,

$$\int X dC = \inf_{P \in \operatorname{core}(C)} \int X dP \quad \text{and} \quad -\int -X dC = \sup_{P \in \operatorname{core}(C)} \int X dP$$

so that the length of the interval $(\int X dC, -\int -X dC)$ is intimately connected with the size of core(C).

After Denneberg (1990), the risk measure $\int X dC$ has been considered in insurance, e.g., by Wang (1996), Wang, Young & Panjer (1997) and De Waegenaere, Kast & Lapied (2003) under the name of Choquet pricing. An overview paper on the applicability of this type of risk measures for solvency purposes is Dhaene, Vanduffel, Tang, Goovaerts, Kaas & Vyncke (2004). Chateauneuf, Kast & Lapied (1996), Wang (2000, 2002) and Hamada & Sherris (2003) used the same mechanism for pricing in financial markets; see also Pelsser (2004). This risk measure is continuous, monotonic and co-monotonic additive. It is subadditive if, and only if, C is concave; it is superadditive for convex C.

Castagnoli, Maccheroni & Marinacci (2004) noticed that the presence on the market of assets without bid-ask spreads turns Choquet prices into standard expectations. More precisely, if an insurance company sets the safety loading according to a probability distortion in agreement with the increasing concave order \leq_{icv} (defined by $X \leq_{\text{icv}} Y \Leftrightarrow \mathbb{E}[\phi(X)] \leq \mathbb{E}[\phi(Y)]$ for all non-decreasing concave functions $\phi : \mathbb{R} \to \mathbb{R}$, provided the expectations exist) then it assigns either strictly positive safety loading to all contracts or zero safety loading to all contracts. Despite its theoretical interest, this result should not prevent the use of Choquet risk measures (since actuarial risk theory imposes strictly positive safety loading, to avoid almost certain ruin).

Kadane & Wasserman (1996) considered $\Omega = [0,1]$, \mathcal{B} the Borel subsets of Ω and symmetric capacities (that is, capacities C such that $\mu[A] = \mu[B] \Rightarrow C[A] = C[B]$, with μ the Lebesgue measure). They characterized coherent symmetric capacities (as those with doubly star-shaped distortion functions; see Kadane & Wasserman (1996) for a definition). If the capacity is further assumed to be concave then Section 4 in Kadane & Wasserman (1996) proves that the distortion functions $f(p) = \min\{p/\alpha, 1\}$ correspond to the extreme points of the set of all distortion functions. This directly implies that spectral risk measures can be represented as mixtures of Tail-VaR.

Note however that Marinacci (1999) proved that symmetric and coherent Choquet capacities turn out to be additive under a fairly mild condition. More specifically, the existence of a single non-trivial unambiguous event (in the sense that the interval (1.1) reduces to a single point) is enough to make them additive. This is a very strong property since excluding the existence of even a single non-trivial unambiguous event seems in general a very stringent assumption (recall that in the additive case, all events are unambiguous). This is obviously related to the result of Castagnoli, Maccheroni & Marinacci (2004) mentioned above.

5 Equivalent utility risk measures in MEU theory

5.1 MEU theory

Gilboa & Schmeidler (1989) suggested the Maximin expected utility (MEU, in short) model, also called multiple prior model. In this context, preference relations have a representation by the functional

$$\mathbb{U}_{MEU}[X|u,\mathcal{M}] = \inf \left\{ \int u(X) dP \middle| P \in \mathcal{M} \right\}$$

where u is a utility function and \mathcal{M} is a closed and convex set of additive probability measures. Thus, CEU coincides with MEU when the set \mathcal{M} is the core of C, that is, $\mathbb{U}_{CEU}[X|u,C] = \mathbb{U}_{MEU}[X|u,\operatorname{core}(C)]$. Moreover, when an arbitrary (closed and convex) set \mathcal{M} is given and one defines $C[A] = \inf\{P[A]|P \in \mathcal{M}\}$, C is not necessarily convex. Furthermore, even if C happens to be convex, \mathcal{M} does not have to be its core.

In addition to the usual assumptions on the preference relation such as transitivity, completeness, continuity and weak monotonicity, Gilboa & Schmeidler (1989) postulated uncertainty aversion and certainty-independence to derive MEU theory. The certainty-independence axiom is weaker than standard independence (since only mixings with degenerate distribution functions corresponding to certainty are considered).

5.2 Equivalent MEU risk measures

Taking u(x) = x gives the class of coherent risk measures in the sense of Artzner, Delbaen, Eber & Heath (1999). More precisely, the equivalent utility principle then yields

$$0 = \mathbb{U}_{MEU}[\varrho[X] - X | u, \mathcal{M}] = \inf_{P \in \mathcal{M}} \int (\varrho[X] - X) dP$$

whence it follows that

$$\varrho[X] = \sup_{P \in \mathcal{M}} \int X dP. \tag{5.1}$$

Risk measures of this type have been used by Castagnoli, Maccheroni & Marinacci (2002) to compute insurance premiums. Huber (1981) already considered functionals (5.1) in robust statistics. The elements of \mathcal{M} are called generalized scenarios in Artzner, Delbaen, Eber & Heath (1999). Huber (1981) proved for the case of a finite set Ω , that a risk measure is coherent if, and only if, it has an upper expectation representation. This result remains valid for more general spaces (see Delbaen (2002)), though in that case "the important conceptual aspects are buried under a mass of technical complications of a measure theoretic and topological nature," as is stated by Huber (1981).

Risk measures of the form (5.1) have been considered for a long time in actuarial science, especially when \mathcal{M} is a moment space (or more generally, when \mathcal{M} is defined by a set of integral constraints); see, e.g., Section 5.3 of Goovaerts, De Vijlder & Haezendonck (1984) for a detailed account.

6 Equivalent utility risk measures in RDEU theory

6.1 RDEU theory

In this section, we present a theory which combines the EU and the DU assumptions, to a certain extent. The rank-dependent expected utility (RDEU, in short) theory was first proposed by Quiggin (1982). It is a generalization of EU and DU theories that preserves the properties of continuity, transitivity and weak monotonicity. The name comes from the fact that cumulative probabilities are transformed, so that the new weight assigned to a particular outcome depends on its rank among possible outcomes. We do not give

the set of axioms connected with RDEU since the modified version of the independence axiom appearing in Quiggin (1982) does not admit a clear behavioral interpretation. For other sets of axioms, see e.g. Chateauneuf & Wakker (1999) and Abdellaoui (2002). The RDEU approach combines the utility function $u(\cdot)$ from EU with a distortion function $f(\cdot)$ from DU. For interesting applications to insurance, see, e.g., Chateauneuf, Cohen & Meilijson (1997).

Under the rank-dependent expected utility model, a decision-maker is characterized by a utility function u (that plays the role of utility on certainty) in conjunction with a distortion function f (that plays the role of a probability perception function). Such a decision-maker prefers the fortune Y to the fortune X if, and only if, $\mathbb{U}_{RDEU}[X|u, f] \leq \mathbb{U}_{RDEU}[Y|u, f]$ where

$$\mathbb{U}_{RDEU}[X|u, f] = \int_{t=-\infty}^{0} \left(f(P[u(X) > t]) - 1 \right) dt + \int_{t=0}^{+\infty} f(P[u(X) > t]) dt$$

with a similar expression for Y. It is easy to see that if f(p) = p, we get the EU model, that is, $\mathbb{U}_{RDEU}[X|u]$, identity] = $\mathbb{U}_{EU}[X|u]$, whereas if u(x) = x we get the DU paradigm, that is, $\mathbb{U}_{RDEU}[X|\text{identity}, f] = \mathbb{U}_{DU}[X|f]$.

Just as in EU, the utility function u is normally taken to be concave or linear. There is less consensus concerning the appropriate shape for the distortion function f. For descriptive purposes, Quiggin (1982) originally proposed an S-shaped function, overweighting extreme low probability outcomes. Other authors suggested to consider convex distortion functions f (this property is akin to risk aversion, as shown above).

Simple distortion functions are those given by the one-parameter family $f_{\alpha}(p) = \min\{p/\alpha, 1\}$ for some $\alpha \in [0, 1]$. In this case, we have

$$\mathbb{U}_{RDEU}[X|u, f_{\alpha}] = \frac{1}{\alpha} \int_{0}^{\alpha} u(F_{X}^{-1}(p)) dp.$$

This family has been considered, e.g., in Bassett, Koenker & Kordas (2004).

Remark 6.1 Following the usage of Knight (1921) in distinguishing uncertainty (i.e., ambiguity) and risk, Gilboa (1987) and Schmeidler (1989) do not take objective probabilities as given but consider subjective probabilities. Capacities are then used and prospects are compared on the basis of Choquet integrals. Wakker (1990) established that, provided weak monotonicity is fulfilled, this approach is equivalent to RDEU. The subtle difference between the two approaches can be summarized as follows: in RDEU, decision-makers know the objective probabilities but distort them when choosing among risky prospects. In CEU, decision-makers use their own (possibly non-additive) subjective probabilities, derived from some capacity.

6.2 Risk aversion

As RDEU distinguishes attitudes towards outcomes and attitudes towards probabilities, risk aversion in RDEU must combine two different concepts. First, there is outcome risk aversion, associated with the idea that the marginal utility of wealth is declining. This is

the standard notion of risk aversion from the EU theory defined by concavity of the utility function. Second, there are attitudes specific to probability preferences. Risk aversion in probability weighting corresponds to pessimism: the decision-maker adopts a set of decision weights that yields an expected value for a transformed risky prospect lower than the mathematical expectation. An alternative, more restrictive, characterization of pessimism leads to a definition of risk aversion in terms of \leq_{icv} . Note that the restriction to concave utility functions does not prohibit risk seeking behaviors in RDEU. It is indeed possible to model risk seeking with diminishing marginal utility of wealth.

The transformation f may be regarded as pessimistic if $f(p) \le p$ for all p. This comes from the fact that $f(p) \le p$ for all p if, and only if, for any random variable X and concave utility u

$$\mathbb{U}_{RDEU}[X|u, f] \leq \mathbb{U}_{RDEU}[\mathbb{E}[X]|u, f] = u(\mathbb{E}[X]),$$

that is, $\mathbb{E}[X]$ is preferred over X. Therefore, concavity of u and pessimism are reasonable conditions for risk aversion.

Let us now turn to the definition of risk aversion as preservation of \leq_{icv} . It can be shown that a decision-maker will be consistent with \leq_{icv} if, and only if, u is concave and f is convex (see, e.g., Quiggin (1993)). Convexity of the distortion function f implies a strong form of pessimism that can be roughly summarized as follows: the worst outcome is overweighted most, and the best outcome underweighted most. Unlike the situation in EU theory, preservation of \leq_{icv} is not equivalent to preference for certainty over risk.

6.3 Equivalent RDEU risk measures

Heilpern (2003) and Tsanakas & Desli (2003) introduce a class of risk measures which can be considered as the solutions of the indifference equations for RDEU. Lemma 1 in Heilpern (2003) establishes that provided u is non-decreasing and concave, $\mathbb{U}_{RDEU}[X|u,f] \leq u(\mathbb{U}_{DU}[X|f])$ whatever the distortion function f. The equivalent RDEU risk measure solves

$$u(0) = \mathbb{U}_{RDEU}[\varrho[X] - X | u, f]. \tag{6.1}$$

In general, there is no explicit solution to this equation. A notable exception is with utility function of the form (2.3) for which

$$\varrho[X] = \frac{1}{c} \ln \left(\int_0^{+\infty} \overline{f} \left(P[\exp(cX) > t] \right) dt \right). \tag{6.2}$$

The equivalent utility risk measures solving (6.1) are great variety. Provided u is non-decreasing and concave and f is a convex distortion function, Heilpern (2003, Theorem 1) proved that the resulting risk measure is translation equivariant, preserves \leq_{icx} and does not entail unjustified safety loading, that is $\varrho[c] = c$ for all real constants c. If u(x) = x, it is moreover positively homogeneous, co-monotonic additive and subadditive. Note that in this case, we are back to the DU paradigm. It is additive for independent random variables only in very special cases (broadly speaking, if u is linear or exponential and f is the identity).

In particular, the risk measure (6.2) is additive for independent random variables only if $\overline{f}(p) = p$ (that is, in the EU case). Moreover, the risk measure (6.2) is strongly monotonic, translation equivariant and convex. Tsanakas & Desli (2003) studied in detail the properties of this risk measure regarding sensitivity to portfolio size and to risk aggregation. They conclude that this risk measure behaves approximately as a coherent risk measure for smaller portfolios of risks, while for larger portfolios the risk aversion induced by the utility function becomes prevalent, and the sensitivity to liquidity and risk aggregation issues gradually increases.

7 Discussion

In this paper, we have derived the great majority of the classical risk measures using the actuarial equivalent utility principle in different theories for decision under risk and uncertainty. As such it makes clear the economic rationale behind these risk measures.

It is now also clear that risk measures encountered in the financial mathematics literature correspond to linear utility of wealth functions. Disregarding utility of wealth does not always seem reasonable from an economic viewpoint. The linearity of the utility function makes risk measures insensitive to liquidity risk: for example, according to positive homogeneity, changes in the size of a portfolio, given that its composition is unchanged, should only affect a proportional change in capital requirements. This disregards that very large portfolios might produce very large losses that, in turn, can make it difficult for the holder of the portfolio to raise sufficient cash to meet his obligations. This has been pointed out, e.g., by Dhaene, Goovaerts & Kaas (2003).

The RDEU framework simultaneously accounts for utility functions and distorted probabilities: the agent is equipped with both a utility function and a probability distortion. Risk measures emerging from RDEU inherit properties from both EU and DU: in the case of a linear utility function equivalent DU risk measures are obtained as special cases while, in the case of a linear probability distortion function, equivalent EU risk measures are found. In general, the properties of equivalent RDEU risk measures lie somewhere in between of these two extreme cases. Their general behavior is nevertheless difficult to determine since it is difficult to know at which point the utility or the distortion function will have the most influence. As pointed out by Tsanakas & Desli (2003), equivalent RDEU risk measures are convex under mild conditions. Note that CEU is even more general.

In particular, the distortion exponential risk measure (6.2) proposed by Tsanakas & Desli (2003) combines the properties of the DU equivalent and (2.4) risk measures. It is obtained as a special case of equivalent RDEU risk measures when an exponential utility function (2.3) is used. It inherits properties from both (2.4) and DU equivalent risk measures. Whether the properties of the distortion exponential risk measures are closer to those of (2.4) or those of equivalent DU risk measures depends to some extent on the underlying risks that are examined. For relatively small portfolios, the distortion exponential risk measure is approximately equivalent to the equivalent DU risk measures and inherits their properties. For larger portfolios, for which liquidity and risk aggregation become an issue, the effect of the utility function (2.3) becomes prevalent and the distortion exponential risk measure inherits its properties from (2.4). The definition of what a large

portfolio is depends on the specific situation and preferences of its holder. It can be controlled by modifying the risk aversion parameter of the exponential function. The numerical illustrations in Tsanakas & Desli (2003) support this evidence.

Alternatives to (6.2) can be obtained by substituting other classes of utility functions to the exponential one (2.3). One could think of CARA (for constant absolute risk aversion) utility functions given by $u(x) = \frac{x^{\gamma}}{\gamma}$ for $\gamma \neq 0$ and $\ln x$ when $\gamma = 0$ that are often used in economic applications.

Acknowledgements. The financial support of the Belgian Government under the contract "Projet d'Actions de Recherche Concertées" ARC 04/09-320 is gratefully acknowledged by Michel Denuit, who also thanks the Banque Nationale de Belgique for financial support under grant "Risk Measures and Economic Capital". Jan Dhaene and Marc Goovaerts acknowledge the financial support of the Onderzoeksfonds K.U. Leuven (GOA/02: Actuariële, financiële en statistische aspecten van afhankelijkheden in verzekerings- en financiële portefeuilles). Roger Laeven acknowledges the support of the Dutch Organization for Scientific Research (No. NWO 42511013).

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