



2 Prediction of contralateral breast cancer: external validation of risk
3 calculators in 20 international cohorts

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18 **Abstract**

19 **Background** Three tools are currently available to predict the risk of contralateral breast cancer (CBC). We aimed to compare
20 the performance of the Manchester formula, CBCrisk, and PredictCBC in patients with invasive breast cancer (BC).

21 **Methods** We analyzed data of 132,756 patients (4682 CBC) from 20 international studies with a median follow-up of
22 8.8 years. Prediction performance included discrimination, quantified as a time-dependent Area-Under-the-Curve (AUC)
23 at 5 and 10 years after diagnosis of primary BC, and calibration, quantified as the expected-observed (E/O) ratio at 5 and
24 10 years and the calibration slope.

25 **Results** The AUC at 10 years was: 0.58 (95% confidence intervals [CI] 0.57–0.59) for CBCrisk; 0.60 (95% CI 0.59–0.61)
26 for the Manchester formula; 0.63 (95% CI 0.59–0.66) and 0.59 (95% CI 0.56–0.62) for PredictCBC-1A (for settings where
27 *BRCA1/2* mutation status is available) and PredictCBC-1B (for the general population), respectively. The E/O at 10 years:
28 0.82 (95% CI 0.51–1.32) for CBCrisk; 1.53 (95% CI 0.63–3.73) for the Manchester formula; 1.28 (95% CI 0.63–2.58) for
29 PredictCBC-1A and 1.35 (95% CI 0.65–2.77) for PredictCBC-1B. The calibration slope was 1.26 (95% CI 1.01–1.50) for
30 CBCrisk; 0.90 (95% CI 0.79–1.02) for PredictCBC-1A; 0.81 (95% CI 0.63–0.99) for PredictCBC-1B, and 0.39 (95% CI
31 0.34–0.43) for the Manchester formula.

32 **Conclusions** Current CBC risk prediction tools provide only moderate discrimination and the Manchester formula was poorly
33 calibrated. Better predictors and re-calibration are needed to improve CBC prediction and to identify low- and high-CBC
34 risk patients for clinical decision-making.

35 **Keywords** Contralateral breast cancer · Risk prediction · Validation · Clinical decision-making

Introduction

A rising number of women with breast cancer (BC) are at risk to develop a new primary tumor in the contralateral breast (CBC) with consequently another cancer treatment and potentially less favorable prognosis [1]. Although CBC incidence is low (~0.4% per year) in the general BC

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population, contralateral preventive mastectomy (CPM) is increasing, also among women with low-CBC risk [2–5].

Three tools are currently available to predict the risk of CBC, although probably none are widely used: (1) the Manchester formula; (2) CBCrisk, and (3) PredictCBC [6–8]. The Manchester group in the United Kingdom (UK) proposed a set of guidelines for counseling women about CPM [8]. Based on a systematic review of the literature, they devised a formula to estimate lifetime CBC risk based on age at first primary BC, family history of BC, estrogen-receptor (ER) status, diagnosis of ductal carcinoma in situ (DCIS), and oophorectomy.

The second tool, CBCrisk, was developed using data on 1921 CBC cases and 5763 matched controls with primary BC [7]. The model uses data on age at first BC diagnosis, age at first birth, first degree family history of BC, high-risk pre-neoplasia, breast density (obtained using the BI-RADS system), ER status, first BC type (pure invasive, pure DCIS, a mix of the two, unknown), and adjuvant endocrine therapy. External validation was performed using two independent studies in the United States (US) of 5185 and 6035 patients with 111 and 117 CBC events [7, 9]. A web-based application provides individualized prediction of CBC risk [10].

Third, PredictCBC was developed, cross-validated and evaluated using data from 132,756 patients with first BC and 4672 CBC events, as part of an international collaboration [5]. PredictCBC predicts CBC risk as a function of family history (first degree) of primary BC, and information of primary BC diagnosis: age, nodal status, size, grade, morphology, ER status, human epidermal growth factor receptor 2 (HER2) status, administration of adjuvant or neoadjuvant chemotherapy, adjuvant endocrine therapy, adjuvant trastuzumab therapy, and radiotherapy. Two versions were developed: PredictCBC version 1A includes presence or absence of a mutation in the *BRCA1* or *BRCA2* genes, an important determinant of CBC [5, 11, 12], while PredictCBC version 1B was developed for untested patients.

External validation in different studies is relevant to assess the prediction performance of prediction models [13]. Our aim was to perform a head-to-head comparison between CBCrisk, PredictCBC and the Manchester formula. We hereto used several large population- and hospital-based studies used to develop and cross-validate the PredictCBC models.

Material and methods

External validation of CBCrisk and the Manchester formula was performed in 20 studies: four with individual patient data from the Netherlands [the Amsterdam Breast Cancer Study (ABCS), the Breast Cancer Outcome Study of Mutation carriers (BOSOM), the Erasmus MC Breast Cancer

Registry (EMC), the Netherlands Cancer Registry (NCR)]; and 16 other studies of the Breast Cancer Association Consortium (BCAC). The latter is an international consortium of 102 studies comprising 182,898 patients (data version: January 2017) with a primary BC diagnosed between 1939 and 2016 [14]. Of these, 16 non-familial BC BCAC studies including invasive non-metastatic European-descent female patients with first primary invasive BC diagnosed from 1990 onwards, and with at least 10 CBC events, were included in the analyses [14]. Details about studies and patient selection, and data imputation were described previously [5].

The outcome was in situ or invasive metachronous CBC. Follow-up started 3 months after invasive first primary BC diagnosis, to exclude synchronous CBCs, and ended at date of CBC, distant metastasis (but not at loco-regional relapse), CPM or last date of follow-up (due to death, being lost to follow-up, or end of study), whichever occurred first. In the BCAC, 27,155 patients were recruited more than 3 months after diagnosis of the first primary BC (prevalent cases); for these patients, follow-up started at date of recruitment (left truncation). Distant metastasis and death due to any cause were competing events.

The Manchester formula provides an estimate of a woman's individual lifetime CBC risk. To assess the prediction performance, we translated the lifetime CBC risk to 5- and 10-year CBC risks (see Supplementary Material). The predictors included in the CBC risk estimation in the Manchester formula, CBCrisk and PredictCBC models are provided in Table 1. Predictors that were sporadically missing were multiply imputed as described elsewhere [5].

Statistical analysis

Discrimination, the ability of the model to differentiate between patients who experienced CBC and those who did not, was calculated by time-dependent Area-Under-the-Curve (AUCs) based on Inverse Censoring Probability Weighting at 5 and 10 years [15, 16]. Values of AUCs close to 1 indicate good discrimination while values close to 0.5 indicate poor discrimination (a coin flip). Calibration is the agreement between observed and predicted risk and is commonly characterized by calibration-in-the-large and slope statistic. Calibration-in-the-large characterizes the overall difference between the observed and predicted risks. It was calculated using the expected/observed (E/O) ratio. An E/O less than 1 indicates that the model systematically underestimates CBC risk, while an E/O above 1 indicates that the model systematically overestimates CBC risk. The expected number of cases was calculated by summing the individual predicted probabilities at 5 and 10 years, based on the patient-specific covariate values [17]. The observed number of cases was estimated by the non-parametric CBC cumulative incidence at 5 and 10 years. The calibration slope was

Table 1 Predictors included in current contralateral breast cancer risk prediction tools

| List of predictors | CBCrisk ^b | Manchester formula ^c | PredictCBC version 1A ^d | PredictCBC version 1B ^d |
|--|----------------------|---------------------------------|------------------------------------|------------------------------------|
| Age at diagnosis | ✓ | ✓ | ✓ | ✓ |
| Age at first birth | ✓ | | | |
| First-degree family history | ✓ | ✓ | ✓ | ✓ |
| <i>BRCA1/2</i> germline mutation | | ✓ | ✓ | |
| First breast cancer behavior type ^a | ✓ | ✓ | | |
| Lymph node status | | | ✓ | ✓ |
| Breast density | ✓ | | | |
| Tumor size | | | ✓ | ✓ |
| Morphology | | | ✓ | ✓ |
| Tumor grade | | | ✓ | ✓ |
| High-risk pre-neoplasia | ✓ | | | |
| ER status | ✓ | ✓ | ✓ | ✓ |
| HER2 status | | | ✓ | ✓ |
| Chemotherapy | | | ✓ | ✓ |
| Endocrine therapy | ✓ | | ✓ | ✓ |
| Radiation to the breast | | | ✓ | ✓ |
| Trastuzumab | | | ✓ | ✓ |
| Oophorectomy under 40 years | | ✓ | | |

ER estrogen-receptor status, HER2 human epidermal growth factor receptor 2

^aContralateral breast cancer risk was calculated including women diagnosed with ductal carcinoma in situ

^bChowdhury et al. [7]

^cBasu et al. [8]

^dGiardiello et al. [5]

143 estimated using a Fine and Gray regression model using the
144 linear predictor of the prediction tools. The linear predictor
145 was vs constructed as the sum of the factors included in each
146 model weighted by the corresponding regression coefficients
147 (or parameters), and then computed in the validation dataset
148 exactly as reported for the development set. The calibration
149 slope is determined as the regression coefficient for this lin-
150 ear predictor when fitted as a single covariate in a regres-
151 sion model of disease outcome in the validation dataset. A
152 well-calibrated model should have a calibration slope of 1;
153 slopes < 1 indicate that coefficients were too optimistic for
154 the validation setting [18]. Calibration results were graphi-
155 cally displayed.

156 Analyses were stratified by geographic groups of studies,
157 since stratification by individual studies would provide too
158 few events in some strata [5, 13, 19]. To allow for heteroge-
159 neity across multiple studies, random-effect meta-analyses
160 were performed. We calculated 95% confidence intervals
161 (CI) and 95% prediction intervals (PI), which indicate the
162 likely range for prediction accuracy of the model in a new
163 dataset, for discrimination and calibration measures. A sen-
164 sitivity analysis was performed to check the consistency of
165 CBCrisk performance measures when metachronous CBC
166 was defined as an event after 6 instead of 3 months since
167 the first BC diagnosis. More details are provided in the

Supplementary Material. All analyses were implemented
using SAS (SAS Institute Inc., NC, USA) and R software
[20].

Results

We included 132,756 patients from 20 studies who experi-
enced 4862 CBC events during a median follow-up of
8.8 years. The main patient and clinical characteristics
across studies and geographic areas are shown in Table 2.

The AUCs at 5 and 10 years was around 0.6: 0.59 (95% CI
0.57–0.61; 95% PI 0.54–0.64) and 0.58 (95% CI 0.57–0.59;
95% PI 0.55–0.61) for CBCrisk (Fig. 1); 0.61 (95% CI
0.60–0.62; 95% PI 0.59–0.63) and 0.60 (95% CI 0.59–0.61;
95% PI 0.58–0.62) for the Manchester formula (Fig. 2). The
E/O ratio at 5 and 10 years was close to 1 for all models:
0.86 (95% CI 0.50–1.46; 95% PI 0.20–3.75) and 0.82 (95%
CI 0.51–1.32; 95% PI 0.21–3.14) for CBCrisk (Table 3);
1.54 (95% CI 0.61–3.92; 95% PI 0.11–20.72, Table 4),
and 1.53 (95% CI 0.63–3.73; 95% PI 0.13–18.52) for the
Manchester formula (Table 4); 1.26 (95% CI 0.57–2.77;
95% PI 0.14–11.34), and 1.28 (95% CI 0.63–2.58; 95% PI
0.18–9.18) for PredictCBC-1A (Table 5); 1.33 (95% CI
0.59–2.99, 95% PI 0.14–12.76), 1.35 (95% CI 0.65–2.77;

Table 2 Description of main patient and clinical factors used for evaluation of the models and formula

| Study ^a /geographic area | Europe—other ^b | Europe—Scandinavia | Europe—United Kingdom | Netherlands—BOSOM | Netherlands—EMC | Netherlands—NCR | United States and Australia |
|---------------------------------------|---------------------------|--------------------|-----------------------|-------------------|-----------------|-----------------|-----------------------------|
| <i>N</i> | 15,183 | 12,928 | 11,921 | 3760 | 3390 | 83,138 | 2436 |
| Age at first diagnosis, years (%) | | | | | | | |
| < 30 | 152 (1.0) | 46 (0.4) | 156 (1.3) | 108 (2.9) | 46 (1.4) | 388 (0.5) | 41 (1.7) |
| 30–39 | 1252 (8.2) | 489 (3.8) | 1811 (15.2) | 842 (22.4) | 374 (11.0) | 4241 (5.1) | 494 (20.3) |
| 40+ | 13,779 (90.8) | 12,393 (95.9) | 9954 (83.5) | 2810 (74.7) | 2970 (87.6) | 78,509 (94.4) | 1901 (78.0) |
| Age at first birth = unknown (%) | 15,183 (100.0) | 12,928 (100.0) | 11,921 (100.0) | 3760 (100.0) | 3390 (100.0) | 83,138 (100.0) | 2436 (100.0) |
| Family history (%) | | | | | | | |
| Yes | 2123 (14.0) | 818 (6.3) | 1371 (11.5) | 737 (19.6) | 591 (17.4) | 0 (0.0) | 319 (13.1) |
| No | 8057 (53.1) | 3158 (24.4) | 8210 (68.9) | 1177 (31.3) | 2482 (73.2) | 0 (0.0) | 1498 (61.5) |
| Unknown | 5003 (33.0) | 8952 (69.2) | 2340 (19.6) | 1846 (49.1) | 317 (9.4) | 83,138 (100.0) | 619 (25.4) |
| First BC type = Pure invasive (%) | 15,183 (100.0) | 12,928 (100.0) | 11,921 (100.0) | 3760 (100.0) | 3390 (100.0) | 83,138 (100.0) | 2436 (100.0) |
| Breast density = unknown (%) | 15,183 (100.0) | 12,928 (100.0) | 11,921 (100.0) | 3760 (100.0) | 3390 (100.0) | 83,138 (100.0) | 2436 (100.0) |
| ER status (%) | | | | | | | |
| Negative | 3387 (22.3) | 1746 (13.5) | 1718 (14.4) | 896 (23.8) | 842 (24.8) | 14,591 (17.6) | 445 (18.3) |
| Positive | 10,071 (66.3) | 9401 (72.7) | 7175 (60.2) | 2024 (53.8) | 2427 (71.6) | 64,790 (77.9) | 1572 (64.5) |
| Unknown | 1725 (11.4) | 1781 (13.8) | 3028 (25.4) | 840 (22.3) | 121 (3.6) | 3757 (4.5) | 419 (17.2) |
| High-risk pre-neoplasia = unknown (%) | 15,183 (100.0) | 12,928 (100.0) | 11,921 (100.0) | 3760 (100.0) | 3390 (100.0) | 83,138 (100.0) | 2436 (100.0) |
| Anti-estrogen therapy (%) | | | | | | | |
| Yes | 7868 (51.8) | 6434 (49.8) | 8712 (73.1) | 809 (21.5) | 1559 (46.0) | 40,214 (48.4) | 363 (14.9) |
| No | 4570 (30.1) | 1947 (15.1) | 2046 (17.2) | 2739 (72.8) | 1821 (53.7) | 42,924 (51.6) | 8 (0.3) |
| Unknown | 2745 (18.1) | 4547 (35.2) | 1163 (9.8) | 212 (5.6) | 10 (0.3) | 0 (0.0) | 2065 (84.8) |
| CBC cumulative incidence (%) | | | | | | | |
| 3-year (95% CI) | 1.0 (0.8–1.2) | 0.7 (0.5–0.9) | 0.5 (0.3–0.7) | 1.7 (1.3–2.1) | 1.7 (1.2–2.1) | 1.3 (1.2–1.4) | 1.8 (0.8–2.8) |
| 5-year (95% CI) | 1.6 (1.4–1.9) | 1.0 (0.8–1.3) | 1.0 (0.8–1.3) | 3.0 (2.5–3.6) | 2.6 (2.1–3.2) | 2.4 (2.3–2.5) | 2.8 (1.7–3.8) |
| 10-year (95% CI) | 3.5 (3.1–3.9) | 2.1 (1.7–2.4) | 1.3 (1.0–1.5) | 5.5 (4.7–6.2) | 5.7 (4.9–6.6) | 4.6 (4.5–4.8) | 4.1 (3.0–5.3) |

More details about the main patient and clinical characteristics by study are available in the supplementary information of [5]

BOSOM Breast Cancer Outcome Study of Mutation carriers, *EMC* Erasmus Medical Center, *NCR* Netherlands Cancer Registry, *BC* breast cancer, *ER* estrogen receptor, *CBC* contralateral breast cancer, *CI* confidence interval

^aThe studies denoted with Europe and United States and Australia are part of the Breast Cancer Association Consortium

^bEurope—other geographic area included studies from Belgium (1), Germany (2), Netherlands (2) and Poland (2)

190 95% PI 0.19–10.24) for PredictCBC-1B (Table 5) [5]. The
 191 calibration slope was close to 1 for CBCrisk (1.26, 95%
 192 CI 1.01–1.50 and 95% PI 1.01–1.50, Tables 3, 4, 5), and
 193 PredictCBC-1A and 1B 0.90 (95% CI 0.79–1.02; 95% PI
 194 0.73–1.08), and 0.81 (95% CI 0.63–0.99; 95% PI 0.50–1.12)
 195 (Table 5), while prognostic effects were far too large for the
 196 Manchester formula (slope: 0.39, 95% CI 0.34–0.43, 95%
 197 PI 0.34–0.43, Tables 4, 5). Calibration plots of CBCrisk
 198 at 5 and 10 years are shown in Supplementary Fig. 1 and

Supplementary Fig. 2. As reported previously [5], the AUCs
 at 5 and 10 years for PredictCBC-1A were 0.63 (95% CI
 0.58–0.67, 95% PI 0.52–0.74), and 0.63 (95% CI 0.59–0.66,
 95% PI 0.53–0.72), respectively; for PredictCBC-1B 0.59
 (CI 0.54–0.63, 95% PI 0.46–0.71, Table 5), and 0.59 (95%
 CI 0.56–0.62, 95% PI 0.52–0.66, Table 5), respectively.

Sensitivity analysis showed that the performance meas-
 ures of CBCrisk did not change when metachronous CBC
 was defined after 6 months since first BC diagnosis (see

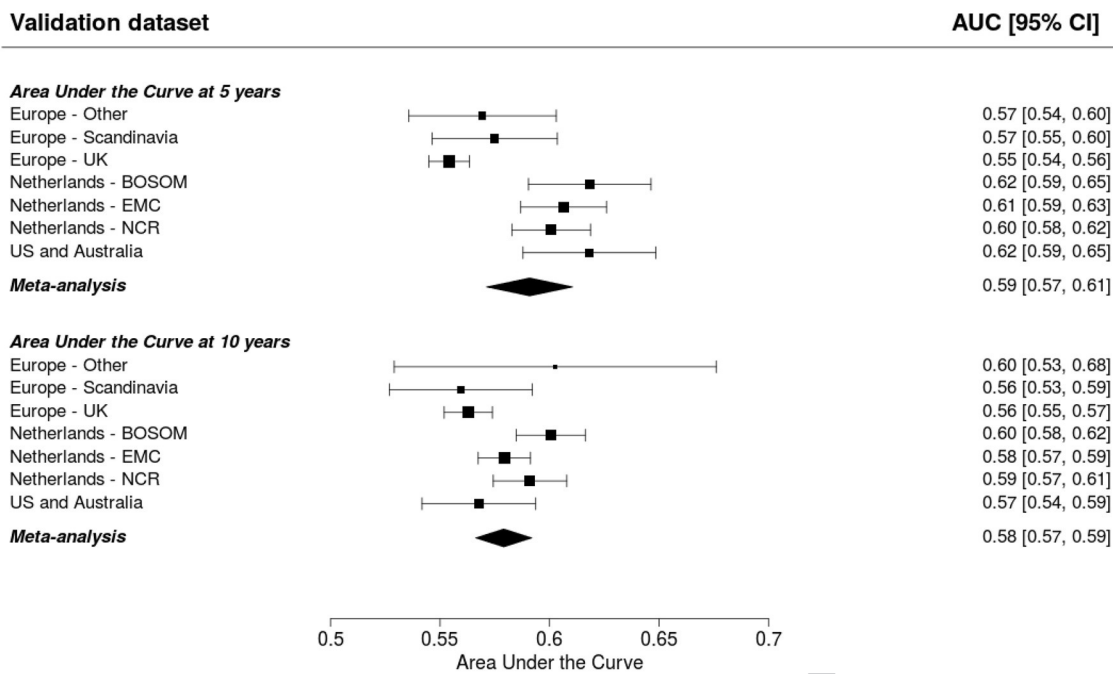


Fig. 1 Prediction performance of the CBCrisk model (Chowdhury et al. [7]). The upper and lower panel show the discrimination assessed by a time-dependent Area-Under-the-Curve at 5 and 10 years, respectively. The black squares indicate the estimated accuracy of a model built on all remaining studies or geographic areas.

The black horizontal lines indicate the corresponding 95% confidence intervals of the estimated accuracy (interval whiskers). The black diamonds indicate the mean with the corresponding 95% confidence interval of the predictive accuracy

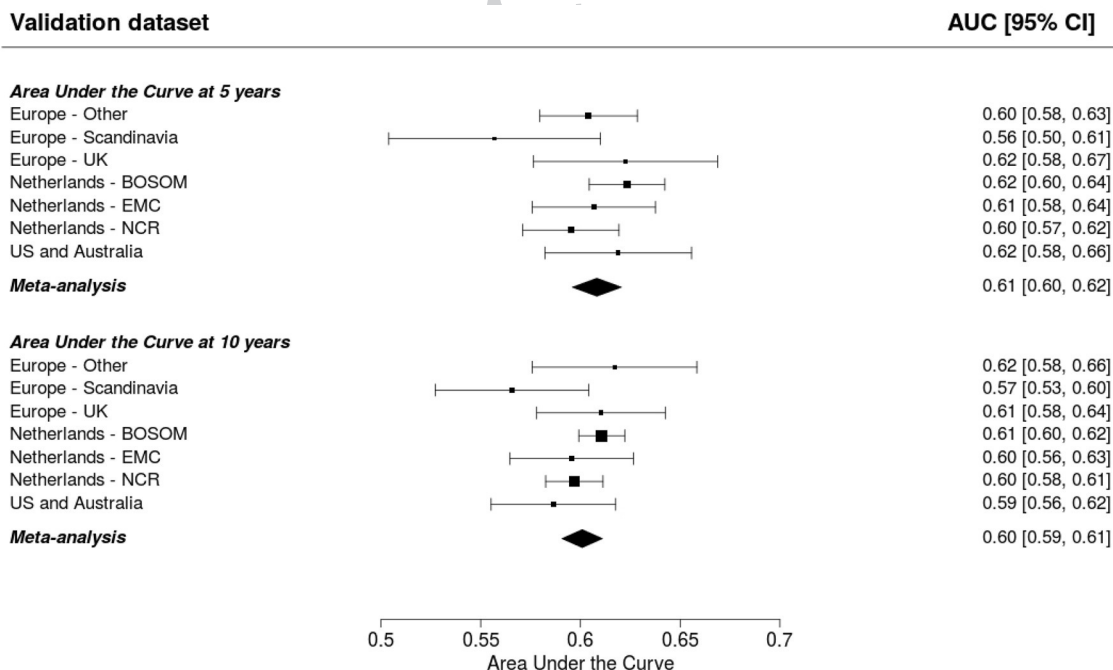


Fig. 2 Prediction performance of the Manchester formula (Basu et al. [8]) The upper and lower panel show the discrimination assessed by a time-dependent Area-Under-the-Curve at 5 and 10 years, respectively. The black squares for each dataset indicate the estimated accuracy of a model built on all remaining studies or geographic areas. The black

horizontal lines indicate the corresponding 95% confidence intervals of the estimated accuracy (interval whiskers). The black diamonds indicate the mean with the corresponding 95% confidence interval of the predictive accuracy

Table 3 Calibration performance of the CBC risk model

| Validation dataset | E/O ratio at 5 years (95% CI) | E/O ratio at 10 years (95% CI) | Calibration slope (95% CI) |
|--------------------|-------------------------------|--------------------------------|----------------------------|
| Europe—Other | 0.87 (0.76 to 0.98) | 0.75 (0.68 to 0.81) | 1.11 (0.40 to 1.83) |
| Europe—Scandinavia | 1.59 (1.28 to 1.91) | 1.23 (1.08 to 1.38) | 0.86 (0.16 to 1.57) |
| Europe—UK | 1.35 (1.38 to 2.17) | 1.82 (1.53 to 2.11) | 0.85 (− 0.03 to 1.73) |
| Netherlands—BOSOM | 0.45 (0.37 to 0.53) | 0.50 (0.43 to 0.57) | 1.34 (0.76 to 1.93) |
| Netherlands—EMC | 0.48 (0.38 to 0.57) | 0.43 (0.37 to 0.50) | 1.19 (0.65 to 1.73) |
| Netherlands—NCR | 0.57 (0.54 to 0.59) | 0.54 (0.52 to 0.56) | 1.40 (1.11 to 1.68) |
| US and Australia | 0.43 (0.33 to 0.54) | 0.56 (0.45 to 0.67) | 1.13 (0.25 to 2.00) |
| Meta-analysis | 0.86 (0.50 to 1.46) | 0.82 (0.51 to 1.32) | 1.26 (1.01 to 1.50) |
| 95% PI | 0.20 to 3.75 | 0.21 to 3.14 | 1.01 to 1.50 |

Chowdhury et al. [7]

E/O expected-observed, CI confidence interval, UK United Kingdom, BOSOM Breast Cancer Outcome Study of Mutation carriers, EMC Erasmus Medical Center, NCR Netherlands Cancer Registry, PI prediction interval

Table 4 Calibration performance of the Manchester formula

| Validation dataset | E/O ratio at 5 years (95% CI) | E/O ratio at 10 years (95% CI) | Calibration slope (95% CI) |
|--------------------|-------------------------------|--------------------------------|----------------------------|
| Europe—Other | 1.64 (1.44 to 1.85) | 1.46 (1.34 to 1.58) | 0.40 (0.29 to 0.50) |
| Europe—Scandinavia | 2.61 (2.09 to 3.12) | 2.11 (1.85 to 2.37) | 0.35 (0.13 to 0.57) |
| Europe—UK | 3.34 (2.60 to 4.08) | 3.49 (2.93 to 4.05) | 0.42 (0.23 to 0.61) |
| Netherlands—BOSOM | 0.81 (0.66 to 0.96) | 0.92 (0.79 to 1.05) | 0.45 (0.33 to 0.56) |
| Netherlands—EMC | 0.94 (0.75 to 1.14) | 0.87 (0.75 to 1.00) | 0.35 (0.21 to 0.49) |
| Netherlands—NCR | 1.00 (0.95 to 1.04) | 1.01 (0.98 to 1.05) | 0.37 (0.33 to 0.42) |
| US and Australia | 0.77 (0.58 to 0.96) | 1.02 (0.82 to 1.23) | 0.51 (0.33 to 0.68) |
| Meta-analysis | 1.54 (0.61 to 3.92) | 1.53 (0.63 to 3.73) | 0.39 (0.34 to 0.43) |
| 95% PI | 0.11 to 20.72 | 0.13 to 18.52 | 0.34 to 0.43 |

Basu et al. [8]

E/O expected-observed, CI confidence interval, UK United Kingdom, BOSOM Breast Cancer Outcome Study of Mutation carriers, EMC Erasmus Medical Center, NCR Netherlands Cancer Registry, PI prediction interval

Table 5 Summary of prediction performance of CBCrisk, Manchester formula, and PredictCBC version 1A and version 1B with the corresponding 95% prediction intervals (PI)

| Characteristics | CBCrisk ^a | Manchester formula ^b | PredictCBC version 1A ^{c,d} | PredictCBC version 1B ^{c,d} |
|--------------------------------|----------------------|---------------------------------|--------------------------------------|--------------------------------------|
| Discrimination | | | | |
| AUC at 5 years (95% PI) | 0.59 (0.54 to 0.64) | 0.61 (0.59 to 0.63) | 0.63 (0.52 to 0.74) | 0.59 (0.46 to 0.71) |
| AUC at 10 years (95% PI) | 0.58 (0.55 to 0.61) | 0.60 (0.58 to 0.62) | 0.63 (0.53 to 0.72) | 0.59 (0.52 to 0.66) |
| Calibration | | | | |
| E/O ratio at 5 years (95% PI) | 0.86 (0.20 to 3.75) | 1.54 (0.11 to 20.72) | 1.26 (0.14 to 11.34) | 1.33 (0.14 to 12.76) |
| E/O ratio at 10 years (95% PI) | 0.82 (0.21 to 3.14) | 1.53 (0.13 to 18.52) | 1.28 (0.18 to 9.18) | 1.35 (0.19 to 10.24) |
| Slope (95% PI) | 1.26 (1.01 to 1.50) | 0.39 (0.34 to 0.43) | 0.90 (0.73 to 1.08) | 0.81 (0.50 to 1.12) |

AUC Area under the curve, PI prediction interval

^aChowdhury et al. [7]

^bBasu et al. [8]

^cGiardiello et al. [5], Fig. 1 and Figure S5

^dversion 1A includes BRCA mutation status as a variable while 1B does not

208 Supplementary Materials, Supplementary Tables 1–2 and
209 Supplementary Fig. 3).

210 Discussion

211 Accurate CBC risk predictions are essential in clinical deci-
212 sion-making around CPM or tailored surveillance among
213 patients with first primary BC. In particular, overestimation
214 of risk can lead to recommending CPM among BC patients
215 with low risks. Underestimation can lead to suboptimal
216 surveillance or hesitation about recommending CPM for
217 patients with substantial risk. Using individual patient data
218 from multiple studies with long follow-up, we externally
219 evaluated the prediction performance accuracy of CBCrisk,
220 a tool developed and validated to provide individualized
221 CBC risk prediction, and the Manchester formula, a heu-
222 ristically derived calculation of CBC lifetime risk [6–9]. In
223 addition, the availability of different European-descendent
224 studies allowed heterogeneity in the performance by geo-
225 graphic area to be assessed.

226 CBCrisk under-predicted the risk of CBC and had moder-
227 ate discrimination ability with considerable heterogeneity
228 between studies. The Manchester formula was empirically
229 derived from a systematic review, and its discrimination
230 accuracy was higher than CBCrisk. This may be explained
231 by the inclusion of *BRCA1/2* mutation carrier information,
232 an important determinant of CBC risk [21]. With the same
233 large individual patient data sets, PredictCBC models had
234 been developed and validated [5]. In particular, PredictCBC
235 version 1A includes information of *BRCA1/2* mutation carri-
236 ers and extensive information about the primary BC includ-
237 ing treatments. The discrimination of all three prediction
238 models was moderate, with AUC values around 0.6.

239 CBCrisk was previously externally validated using two
240 independent clinical studies from Johns Hopkins University
241 (JH) and MD Anderson Cancer Center (MDA) in the US [9].
242 Discrimination ability was 0.61 and 0.65 at 3 years, and 0.62
243 and 0.61 at 5 years for JH and MDA, respectively. The risk
244 of CBC was overestimated in JH with E/O ratios of 2.02 and
245 1.56 at 3 and 5 years, while underestimated in MDA with
246 E/O ratios of 0.61 and 0.62, respectively.

247 The considerable heterogeneity in all CBC risk calcula-
248 tors, especially in the CBCrisk and the Manchester formula,
249 reflects the different CBC incidences in every study [13].
250 Another potential source of heterogeneity is the carrier fre-
251 quency of germline mutations associated with CBC that may
252 vary among studies, especially in the CBC calculators not
253 including information of *BRCA1/2* mutation as CBCrisk and
254 the PredictCBC-1B [22]. In addition, heterogeneity may be
255 due to the different proportions of the use of (neo)adjuvant
256 systemic therapies explained by the different distribution of
257 tumor subtypes among studies [4]. Besides, inter-observer

258 variation in pathological examination of BC among stud-
259 ies may lead to different adjuvant systemic therapy advice
260 and, consequently, prediction of CBC risk [23]. Variation
261 in prediction performance and limited generalizability of
262 CBC risk calculators can also be partially explained by dif-
263 ferences in how predictors are measured among studies [24,
264 25]. For example, lack of family history knowledge may
265 lead to uncertainty in risk prediction and varies according
266 to demographics of the patients [26]. In particular, if in some
267 studies BC patients misreported information about family
268 history, the CBC risk would be over(under)estimated caus-
269 ing inappropriate decision-making regarding CPM or tai-
270 lored surveillance. Some limitations of our study must be
271 recognized. First, our dataset, while large, had missing data
272 for three covariates that were used in the CBCrisk model:
273 breast density, age at first birth, and high-risk pre-neoplas-
274 ia. The authors of CBCrisk estimated the relative risks for
275 patients with the unknown characteristics, but the use of
276 the missing indicator variable is suboptimal compared to
277 having the prognostic information available. It may lead to
278 over or under-estimation of absolute CBC risk [27]. For this
279 reason, we suggest that it is preferable to use multiple im-
280 putation of missing data, as is done in the PredictCBC models
281 [28, 29]. In addition, investigation of the potential source of
282 model misspecification due to possible different definitions
283 or measurement error was not possible [30–32].

284 In conclusion, current statistical risk prediction mod-
285 els and heuristic formulas provided moderate CBC indi-
286 vidualized prediction performance. Careful re-calibration
287 is required before considering these models for clinical
288 decision-making. A more direct comparison between the
289 current CBC risk prediction models using a large external
290 dataset with complete information on all factors included in
291 all CBC prediction models would be ideal, but is currently
292 unavailable. There is an ongoing debate about improvements
293 of clinical prediction performance using machine learning
294 approaches compared to standard regression approaches for
295 risk prediction [33, 34]. However, irrespective of the meth-
296 odology, better predictors are needed to predict CBC more
297 accurately. Deeper biological insights and potential inclusion
298 of other genetic markers such as *CHEK2* c.1100del mutation
299 status and polygenic risk scores based on common genetic
300 variants may improve CBC risk prediction, although rare
301 mutations are unlikely to contribute substantially to CBC
302 risk in the general population [35, 36]. Life-style factors
303 such as body mass index, alcohol consumption, and smok-
304 ing also may help to better stratify high- and low-CBC risk
305 patients even though these factors are difficult to measure
306 accurately. Moreover, breast density may be important. More
307 detailed information about adjuvant systemic therapies may
308 better identify patients with low- and high-CBC risk since
309 chemotherapy and especially endocrine therapy reduce CBC
310 risk [4]. After extension and further external validation of

311 prediction models for CBC risk, investigation of their poten-
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Compliance with ethical standards

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