

1 **Original Article**2 **Reduced Lung-Cancer Mortality with Volume CT**
3 **Screening in a Randomized Trial**

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33 **Abstract**34 **Background**

35 There are limited data from randomized trials regarding whether volume-based,
 36 low-dose computed tomographic (CT) screening can reduce lung-cancer
 37 mortality among male former and current smokers.

38 **Methods**

39 A total of 13,195 men (primary analysis) and 2594 women (subgroup analyses)
 40 between the ages of 50 and 74 were randomly assigned to undergo CT screening
 41 at T0 (baseline), year 1, year 3, and year 5.5 or no screening. We obtained data
 42 on cancer diagnosis and the date and cause of death through linkages with
 43 national registries in the Netherlands and Belgium, and a review committee
 44 confirmed lung cancer as the cause of death when possible. A minimum follow-
 45 up of 10 years until December 31, 2015, was completed for all participants.

1 Results

2 Among men, the average adherence to CT screening was 90.0%. On average,
3 9.2% of the screened participants underwent at least one additional CT scan
4 (initially indeterminate). The overall referral rate for suspicious nodules was
5 2.1%. At 10 years of follow-up, the incidence of lung cancer was 5.58 cases per
6 1000 person-years in the screening group and 4.91 cases per 1000 person-years
7 in the control group; lung-cancer mortality was 2.50 deaths per 1000 person-
8 years and 3.30 deaths per 1000 person-years, respectively. The cumulative rate
9 ratio for death from lung cancer at 10 years was 0.76 (95% confidence interval
10 [CI], 0.61 to 0.94) in the screening group as compared with the control group,
11 similar to the values at years 8 and 9. Among women, the rate ratio was 0.67
12 (95% CI, 0.38 to 1.14) at 10 years of follow-up, with values of 0.41 to 0.52 in
13 years 7 through 9.

14 Conclusions

15 In this trial involving high-risk persons, lung-cancer mortality was significantly
16 lower among those who underwent volume CT screening than among those who
17 underwent no screening. There were low rates of follow-up procedures for
18 results suggestive of lung cancer. (Funded by the Netherlands Organization of
19 Health Research and Development and others; NELSON Netherlands Trial
20 Register number, NL580.)

21 Lung cancer is the leading cause of death from cancer worldwide (18.4% of all
22 cancer deaths) and causes more deaths than breast, colorectal, and cervical
23 cancers combined — cancers for which population-based screening programs
24 exist.¹ Only 15% of patients with lung cancer are still alive 5 years after
25 diagnosis, because approximately 70% of patients have advanced disease at the
26 time of diagnosis.² Although smoking prevalence is decreasing in Western
27 countries, 17 to 28% of adults currently still smoke, and smoking initiation
28 remains substantial in youths.³ Lung cancer and other tobacco-related diseases
29 are expected to remain important health problems worldwide for decades.^{2,4}

30 The U.S.-based National Lung Screening Trial (NLST) showed that a strategy
31 of three annual computed tomographic (CT) screenings resulted in 20.0% lower
32 mortality from lung cancer than screening with the use of chest radiography
33 among 53,454 participants at high risk for lung cancer after a median follow-up
34 of 6.5 years, and the trial recently confirmed a 19% (maximum) lower mortality
35 at a median follow-up of 5.5 and 6.0 years.^{5,6} The U.S. Preventive Services Task
36 Force requested an independent review and a modeling study,^{7,8} which resulted in
37 the recommendation to annually screen persons 55 to 80 years of age with a
38 smoking history of 30 or more pack-years, who currently smoke or quit smoking
39 within the past 15 years. No other trial of lung-cancer screening has yet reported
40 benefits with respect to mortality.⁹

41 The Dutch–Belgian lung-cancer screening trial (Nederlands–Leuvens
42 Longkanker Screenings Onderzoek [NELSON]), a population-based, randomized,
43 controlled trial initiated in 2000, aimed to show a reduction in lung-cancer

1 mortality of 25% or more with volume-based, low-dose CT lung-cancer screening
2 in high-risk male participants at 10 years of follow-up. Here, we report lung-
3 cancer incidence, mortality, and the performance of the four screening rounds in
4 the NELSON trial among male participants (main analysis) and female
5 participants (subgroup analyses).

6 **Methods**

7 **Trial Oversight**

8 The trial was approved by the Dutch Minister of Health and the medical ethics
9 committee at each participating site.¹⁰ Conceptualization of the trial, funding
10 acquisition, data collection and curation, analysis of the primary outcome, the
11 writing of the first draft of the manuscript, and revision of the manuscript based
12 on review comments were performed by Erasmus MC and University Medical
13 Center Groningen (UMCG). CT screening and follow-up were performed by the
14 four screening sites (UMCG, University Medical Center Utrecht, Spaarne
15 Gasthuis, and University Hospital Leuven). An independent cause-of-death
16 committee defined the cause of death for some of the deceased participants (see
17 the Supplementary Appendix, available with the full text of this article at NEJM.
18 org). Data on workup, cancer diagnosis and stage, treatment, vital status, and
19 cause of death were obtained through linkages with the Dutch Center for
20 Genealogic and Heraldic Studies, Statistics Netherlands, and the Dutch Cancer
21 Registry. Primary outcome data were kept confidential until unblinding. None of
22 the funders had any role in the trial design, the collection or analysis of the
23 data, or the writing of the manuscript. The authors vouch for the completeness
24 and accuracy of the data and for the fidelity of the trial to the protocol (available
25 at NEJM.org). No one who is not an author contributed to the writing of the
26 manuscript.

27 **Power Calculation and Eligibility Criteria**

28 An overview of the previously published power calculation and trial design is
29 available in the Supplementary Appendix.¹¹⁻¹³ The preferred risk-based selection
30 scenario (scenario D¹¹) required 17,300 to 27,900 participants (current or former
31 smokers [those who had quit ≤ 10 years ago] who had smoked >15 cigarettes a
32 day for >25 years or >10 cigarettes a day for >30 years) to show a lung-cancer
33 mortality that was lower by 20 to 25% in the screening group than in the
34 control group at 10 years of follow-up, given the following conditions: one-sided
35 testing, based on experience with the European Randomized Study of Screening
36 for Prostate Cancer (two-sided testing was used for the final analyses); 90%
37 power; 95% adherence in the screening group; 5% contamination (i.e., lung-
38 cancer screening) in the control group; and an expected lung-cancer mortality of
39 3.4 per 1000 person-years without screening at 10 years of follow-up.¹¹ Exclusion
40 criteria were patient report of moderate or severe health problems and an
41 inability to climb two flights of stairs; a body weight of more than 140 kg;
42 current or past renal cancer, melanoma, or breast cancer; a diagnosis of lung

1 cancer or treatment related to lung cancer within the past 5 years; or a chest CT
2 scan within the past year.^{11,12} A current smoker was defined as a person who had
3 smoked cigarettes during the last 2 weeks.

4 The trial focused on men (see the Supplementary Appendix).¹¹ At the time of
5 initiation (2000 through 2004), only a small number of women were eligible,
6 because smoking was much less prevalent and much less intensive among
7 women than among men. Because of the importance of the inclusion of women,
8 a sample of high-risk women was approached for participation.

9 **Recruitment**

10 On the basis of population registries, 606,409 persons 50 to 74 years of age who
11 lived in four selected regions in the Netherlands and Belgium were approached
12 with a general questionnaire and brief information about the trial in 2003 (first
13 recruitment) or 2005 (second recruitment) (see the Supplementary Appendix,
14 including Fig. S2).¹⁴ A total of 30,959 respondents of the 150,920 who returned
15 questionnaires were eligible. Eligible persons were invited to participate; 15,822
16 persons (51.1%), who provided written informed consent, underwent the initial
17 randomization (in a 1:1 ratio) from December 2003 through July 2006 (median
18 randomization date, November 2004) (Fig. S7).^{11,13,14} After linkage with Statistics
19 Netherlands and the Dutch Center for Genealogic and Heraldic Studies, 30
20 participants had died after providing informed consent and before the
21 randomization date, which resulted in 15,792 formal participants (13,195 men,
22 2594 women, and 3 participants with unknown sex) (Table S1).

23 **Screening Rounds and Nodule-Management Protocol**

24 The screening rounds and the nodule-management protocol have been described
25 previously (summarized in Fig. S8).^{13,15-19} In short, from January 2004 through
26 December 2012, participants in the screening group were invited to undergo
27 four rounds of low-dose CT screening for lung cancer that were performed in
28 the four CT screening sites with intervals of 1, 2, and 2.5 years.

29 For CT screening, low-dose 16-multidetector or, in later rounds,
30 64-multidetector CT systems were used to acquire isotropic volume data, without
31 administration of contrast medium. Apart from local readings, all images were
32 analyzed centrally at UMCG with the use of semiautomated software (LungCare,
33 version Somaris/5 VA70C-W, Siemens Medical Solutions). The analysis included
34 the semiautomated segmentation of nodules and determination of the nodule
35 volume.²⁰ If the software was not able to segment a nodule accurately, the
36 volume was corrected manually by the radiologist.²¹ Depending on the volume
37 and volume-doubling time, a screening could be negative, indeterminate, or
38 positive (Fig. S3). Participants in the control group underwent no screening.

39 **Follow-up Data**

40 Follow-up data were retrieved from national linkages at approximately 5, 7, and
41 10 to 11 years of complete follow-up. A total of 18 persons (13 men and 5
42 women) could not be linked, because a digital consent form could not be
retrieved. Population data were available regarding randomization date, sex, date

1 of lung-cancer diagnosis, and date and cause of death for all deceased Belgian
2 persons up to December 2013 and September 2018 through linkages in January
3 2016 and October 2018, respectively.

4 **Cause-of-Death Review**

5 The primary outcome of the NELSON trial was lung cancer–specific mortality. A
6 clinical expert committee was formed to assign the cause of death by an
7 evaluation process using a flow chart and predetermined criteria.²² A total of
8 296 completed and blinded medical files of 426 deceased Dutch male patients
9 with lung cancer (69.5%) were reviewed and compared with official death
10 certificates (cutoff, 10 years of follow-up or December 31, 2015). The overall
11 concordance among members of the expert committee was 86.1%. The
12 sensitivity and specificity of the official death certificate were 92.6% and 98.8%,
13 respectively.²³ Death from lung cancer was considered valid only if the expert
14 committee had concluded that lung cancer was the cause of death. The
15 international mortality advisory committee deemed possible biases to be
16 relatively small and agreed on further use of official statistics for the primary
17 outcome, if lung cancer as the cause of death was recorded in the national
18 registry for vital statistics.

19 **Statistical Analysis**

20 The primary analysis of the trial consisted of a comparison of lung-cancer
21 mortality between the screening group and the control group (main analysis,
22 men; subanalyses, women), according to the intention-to-screen principle.
23 Specifically, the rate ratio for death from lung cancer was compared between the
24 two groups; the rate ratio was derived as the ratio of event rates, under the
25 assumption of a Poisson distribution for the number of events (two-sided test).
26 Secondary analyses compared all-cause mortality and the incidence of first
27 recorded diagnosis of lung cancer between the two groups. The date of
28 censoring of data for first recorded lung cancer{q1}, death from lung cancer, and
29 death from any cause was December 31, 2015, or 10 years of follow-up since
30 randomization (whichever came first). Event rates were defined as the ratio of
31 the number of events to the person-years at risk for the event. For the incidence
32 of first recorded lung cancer, person-years were measured from the time of
33 randomization to the date of diagnosis of lung cancer, death, or censoring of
34 data (whichever came first); for mortality{q2}, person-years were measured from
35 the time of randomization to the date of death or censoring of data (whichever
36 came first). Previously published definitions are summarized in the
37 Supplementary Appendix.^{13,15,16}

38 Continuous variables are presented as means and standard deviations (normal
39 distribution) or as medians, interquartile ranges, and ranges (skewed
40 distribution). Differences in distributions of baseline characteristics of
41 participants in the screening group and participants in the control group were
42 analyzed with the use of Pearson's chi-square test for nominal or categorical
43 variables and the Mann–Whitney test for ordinal or continuous variables with a

1 nonnormal distribution. Analyses were performed with the use of Stata software,
2 R statistical packages, and SPSS software, version 25. Exact methods were used
3 to calculate confidence intervals for the rate ratios. P values were calculated with
4 the use of two-sided exact tests; a P value of less than 0.05 was considered to
5 indicate statistical significance. No corrections for multiple comparisons were
6 included. Missing data for the primary outcome were negligible owing to the
7 linkages with the national registries (>98% coverage).

8 **Results**

9 **Baseline Characteristics of Male Participants**

10 A total of 13,195 male participants were randomly assigned to either the
11 screening group (6583 men) or the control group (6612 men). Baseline
12 characteristics did not differ significantly between the two groups, except for
13 duration of smoking (Table 1). At randomization, the median age of the male
14 participants was 58 years in each group (interquartile range, 55 to 63 in the
15 screening group and 54 to 63 in the control group), with a median smoking
16 history of 38.0 pack-years (interquartile range, 29.7 to 49.5) in each group.
17 Overall, 44.9% of the male participants were former smokers.

18 **Screening Results in Male Participants**

19 In total, 22,600 CT scans were performed, and screening uptake was on average
20 90.0% (95% confidence interval [CI], 76.9 to 95.8) (Table 2). In 9.2% of the
21 scans (2069 of 22,600), an indeterminate screening test required a repeat CT
22 scan to calculate volume-doubling time before the final screening-test outcome
23 could be defined. At baseline, the percentage of indeterminate tests was highest
24 (19.7%), after which it decreased to between 1.9% and 6.7% at year 1 through
25 year 5.5. In follow-up rounds, 55% of new nodules resolved.²⁴ Finally, 467 of
26 22,600 CT scans (2.1%) were test-positive and required further workup by the
27 pulmonologist, leading to 203 screening-detected lung cancers. The overall
28 positive predictive value of a positive screening test was 43.5%. This means that
29 264 of 19,327 screened participants (1.4%) had a false positive test. No adverse
30 events were reported. After a positive screening test, the national guidelines for
31 treatment of lung cancer were applied by the local hospitals.

32 **Lung Cancer in Male Participants**

33 Figure 1A shows the cumulative incidence of lung cancer according to follow-up
34 period and trial group. (Results for lung cancer of stage III or higher are
35 provided in Fig. S5.) At 10-year follow-up, the cumulative incidence of lung
36 cancer was 5.58 cases per 1000 person-years (341 lung cancers with a known
37 date of diagnosis) among male participants in the screening group and 4.91
38 cases per 1000 person-years (304 lung cancers with a known date of diagnosis)
39 among those in the control group (rate ratio, 1.14; 95% CI, 0.97 to 1.33). A total
40 of 59.0% (203 of 344) of all lung cancers in the screening group were detected
41 on screening (Table 3), and 12.8% (44 of 344) were interval cancers. Screening-
42 detected lung cancers were substantially more often diagnosed in stage IA or IB

1 (58.6%), whereas only 14.2% (screening group) and 13.5% (control group) of the
2 participants with non–screening-detected lung cancers received a diagnosis in
3 stage IA or IB ($P<0.001$). Stage IV cancer was diagnosed in almost half the
4 participants with non–screening-detected lung cancers (51.8% in the screening
5 group and 45.7% in the control group), whereas only 9.4% of the screening-
6 detected lung cancers were diagnosed in stage IV ($P<0.001$). Most (screening-
7 detected) lung cancers were adenocarcinomas (52.0% in the screening group and
8 43.8% in the control group, $P=0.03$).

9 **Mortality**

10 At 10 years of follow-up, 156 men with a known date of lung-cancer diagnosis
11 in the screening group and 206 in the control group had died from lung cancer
12 (2.50 deaths per 1000 person-years and 3.30 deaths per 1000 person-years,
13 respectively), which resulted in a cumulative rate ratio for death from lung
14 cancer of 0.76 (95% CI, 0.61 to 0.94). Similar rate ratios, which differed
15 significantly between the two groups, were observed at years 8, 9, and 11 (Fig. 1
16 and Table S3). Table 4 shows the causes of death in the two groups. All-cause
17 mortality at 10 years of follow-up was 13.93 deaths per 1000 person-years among
18 male participants in the screening group and 13.76 deaths per 1000 person-years
19 among those in the control group (rate ratio, 1.01; 95% CI, 0.92 to 1.11).

20 Analyses of data from the small subsample of women (with a known date of
21 lung-cancer diagnosis) showed a rate ratio for death from lung cancer of 0.67
22 (95% CI, 0.38 to 1.14) at 10 years of follow-up. The rate ratio was 0.46 (95% CI,
23 0.21 to 0.96) at 7 years, 0.41 (95% CI, 0.19 to 0.84) at 8 years, and 0.52 (95% CI,
24 0.28 to 0.94) at 9 years.

25 **Sensitivity Analyses**

26 At the 11-year follow-up (up to December 2016), the rate ratio for death from
27 lung cancer among male participants was 0.78 (95% CI, 0.63 to 0.95). After 10
28 years of follow-up, the subgroup of men 50 to 54 years of age — not included in
29 the NLST — had a rate ratio of 0.85 (95% CI, 0.48 to 1.50). The subgroup of
30 men 65 to 69 years of age had the lowest rate ratio of any age group, at 0.59
31 (95% CI, 0.35 to 0.98) (Table S2).

32 Approximately 50% of the participants in the NELSON trial met the eligibility
33 criteria of the NLST. Among NLST-eligible men, the rate ratio at 10 years of
34 follow-up was 0.82 (95% CI, 0.64 to 1.05). If all deaths from lung cancer, with
35 no restriction regarding known date of diagnosis, were included, the rate ratio
36 would be 0.76 (95% CI, 0.62 to 0.94) among all men in the NELSON trial and
37 0.81 (95% CI, 0.63 to 1.04) among NLST-eligible men.

38 **Discussion**

39 In the NELSON trial, volume CT lung-cancer screening of high-risk former and
40 current smokers, with the introduction of growth-rate assessment as an imaging
41 biomarker for indeterminate tests, resulted in low referral rates for additional

1 assessments and substantially lower lung-cancer mortality (in both sexes) than
2 no screening, despite screening intervals that increased over time. Adherence to
3 CT screening was very high; at least 87.6% of the male participants underwent
4 three screenings. In line with the mortality outcomes, volume CT screening in
5 the NELSON trial has led to a substantial shift to lower-stage cancers at the
6 time of diagnosis as well as to more frequent eligibility for curative treatment
7 (mainly surgical).²⁶ Because only modest differences were found between
8 participants and eligible nonrespondents,¹⁴ we expect the results to be highly
9 generalizable.

10 In the small subsample of women, the effects of screening on lung-cancer
11 mortality were consistently more favorable. Post hoc analyses from the NLST
12 also showed weak evidence of a differential effect size according to sex and
13 histologic type.²⁷ In addition, the recently reported rate ratio for death from lung
14 cancer (screening vs. no screening) {q3} in the NLST was 0.95 (95% CI, 0.83 to
15 1.10) among men and 0.80 (95% CI, 0.66 to 0.96) among women (dilution-
16 adjusted analysis).⁶ Recently, the German Lung Cancer Screening Intervention
17 Trial showed a significant benefit with respect to lung-cancer mortality in the
18 small subgroup of women who were invited to undergo screening (hazard ratio,
19 0.31; 95% CI, 0.10 to 0.96).²⁸ These outcome data are also consistent with
20 differences between the sexes in the screening-detectable preclinical period (i.e.,
21 the period in which the lung cancer is detectable through CT screening but has
22 not yet clinically manifested itself through symptoms).²⁹ Ad hoc analyses of data
23 from male participants in the NELSON trial who met the eligibility criteria of
24 the NLST (although not powered and with overlapping confidence intervals)
25 suggest more favorable effects on lung-cancer mortality than in the NLST,
26 despite lower referral rates for suspicious lesions. Important differences were
27 seen in screening results at baseline in the NELSON trial (volume-based nodule-
28 management protocol) as compared with the NLST (diameter-based nodule-
29 management protocol): the percentage of patients with a positive test was 2.1%
30 in the NELSON trial and 24% in the NLST, and the positive predictive value was
31 43.5% and 3.8%, respectively.⁵

32 At baseline, participants in the screening group reported a longer duration of
33 smoking than those in the control group but the same number of pack-years.
34 Furthermore, smoking behavior was similar (intention-to-treat analyses) in the
35 two groups after 2 years of follow-up.³⁰ Bias in screening effect in favor of the
36 screening group is therefore not expected. The NELSON trial was not powered
37 to show a possible favorable difference in all-cause mortality (expected within
38 the range of 2.5%), because it would have required unrealistic sample sizes.³¹
39 Comparisons of other causes of death showed no meaningful differences
40 between the screening group and the control group.

41 Concerns have been raised about the potential for overdiagnosis in lung-
42 cancer screening. Excess-incidence analysis of data from the NLST estimated an
43 upper boundary of overdiagnosis risk of 18.5%.³² In the NELSON trial, an excess
44 of 40 cases (344 vs. 304) was found among the male participants in the

1 screening group 10 years after randomization (4.5 years after the final screening
2 round), which suggests an excess-incidence overdiagnosis rate of 19.7%
3 (bootstrapped 95% CI, -5.2 to 41.6) for screening-detected cases. However,
4 extending the follow-up to 11 years after randomization (5.5 years after the final
5 screening round) reduced the number of excess cases to 18, yielding an excess-
6 incidence overdiagnosis rate of 8.9% (bootstrapped 95% CI, -18.2 to 32.4) for
7 screening-detected cases. This is in line with modeling analyses suggesting that
8 the lead time of CT screening can be as long as 9 to 12 years for some cancers,
9 which indicates that appropriate estimation of the level of overdiagnosis in the
10 NELSON trial requires additional years of follow-up.³³ Because of this, an
11 overdiagnosis rate of 8.9% for screening-detected cases may be considered as the
12 upper limit of overdiagnosis in the NELSON trial. The clinical management
13 strategy in the NELSON trial was highly restrictive with respect to invasive
14 diagnosis and treatment of persistent subsolid nodules.

15 The high adherence to CT screening may reflect a high level of
16 conscientiousness among trial participants. In the future, improvement in
17 screening selection (personalized risk-based approach) will probably result in a
18 more favorable trade-off between harms and benefits of CT lung-cancer
19 screening.^{4,9,34-38}

20 The NELSON trial showed that volume CT lung-cancer screening, with low
21 rates of follow-up procedures for test results suggestive of lung cancer, resulted
22 in substantially lower lung-cancer mortality than no screening among high-risk
23 persons. Volume CT screening enabled a significant reduction of harms (e.g.,
24 false positive tests and unnecessary workup procedures), without jeopardizing
25 favorable outcomes. Trial data suggest greater benefits in women than in men,
26 but in a subgroup with a relatively low number of women. More research is
27 required in women, as well as in other subgroups.

Data sharing

28 A data sharing statement provided by the authors is available with the full text of this article at
29 NEJM.org.

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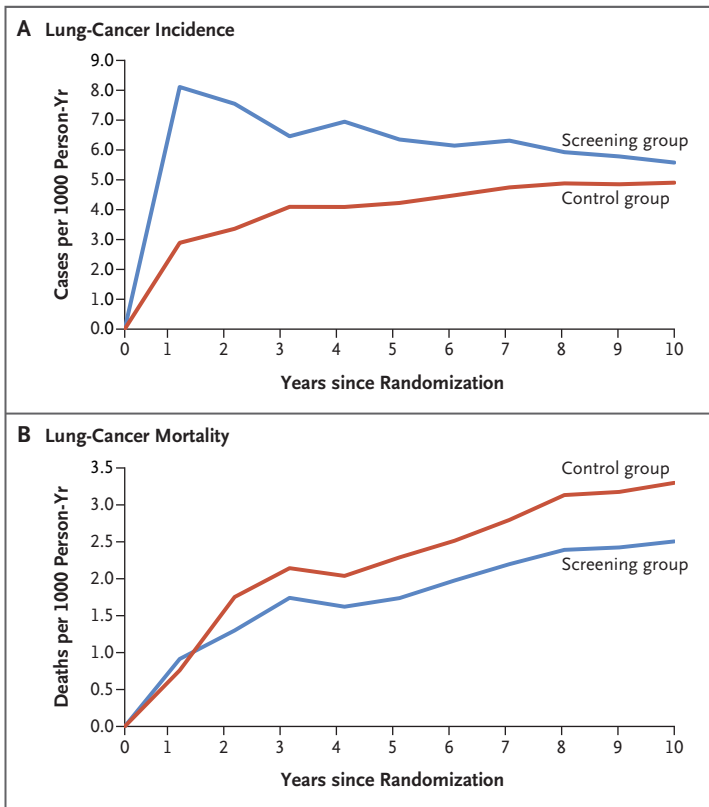


Figure 1. Lung-Cancer Incidence and Lung-Cancer Mortality among Male Participants.

Panel A shows the cumulative lung-cancer incidence (per 1000 person-years) according to follow-up year since randomization. Panel B shows the cumulative lung-cancer mortality (per 1000 person-years) according to follow-up year since randomization. Cause of death (with known date of lung-cancer diagnosis) was defined by the cause-of-death committee, if available, or by vital-statistics registries.

Table 1. Baseline Characteristics of the Male Participants at Randomization.*			
Characteristic	Screening Group (N=6583)	Control Group (N=6612)	P Value
Age			
Median (IQR) — yr	58 (55–63)	58 (54–63)	0.35
Range — yr	46–76	34–89	
Distribution — no./total no. (%)†			
<50 yr	3/6560 (<0.1)	6/6571 (0.1)	
50–54 yr	1611/6560 (24.6)	1694/6571 (25.8)	
55–59 yr	2226/6560 (33.9)	2231/6571 (34.0)	
60–64 yr	1554/6560 (23.7)	1475/6571 (22.4)	
65–69 yr	797/6560 (12.1)	781/6571 (11.9)	
70–74 yr	329/6560 (5.0)	337/6571 (5.1)	
≥75 yr	40/6560 (0.6)	47/6571 (0.7)	
Pack-yr of smoking‡			
Median (IQR)	38.0 (29.7–49.5)	38.0 (29.7–49.5)	0.26
Range	0.4–159.5	1.3–156.0	
Cigarettes smoked per day — no./total no. (%)			
			0.69
≤10	20/6565 (0.3)	18/6596 (0.3)	
11–15	1470/6565 (22.4)	1437/6596 (21.8)	
16–20	1859/6565 (28.3)	1859/6596 (28.2)	
21–25	1732/6565 (26.4)	1779/6596 (27.0)	
26–30	669/6565 (10.2)	723/6596 (11.0)	
31–40	454/6565 (6.9)	437/6596 (6.6)	
>40	361/6565 (5.5)	343/6596 (5.2)	
Duration of smoking — no./total no. (%)			
			0.03
≤25 yr	25/6563 (0.4)	21/6594 (0.3)	
26–30 yr	657/6563 (10.0)	722/6594 (10.9)	
31–35 yr	1652/6563 (25.2)	1700/6594 (25.8)	
36–40 yr	2030/6563 (30.9)	2105/6594 (31.9)	
41–45 yr	1451/6563 (22.1)	1317/6594 (20.0)	
≥45 yr	748/6563 (11.4)	729/6594 (11.1)	
Age at initiation of smoking — no./total no. (%)			
			0.54
<15 yr	1153/6560 (17.6)	1141/6588 (17.3)	
15–29 yr	5376/6560 (82.0)	5407/6588 (82.1)	
≥30 yr	31/6560 (0.5)	40/6588 (0.6)	
Smoking status — no./total no. (%)			
			0.40
Current	3643/6566 (55.5)	3611/6595 (54.8)	
Former	2923/6566 (44.5)	2984/6595 (45.2)	
Years since cessation of smoking — no./total no. (%)			
			0.72
<1	489/2908 (16.8)	493/2963 (16.6)	
1–5	1316/2908 (45.3)	1334/2963 (45.0)	
6–10	1054/2908 (36.2)	1096/2963 (37.0)	
>10	49/2908 (1.7)	40/2963 (1.3)	

* Percentages may not total 100 because of rounding. IQR denotes interquartile range.

† The trial was designed for persons 50 to 74 years of age. Some men who were younger or older than the birth cohort that was approached underwent randomization and were included in the analysis.

‡ Some men who had a lower smoking history than the inclusion criterion underwent randomization and were included in the analysis.

Table 4. Cause of Death of Deceased Male Participants at 10 Years of Follow-up or until the Data-Cutoff Date of December 31, 2015.*

Variable	Screening Group (N = 868)	Control Group (N = 860)	Total (N = 1728)	Rate Ratio (95% CI)
	<i>number (percent)</i>			
Cause of death — no. (%)				
Lung cancer	160 (18.4)	210 (24.4)	370 (21.4)	0.76 (0.62–0.94)
No lung cancer after cause-of-death review, no other specification	6 (0.7)	11 (1.3)	17 (1.0)	0.55 (0.17–1.61)
Other neoplasm	318 (36.6)	289 (33.6)	607 (35.1)	1.10 (0.94–1.30)
Cardiovascular disease	189 (21.8)	181 (21.0)	370 (21.4)	1.05 (0.85–1.29)
Respiratory disease	42 (4.8)	43 (5.0)	85 (4.9)	0.98 (0.62–1.53)
Symptoms, signs, and abnormal clinical and laboratory findings, not elsewhere classified	37 (4.3)	20 (2.3)	57 (3.3)	1.86 (1.05–3.37)
Diseases of the digestive system	30 (3.5)	21 (2.4)	51 (3.0)	1.43 (0.79–2.63)
External causes of illness and death	24 (2.8)	19 (2.2)	43 (2.5)	1.27 (0.67–2.45)
Endocrine, nutritional, and metabolic diseases	21 (2.4)	9 (1.0)	30 (1.7)	2.34 (1.03–5.80)
Diseases of the nervous system	9 (1.0)	19 (2.2)	28 (1.6)	0.48 (0.19–1.10)
Other cause of death	26 (3.0)	28 (3.3)	54 (3.1)	0.93 (0.52–1.65)
Unknown	6 (0.7)	10 (1.2)	16 (0.9)	0.60 (0.18–1.83)
Total person-yr at risk	62,298	62,484	124,782	
All-cause mortality — deaths per 1000 person-yr	13.93	13.76	13.85	1.01 (0.92–1.11)

* Percentages may not total 100 because of rounding.

Queries

q1. AU: OK as edited?

q2. AU: OK as edited?

q3. AU: Correct?

Data Sharing Statement

de Koning HJ, van der Aalst CM, de Jong PA, et al. Reduced Lung-Cancer Mortality with Volume CT Screening in a Randomized Trial. N Engl J Med. DOI: 10.1056/NEJMoa1911793.

Question	Authors' Response
Will the data collected for your study be made available to others?	Yes
Would you like to offer context for your decision?	A data advisory board will review requests.
Which data?	—
Additional information about data	—
How or where can the data be obtained?	—
When will data availability begin?	—
When will data availability end?	—
Will any supporting documents be available?	—
Which supporting documents?	—
Additional information about supporting documents	—
How or where can supporting documents be obtained?	—
When will supporting documents availability begin?	—
When will supporting documents availability end?	—
To whom will data be available?	—
For what type of analysis or purpose?	—
By what mechanism?	—
Any other restrictions?	—
Additional information	—

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