

# Milk homogenization monitoring: fat globule size estimation from scattering spectra of milk

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## Abstract

The fat globule size distribution of raw milk and milk with an increasing degree of homogenization was estimated based on their bulk light scattering properties in the Vis/NIR wavelength range. The particle size distribution (PSD) was approximated as a lognormal distribution, of which the parameters were estimated simultaneously with the fat concentration. This resulted in a good agreement between the estimated PSDs and the reference PSDs obtained by laser diffraction in case of raw and strongly homogenized samples. The accuracy increased if a known fat concentration was incorporated, or when the scattering coefficient and anisotropy factor spectra were used simultaneously as input. For mildly homogenized samples, the lognormal distribution was unable to fit the bimodal PSD correctly and focused on the largest fat globules. In this case, the estimated PSDs provided still relatively accurate information on D90, D32 and the right distribution tail, which contains the largest fat globules.

*Industrial relevance:* The presented estimation method demonstrates the potential of bulk scattering spectra for determining the PSD and concentration of scattering particles in turbid media. Further development of this technology can lead to new solutions for spectroscopic PSD determination, allowing on-line monitoring systems for a wide range of food and non-food products.

**Keywords:** Particle size distribution, milk homogenization, light scattering properties

## 1 Introduction

The composition of milk is an important characteristic with regard to food quality and further processing into derived products. The concentrations of the main components besides water, namely fat, protein and lactose, are often optically determined based on their light absorption characteristics in the infrared wavelength region (Fox & McSweeney, 1998; Walstra, Jenness, & Badings, 1984; Lynch, Barbano, Schweisthal, & Fleming, 2006). The infrared light is preferred because of the clear absorption bands and the lower influence of light scattering on the acquired spectra. In general, milk contains two types of scattering particles suspended in the milk serum: milk fat globules and casein micelles. The average fat content in bovine milk is about 4.0% w/w (range 2.5-5.5%), while casein is on average present at about 2.6% w/w (range 1.7-3.5%) (Walstra, Wouters, & Geurts, 2005). However, not only the concentrations play a role, but also the globule size determines the physical properties of milk. Fat globules in raw milk are reported to range from 0.1  $\mu\text{m}$  to 15-20  $\mu\text{m}$  diameter (Fox & McSweeney, 1998). Their size shows biological variation

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### Abbreviations

BOP	bulk optical properties	$d_{\text{mean}}$	mean diameter	UT	unscattered transmission
D...	...% percentile	LB	lower boundary	VF	volume fraction
D32	Sauter mean diameter	$par1$ & $par2$	generalized distribution parameters	$\mu_s$	scattering coefficient
D43	DeBrouckere mean diameter	PSD	particle size distribution	$\mu'_s$	reduced scattering coefficient
DIS	double integrating sphere	UB	upper boundary	$g$	anisotropy factor

31 with breed, age, health status of the animal etc. The fat globule size can be reduced by means of a  
32 homogenization process (Walstra, Geurts, Noomen, Jellema, & van Boekel, 1999), which stabilizes  
33 the milk against creaming and partial coalescence, and changes the viscosity (Walstra et al., 1999).  
34 Light scattering is wavelength dependent and influenced by the particle size. The smaller casein  
35 micelles, with a size ranging from 20 nm up to 400 nm (Walstra et al., 2005), particularly scatter  
36 light in the ultraviolet and visible wavelength range, while the larger fat globules have a scattering  
37 effect up to the infrared range.

38 To monitor the fat globule size before and after homogenization, measurements based on light  
39 spectroscopy, light scattering or electronic counting can be used (Fox & McSweeney, 1998; Bylund,  
40 2003). These are usually bench-top instruments that report a full particle size distribution (PSD).  
41 However, they only provide measurements at regular time intervals, at which a sample has to  
42 be transferred from the production line to the lab for preparation and analysis. Most methods  
43 also require severe dilution of the small sample volumes, which can alter the sample as aggregates  
44 might break up.

45 Several researchers have investigated the potential of Vis/NIR spectroscopy for particle sizing  
46 purposes. Bogomolov, Melenteva, and Dahm (2013) reported diffuse transmission spectra of in-  
47 creasingly homogenized milk in the 400-1000 nm wavelength range and used the representative  
48 layer theory to attribute the spectral changes to a decreasing fat globule size. The next step  
49 of extracting PSD information from spectra by inverse estimation was for example done by Di  
50 Marzo, Cree, and Barbano (2016). They estimated PSD parameters such as the mean and 90%  
51 quantile based on infrared spectra using partial least squares models. However, they were not able  
52 to predict a complete PSD and such data-based models can only be used on samples very similar  
53 to the ones used for training. On the other hand, Cabassi, Profaizer, Marinoni, Rizzi, and Catta-  
54 neo (2013) made PSD estimations based on NIR transmission spectra of raw milk by assuming a  
55 Weibull distribution as PSD shape. However, their validation was limited to the estimated Sauter  
56 mean diameters (D32).

57 Instead of these (semi-)empirical approaches, the underlying physics of light scattering can be  
58 used to access the particle size of turbid media like milk. Mie theory provides a direct relation  
59 between the size and scattering by spherical particles, such as fat globules. Aernouts et al. (2015b)  
60 found that the bulk optical properties (BOP) of milk are strongly linked to the size distribution  
61 of the milk fat globules. These BOP include the bulk scattering coefficient  $\mu_s$ , which indicates the  
62 probability of photon scattering per infinitesimal path length ( $[\mu\text{m}^{-1}]$ ), and the anisotropy factor  
63  $g$ , a measure for the direction of scattering ( $0 = \text{isotropic}$ ,  $1 = \text{completely forward}$ ). These two bulk  
64 scattering properties can be combined into the reduced scattering coefficient  $\mu'_s$ , according to the  
65 expression  $\mu'_s = \mu_s \times (1 - g)$ . Furthermore, we have shown that the PSDs of monomodal polystyrene  
66 suspensions can be estimated from the Vis/NIR BOP extracted from double integrating spheres  
67 measurements (Postelmans, Aernouts, & Saeys, 2019). This requires a robust inversion of the Mie  
68 scattering theory. Recently, Stocker et al. (2017) succeeded in estimating the milk fat globules  
69 PSD in raw and homogenized milk from the Vis/NIR scattering and reduced scattering coefficient  
70 spectra using an inversion of Mie theory. Nevertheless, they did not explore the potential of the  
71 Vis/NIR scattering anisotropy factor, as well as a combination of the different scattering properties  
72 to further improve the PSD estimation.

73 Therefore, the objective of this study is to evaluate the potential of the scattering anisotropy  
74 factor and the combination of scattering properties for estimating the fat globules size distribution  
75 in raw milk and an extensive set of milk samples with different degrees of homogenization. To  
76 improve the robustness, the estimation routine includes a procedure to obtain relevant starting  
77 points and to prevent local minima and non-converged solutions from being accepted as final PSD  
78 estimate. On top of that, it is investigated if including information on the concentration of the  
79 scatterers (the fat content), for example based on spectroscopic absorption measurements or  $\mu_a$   
80 (Aernouts, Polshin, Lammertyn, & Saeys, 2011; Aernouts, Polshin, Saeys, & Lammertyn, 2011;  
81 Aernouts et al., 2015a) can improve the accuracy of the PSD estimates. This would combine the  
82 benefits of extracting composition information from the absorption properties and particle size  
83 information from the light scattering properties.

## 2 Materials and methods

### 2.1 Milk samples and reference PSDs

A raw bulk milk sample was collected from a Belgian dairy research farm (Hooibeekhoeve, Geel). The cooled tank (at 4°C) contained the milk of 70 Holstein-Friesian dairy cows produced over two days and was regularly stirred. The fat and protein content of the milk (respectively 42.8 g/l and 33.7 g/l) was determined by the Milk Control Center Flanders with a MilkoScan FT+ (Foss, Hillerød, Denmark) according to ISO 9622:2013 (ISO, 2013). After warming up the sample to 37 °C and gently stirring, seven subsamples of 20 ml were taken. The subsamples were subjected to ultrasonic homogenization for respectively 0 s (raw milk), 30 s, 60 s, 120 s, 240 s, 480 s and 960 s. The sonication was performed using a Vibra-Cell VCX400 (400 W, 20 kHz) in combination with a CV26 converter, a 13 mm horn and a 3 mm microtip (Sonics & Materials Inc., Danbury, USA). The amplitude of the sonicator was set at 20% (80 W). The microtip was immersed about 1 cm into the milk sample, which was contained in a 50 ml conical polypropylene Falcon tube. To prevent overheating of the milk during homogenization, the samples were placed in a water bath of 24 °C for homogenization times up to 240 s, or 19 °C for longer homogenization times. Three-fold diluted samples (10 ml milk with 20 ml deionized water) were used in the optical measurements to ensure the independent scattering assumption was valid (Aernouts et al., 2015b).

The PSDs of the fat globules and casein micelles were determined using laser diffraction with a Mastersizer 3000 instrument (Malvern, UK). The milk was added drop-wise to a beaker of deionized water until a red laser obscuration of 5-9% was reached, with the rotor of the Hydro EV dispersion unit stirring at 2400 rpm. The resulting PSDs were the average of five consecutive measurements of 5 s without delay. The particle refractive index was set at  $1.46 + i5 \times 10^{-5}$ , as indicated in the Mastersizer software for milk fat. Particles were assumed to be spherical (Mie theory) and the general purpose analysis type was applied. The PSD of casein was identified as the first mode of the bimodal PSD of raw milk, with a clear separation from the second mode (milk fat). Subsequently, the PSD of casein was subtracted from all PSDs, and the resulting fat PSDs were converted to probability density functions.

### 2.2 Experimental BOP determination

To determine the BOP of the milk samples experimentally, the double integrating sphere (DIS) set-up combined with an unscattered transmission (UT) measurement path described by Aernouts et al. (2013) was used. A supercontinuum laser coupled to a monochromator illuminated the samples and allowed a sequential scan over the desired wavelength range. For more details on this set-up, the reader is referred to Aernouts et al. (2013), while more specific information on the milk measurements can be found in (Aernouts et al., 2015a, 2015b).

The milk samples were loaded into borosilicate cuvettes with a sample thickness of 0.55 mm and placed between the two integrating spheres to measure the total reflectance and total transmittance. To measure the unscattered transmittance, a 0.155 mm thick cuvette was placed in the UT path. Replicates were obtained by reloading the cuvettes five times from the same subsample. Measurements were performed in the wavelength range from 550 nm to 1350 nm in steps of 10 nm.

The acquired reflectance and transmittance spectra were passed on to the inverse adding doubling routine to calculate the BOP, as implemented by Prahl (2011). In addition to these spectra, the sample thickness and the sample refractive index were given as inputs. The refractive index of milk was calculated based on the weight fractions of the different milk components (Walstra et al., 1984) and the wavelength-dependent refractive index of water (Segelstein, 1981; Aernouts et al., 2014). Since the fat globules in raw milk are larger than 25% of the wavelength, they do not contribute to the refractive index (Walstra et al., 1984). For the homogenized samples, comparison of inverse adding-doubling results calculated based on a sample refractive index with and without a contribution of fat showed that the effect of including the fat concentration is negligible. Therefore, the milk refractive index used for all samples was calculated without contribution of milk fat.

134 Scattering spectra ( $\mu_s$ ,  $\mu'_s$  and  $g$ ) were smoothed by a third order Savitsky-Golay filter with a  
 135 window width of 60 nm. Since the experimental BOP contained scattering contributions of both  
 136 milk fat and casein micelles, the casein scattering was removed by correcting the experimental  
 137 spectra according to Eq. (1), based on simulated casein spectra (see section 2.3). Eq. (1)b was  
 138 obtained by substitution using the definition of  $g$  ( $g = 1 - \mu'_s/\mu_s$ ) and the fact that Eq. (1)a is  
 139 also valid for  $\mu'_s$ .

$$\mu_{s,meas,fat} = \mu_{s,meas} - VF_{casein} \times \mu_{s,sim,casein} \quad (1a)$$

$$g_{meas,fat} = [g_{meas} \times \mu_{s,meas} - VF_{casein} \times \mu_{s,sim,casein} \times g_{sim,casein}] / \mu_{s,meas,fat} \quad (1b)$$

### 141 2.3 Forward simulation of BOP

142 The BOP ( $\mu_s$ ,  $\mu'_s$ ,  $g$ ) of all measured milk fat PSDs and the casein PSD were simulated using  
 143 the iterative tool for polydisperse systems developed by Aernouts et al. (2014). All spectra were  
 144 calculated in the 0.55-1.35  $\mu\text{m}$  wavelength range with a step of 0.01  $\mu\text{m}$ . A particle volume fraction  
 145 (VF) of 1% was adopted, as this is in agreement with the requirements for independent scattering.  
 146 In this regime, the anisotropy factor is independent of the VF, while the scattering coefficient  
 147 and reduced scattering coefficient are linearly proportional to the VF. Therefore, multiplying the  
 148 simulated  $\mu_s$  and  $\mu'_s$  spectra with the desired concentration performs a correct scaling.

149 The scattering by casein micelles was simulated based on the measured reference PSD of  
 150 casein. The casein content was fixed at 75% w/w of the measured crude protein fraction, in  
 151 accordance with the average casein fraction reported for crude protein in milk (Walstra et al.,  
 152 1999). The obtained value was divided by three to match the three-fold dilution of the samples in  
 153 the optical measurements. The bulk scattering spectra of the raw and homogenized milk samples  
 154 were simulated using the reference PSD of the fat globules. The real part of the milk serum  
 155 refractive index was obtained by adding a baseline to the refractive index of water (Segelstein,  
 156 1981), following the formula given by Walstra et al. (1999) and using the three-fold diluted average  
 157 concentration of the milk components. The real and complex parts of the refractive indices of milk  
 158 fat and casein were calculated based on the absorption coefficients of a mixture of milk components  
 159 as described by Aernouts et al. (2015b).

### 160 2.4 Estimation of PSD and VF

161 Milk fat PSDs were estimated from the  $\mu_s$ ,  $\mu'_s$  and  $g$  spectra with a procedure similar to those  
 162 described by Postelmans, Aernouts, and Saeys (2018) and Postelmans et al. (2019). The PSDs were  
 163 approximated by a lognormal probability density function since this type performed best when  
 164 fitting lognormal, normal and weibull distributions directly to the measured milk fat PSDs (results  
 165 not shown). These directly fitted lognormal distributions were also used to set the parameter  
 166 boundaries of the constrained optimization in the PSD estimation routine. Lower boundaries were  
 167 set at 70% of the minimal fitted distribution parameter values, while upper boundaries were set  
 168 at 130% of the maximal values. This resulted in the range -2.70 to 0.86 for lognormal distribution  
 169 parameter  $\mu$  and 0.29 to 1.89 for parameter  $\sigma$ . VF was limited by a minimum of zero fat content  
 170 and a maximum of 3% v/v fat in three-fold diluted milk, given that the fat content in undiluted  
 171 milk will rarely exceed 5.5% w/w (6.21% v/v) (Walstra et al., 1999). During optimization, all  
 172 parameters were scaled to the range of 0.5-1.5 to reduce possible effects of differences in magnitude.

173 The PSD estimation routine consists of three steps (Postelmans et al., 2019): (1) defining  
 174 the starting points of the optimization, (2) the optimization step in which the calculated scat-  
 175 tering spectrum of the PSD estimate is iteratively updated to match the experimental scattering  
 176 spectrum, and (3) a selection procedure on the solutions to retrieve the final PSD estimate.

177 (1) In order to determine a set of limited but relevant starting points, the cost based on nor-  
 178 malized spectra (in Eq. (2) shown for  $\mu_s$ ) was evaluated for a grid of 75x75 points, equidistantly  
 179 distributed in the scaled distribution parameter space. Local minima were detected using Mat-  
 180 lab's 'imregionalmin' function (Image Processing Toolbox, Matlab R2016b, The Mathworks Inc.,

181 Massachusetts, USA). In case of  $\mu_s$  and  $\mu'_s$ , the ratio of the mean input spectrum and the mean  
 182 of the spectrum calculated for the local minima is a rough estimate for VF. This value had to  
 183 be within the VF boundaries, otherwise this combination of PSD parameters was discarded as  
 184 starting point. For  $g$ , the ratio of spectral means had to be in the 0.75-1.25 range, since a maximal  
 185 multiplicative baseline of  $\pm 25\%$  was tolerated. If more than ten local minima remained, only the  
 186 ten with the lowest cost were retained as starting points. In case of estimation on  $\mu_s$  and  $\mu'_s$ , an  
 187 initial VF value was provided by the above-mentioned ratio of spectral means.

$$\min \log_{10} \left[ \sum_{i=1}^{N_\lambda} \left( \frac{\frac{\mu_{s,i}}{\text{mean}(\mu_s)} - \frac{\widehat{\mu_{s,i}}(\text{par1}, \text{par2})}{\text{mean}(\widehat{\mu_s})}}{\frac{\mu_{s,i}}{\text{mean}(\mu_s)}} \right)^2 \right] \quad (2)$$

188 (2) The PSD estimation routine used the ‘patternsearch’ algorithm, a non-gradient based  
 189 optimizer, as implemented in the Global Optimization Toolbox of Matlab R2016b (The Mathworks  
 190 Inc., Massachusetts, USA; Conn, Gould, & Toint, 1997). In case of  $\mu_s$  or  $\mu'_s$ , the PSD and VF  
 191 were estimated simultaneously using the cost in Eq. (3). For PSD estimates based on  $g$ , only PSD  
 192 parameters  $par1$  and  $par2$  remain in Eq. (3), since  $g$  is concentration independent in the assumed  
 193 independent scattering regime.

$$\min \log_{10} \left[ \sum_{i=1}^{N_\lambda} \left( \frac{\mu_{s,i} - \widehat{\mu_{s,i}}(\text{par1}, \text{par2}, VF)}{\mu_{s,i}} \right)^2 \right] \quad (3)$$

194 (3) A selection procedure was applied to the solutions found for the different starting points.  
 195 First, all non-converged end points were discarded, as well as solutions that reached one or more  
 196 of the parameter boundaries. In case of  $g$ , the ratio of the mean input spectrum and the mean  
 197 calculated spectrum had to be between 0.75-1.25, since a multiplicative baseline error of maximal  
 198 25% on the measurements was tolerated. From the remaining end points, only those with a cost  
 199 value within 2.5% of the lowest remaining cost were considered. If they formed one group, i.e.  
 200 absolute difference between scaled distribution not more than 0.05, the solution with the minimal  
 201 cost was considered as final PSD estimate. If not, the estimated PSDs were considered non-unique  
 202 and no final estimate was selected.

## 203 2.5 Estimation of PSD and VF on $\mu_s$ and $g$ simultaneously

204 PSD and VF estimates were also made using  $\mu_s$  and  $g$  spectra simultaneously. The distribution  
 205 parameter boundaries and the steps of the optimization routine remained identical to those de-  
 206 scribed in section 2.4. Only the cost functions were replaced to include both  $\mu_s$  and  $g$ , respectively  
 207 Eq. (2) by Eq. (4) for the grid calculation based on normalized spectra, and Eq. (3) by Eq. (5)  
 208 in the optimization. In order to attribute an equal weight to  $\mu_s$  and  $g$ , a cost function based on the  
 209 ratio of the sum of squared errors (SSE) to the total squared errors (SST) was preferred over one  
 210 based on the relative difference between the spectra.

$$\begin{aligned} \min & \left[ \frac{SSE_{\mu_s, \text{norm.}}}{SST_{\mu_s, \text{norm.}}} + \frac{SSE_{g, \text{norm.}}}{SST_{g, \text{norm.}}} \right] \\ & = \min \left[ \frac{\sum_{i=1}^{N_\lambda} \left( \frac{\mu_{s,i}}{\text{mean}(\mu_s)} - \frac{\widehat{\mu_{s,i}}(\text{par1}, \text{par2})}{\text{mean}(\widehat{\mu_s})} \right)^2}{\sum_{i=1}^{N_\lambda} \left( \frac{\mu_{s,i}}{\text{mean}(\mu_s)} - \text{mean}\left(\frac{\mu_s}{\text{mean}(\mu_s)}\right) \right)^2} + \frac{\sum_{i=1}^{N_\lambda} \left( \frac{g_i}{\text{mean}(g)} - \frac{\widehat{g}_i(\text{par1}, \text{par2})}{\text{mean}(\widehat{g})} \right)^2}{\sum_{i=1}^{N_\lambda} \left( \frac{g_i}{\text{mean}(g)} - \text{mean}\left(\frac{g}{\text{mean}(g)}\right) \right)^2} \right] \quad (4) \end{aligned}$$

$$\min \left[ \frac{SSE_{\mu_s}}{SST_{\mu_s}} + \frac{SSE_g}{SST_g} \right] = \min \left[ \frac{\sum_{i=1}^{N_\lambda} (\mu_{s,i} - \widehat{\mu_{s,i}}(\text{par1}, \text{par2}, VF))^2}{\sum_{i=1}^{N_\lambda} (\mu_{s,i} - \text{mean}(\mu_s))^2} + \frac{\sum_{i=1}^{N_\lambda} (g_i - \widehat{g}_i(\text{par1}, \text{par2}))^2}{\sum_{i=1}^{N_\lambda} (g_i - \text{mean}(g))^2} \right] \quad (5)$$

211 **2.6 PSD estimation with fixed VF**

212 All PSD estimations were repeated with the VF value fixed at the three-fold dilution of the fat  
 213 content determined in the reference analysis. This way, a degree of freedom is eliminated from the  
 214 optimization routine, leaving only the distribution parameters to be determined. For estimations  
 215 based on  $\mu_s$  and  $\mu'_s$ , the cost function used for both the starting point determination and for the  
 216 optimization is based on non-normalized spectra, cf. Eq. (3) and Eq. (5), although VF is a fixed  
 217 value instead of a parameter to be estimated.

218 **3 Results and discussion**

219 **3.1 Reference PSDs**

220 The PSD of casein was identified as the first mode of the bimodal PSD of raw milk and is shown  
 221 in Fig. 1a. The casein micelles are clearly smaller than the fat globules in raw milk, but the higher  
 222 the degree of homogenization, the more their PSDs overlap. The measured peak for casein shows  
 223 a high similarity with the casein PSD reported by Aernouts et al. (2015b). However, both are  
 224 clearly underestimating the average size of casein micelles that lies around 150-200 nm according  
 225 to literature (Walstra et al., 1999; C. de Kruif, 1998; C. G. K. de Kruif & Huppertz, 2012). Stocker  
 226 et al. (2017) measured a mean diameter of 189 nm for casein in commercial skim milk, but it is  
 227 unclear whether this PSD obtained by dynamic light scattering refers to an intensity or volume  
 228 based PSD. By fitting  $\mu_s$  and  $\mu'_s$  spectra based on Mie theory, they also estimated a casein PSD  
 229 with a mean diameter of 211 nm. The discrepancy between our results and those reported by other  
 230 researchers was most likely caused by the restriction of laser diffraction measurements to assume  
 231 only one type of particles. More specific, only one particle refractive index could be defined, even  
 232 if the sample is known to contain multiple particle types, as was the case here. Since milk fat  
 233 globules were the scatterers of interest, the refractive index of milk fat was used. Because of this,  
 234 the size of the casein micelles was calculated with a too low refractive index, resulting in a casein  
 235 peak shifted to smaller sizes.

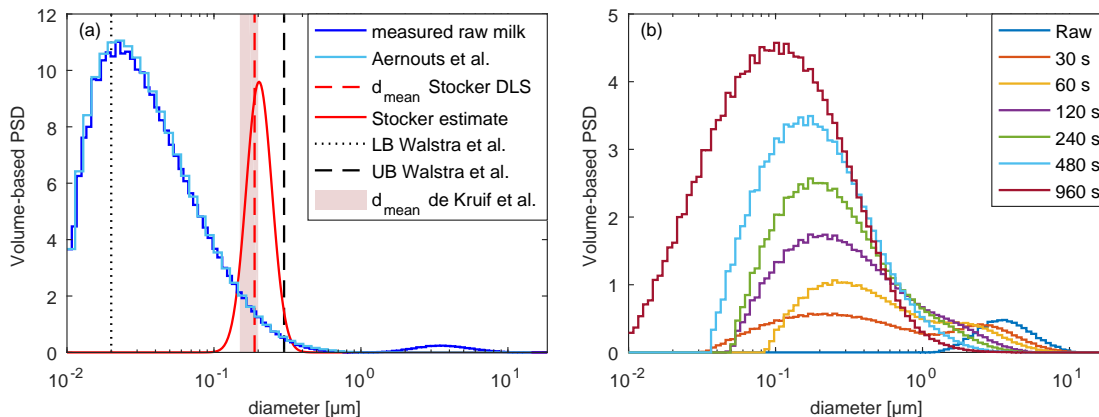


Figure 1: a) Measured PSD of raw milk compared to the casein PSD of Aernouts et al. (2015b), Stocker et al. (2017), Walstra et al. (1999), C. de Kruif (1998) and C. G. K. de Kruif and Huppertz (2012). b) Measured PSDs of milk fat (contribution of casein removed). LB = lower boundary, UB = upper boundary.

236 The measured PSDs for the milk fat globules, the scatterers of interest, are shown in Fig. 1b.  
 237 They were obtained by subtracting the casein PSD in Fig. 1a, from all sample PSDs before  
 238 converting them to probability density functions. The fat globule size distribution in raw milk  
 239 is monomodal with the peak around 3.5  $\mu\text{m}$  diameter. During homogenization, a second peak of  
 240 smaller particles around 0.2  $\mu\text{m}$  in diameter appears. The longer the homogenization time, the

241 more the relative importance of this new peak increases, rather than shifting the original peak to  
 242 smaller particle sizes.

243 The PSD of fat globules in the unhomogenized raw milk sample was found to be monomodal,  
 244 just as the example shown in the work of Cabassi et al. (2013). On the other hand, Jhanwar  
 245 and Ward (2014) obtained a bimodal PSD for whole milk, even if casein was removed before the  
 246 PSD measurement. Stocker et al. (2017) also noticed this in their PSD measurements, but they  
 247 concluded that it was an artifact and assumed a monomodal PSD for the raw fat globules.

### 248 3.2 Measured & simulated BOP

249 The mean bulk scattering spectra calculated from DIS and UT measurements (after casein cor-  
 250 rection) are shown in Fig. 2. The  $\mu_s$  spectra were well reproducible with a low noise level, while  $g$   
 251 and  $\mu'_s$  were more susceptible to noise. This is possibly due to low total reflectance values in the  
 252 DIS measurements and therefore a low signal-to-noise ratio. Moreover, such errors on the total re-  
 253 flection spectra could also be the cause of the baseline mismatch between measured and simulated  
 254  $g$  and  $\mu'_s$  spectra in case of the ‘raw’, ‘30 s’ and ‘60 s’ samples. A slight underestimation in  $\mu'_s$ , in  
 255 combination with a correct  $\mu_s$  (less depending on DIS measurements), leads to an overestimated  
 256  $g$  through the relation  $g = 1 - \mu'_s/\mu_s$ . Inaccuracies on the sample or particle refractive index can  
 257 also contribute to such baseline effects (Postelmans et al., 2018), as an approximate formula for  
 258 the refractive index of milk and milk serum was used.

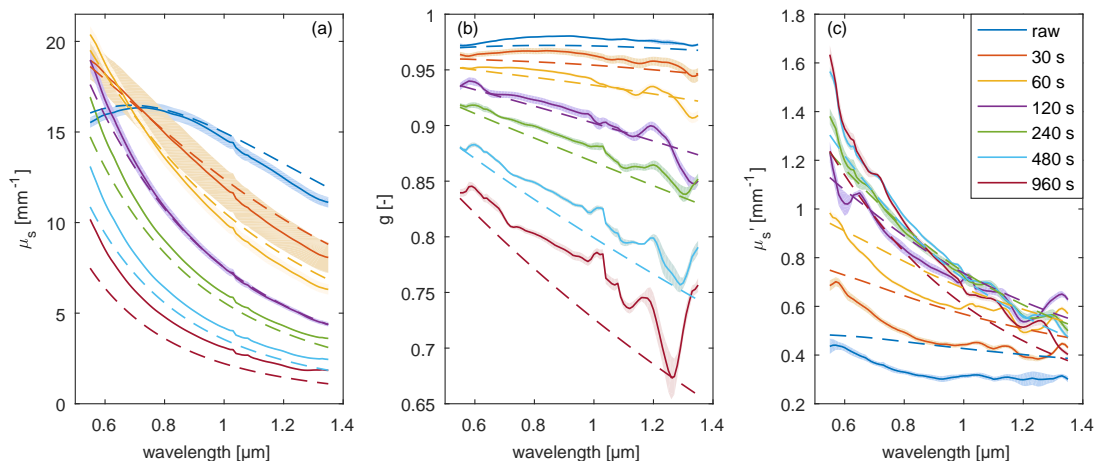


Figure 2: Mean experimental spectra with casein correction  $\pm$  standard deviation for a)  $\mu_s$ , b)  $g$  and c)  $\mu'_s$ . Dashed lines indicate the corresponding spectra simulated based on the reference PSDs.

259 In Fig. 2, the increasing underestimation of simulated  $\mu_s$  spectra compared to the experimental  
 260 casein corrected spectra suggests there might be a systematic effect on top of the measurement  
 261 errors. For samples with homogenization times ranging from 240 s to 960 s, the baseline mismatch  
 262 clearly aggravates in both  $\mu_s$  and  $g$ . It is therefore thought to be related to the homogenization  
 263 process itself. In raw milk, the fat globule membrane mainly consists of phospholipids and proteins,  
 264 and has an average thickness around 15 nm (Walstra et al., 1999). Upon homogenization, the  
 265 total globule surface area increases due to the newly formed small fat globules. To cover the  
 266 increased fat-milk plasma interface, the original membrane material is completed with adhering  
 267 casein and serum proteins (Strawbridge, Ray, Hallett, Tosh, & Dagleish, 1995). Walstra et al.  
 268 (1999) reported an average protein load per surface area of 10 mg/m<sup>2</sup>. Casein (sub)micelles are  
 269 preferentially adsorbed over serum proteins and make up about 93% of the proteins in the new  
 270 surface layer, with a preference for the largest micelles (Walstra et al., 1999). Therefore, the  
 271 concentration of free casein micelles also decreases.

272 The discrepancy between simulated and experimental scattering spectra, especially for the  
 273 milk samples that were homogenized for a longer time, suggests that the full complexity of the  
 274 homogenization process is not captured in the simulations. Both the casein coating and the  
 275 decrease in free casein micelles would result in an increase of the scattering coefficient and the  
 276 anisotropy factor, bringing the simulated and experimental profiles closer together (results not  
 277 shown). However, as information on the exact refractive indices of casein and milk fat, the fat  
 278 globule membrane thickness and the volume fraction of free casein is not available, it was not  
 279 possible to take this full complexity into account.

### 280 3.3 PSD estimation on $\mu_s$

281 The PSDs estimated on simulated and experimental  $\mu_s$  spectra are plotted together in Fig. 3 (&  
 282 Fig. S-1). The estimations on simulated spectra provide the ‘ideal’ case, since the simulations were  
 283 noise-free and the same refractive indices were used for the forward simulation and the inverse  
 284 estimation. In all cases shown in red, the VF was estimated simultaneously with the distribution  
 285 parameters. The estimated VF values are discussed separately in section 3.7.

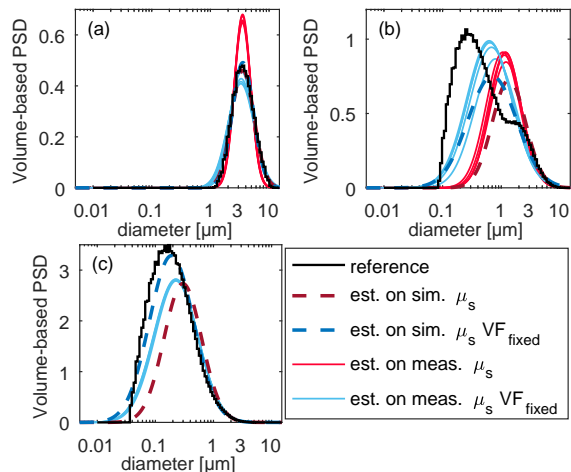


Figure 3: PSDs of milk fat globules estimated on simulated and experimental  $\mu_s$  spectra, with estimated or fixed VF. a) ‘raw’, b) ‘60 s’, c) ‘480 s’.

286 The PSD of the ‘raw’ milk sample (Fig. 3a) was well estimated based on  $\mu_s$  spectra, as the  
 287 lognormal distribution properly fits the shape of the measured PSD. A lognormal PSD was also  
 288 used by Stocker et al. (2017), while Cabassi et al. (2013) preferred a Weibull distribution. For  
 289 homogenized samples, the estimated PSD is mainly dominated by the larger fat globules (Fig. 3b)  
 290 since their scattering is more pronounced than this of the smaller particles. A lower sensitivity  
 291 to submicron particles due to the similar shape of their  $\mu_s$  spectra was already observed when  
 292 estimating PSDs (Postelmans et al., 2019). For the most intensively homogenized samples (‘480 s’  
 293 and ‘960 s’), even no valid estimates were retained by the selection procedure. Stocker et al. (2017)  
 294 reported similar difficulties: a small difference in mean particle size can have a large effect on  $\mu_s$ ,  
 295 but the estimation algorithm may attribute it to a change in particle concentration.

296 If, however, the VF was incorporated as a fixed value (1.56% v/v, three-fold dilution of the  
 297 reference analysis) instead of being estimated, it resulted in valid PSD estimates for all samples.  
 298 Moreover, the underestimation of the small globule fraction reduced (Fig. 3 blue lines), because  
 299 errors in scattering level due to an incorrect particle size could no longer be compensated for  
 300 by adapting the particle concentration. Cabassi et al. (2013) also used a fixed fat concentration  
 301 when estimating PSDs, although their estimation routine included an additional correction factor  
 302 besides the two distribution parameters. Stocker et al. (2017) estimated the concentration of  
 303 milk fat simultaneously with the PSD parameters, but no reference analysis was done to confirm



304 the accuracy of their VF estimates. In practical applications, the VF of fat globules could be  
 305 determined based on absorption spectroscopy or the bulk absorption coefficient  $\mu_a$ .

### 306 3.4 PSD estimation on $g$

307 The PSDs estimated on  $g$  spectra are shown in Fig. 4 (& Fig. S-2). In case of the ‘raw’ sample,  
 308 no valid estimate was obtained based on experimental spectra as all replicates reached the lower  
 309 boundary for  $\sigma$ . The PSDs of homogenized samples are more accurate: the higher the degree of  
 310 homogenization, the more the distribution peak shifts to smaller sizes. Furthermore, the distri-  
 311 butions are wider and approximate the left distribution tail more accurately. Nevertheless, PSDs  
 312 estimated on experimental spectra tend to be less wide than their respective counterpart estimated  
 313 on simulations, although the difference decreased with increasing degree of homogenization.

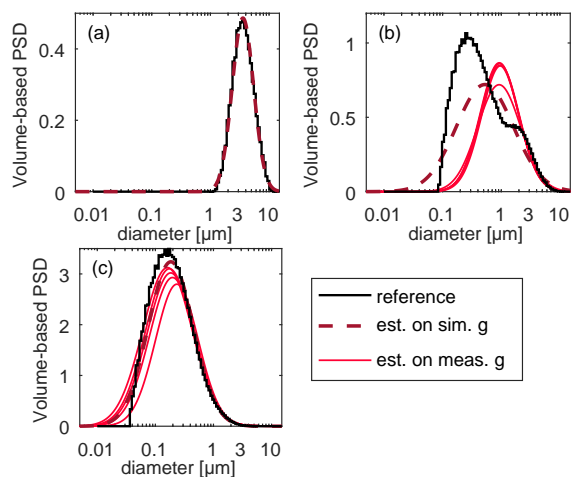


Figure 4: PSDs of milk fat globules estimated on simulated and experimental  $g$ . a) ‘raw’, b) ‘60 s’, c) ‘480 s’.

### 314 3.5 PSD estimation on $\mu'_s$

315 In general, there was less variation in the level of  $\mu'_s$  spectra compared to  $\mu_s$  and the spectra  
 316 showed similar noise as  $g$ . This negatively affected the PSD estimation, since only five valid PSD  
 317 estimates were obtained when estimated simultaneously with the VF (Fig. 5 red lines, Fig. S-3).  
 318 Similarly, problems with the distribution width estimated based on  $\mu'_s$  were already reported for  
 319 polystyrene particle suspensions (Postelmans et al., 2018, 2019).

320 Fixing VF drastically increased the number and quality of the valid PSD estimates, namely all  
 321 samples ranging from 60 s to 960 s homogenization (Fig. 5b-c, blue lines). Stocker et al. (2017) also  
 322 fixed the VF of fat while estimating a bimodal lognormal distribution on  $\mu'_s$ . However, the effect  
 323 of incorporating a known VF cannot be investigated based on their data since no PSD estimates  
 324 with VF estimation were reported by them, nor any reference PSD measurements.

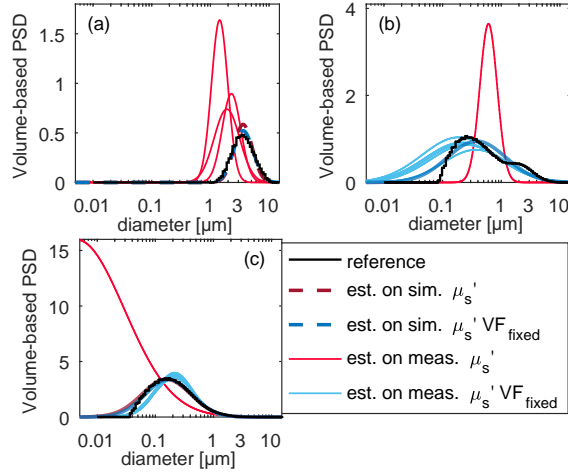


Figure 5: PSDs of milk fat globules estimated on simulated and experimental  $\mu'_s$  spectra, with estimated or fixed VF. a) ‘raw’, b) ‘60 s’, c) ‘480 s’.

### 3.6 PSD estimation more robust by combining $\mu_s$ and $g$ in cost

325

326 When looking at the PSDs estimated on  $\mu_s$ , the input spectrum is fitted relatively accurate. How-  
 327 ever, the  $g$  spectrum of these estimated PSDs often does not match well with the corresponding  
 328  $g$  spectrum. On the contrary, the PSDs estimated on  $g$ , have most often a well-matching normal-  
 329 ized  $\mu_s$  spectrum, but no VF estimation could be made. Therefore, it was investigated if PSDs  
 330 estimated on a combination of  $\mu_s$  and  $g$  would inherit ‘the best of both’: the PSD estimates on  $g$   
 331 with a VF estimate on  $\mu_s$ .

332 Using both scattering spectra as input produces a valid PSD and VF estimate for all samples  
 333 except raw milk, as can be seen in Fig. 6 (red lines, Fig. S-4). The estimates for raw milk stranded  
 334 at the lower boundary of distribution parameter  $\sigma$  in an attempt to fit both the level and shape  
 335 of the experimental spectra, just like the estimates on solely  $g$ . The estimated PSDs are generally  
 336 wider than those estimated on  $\mu_s$  and resemble more those estimated from the  $g$  spectra. The  
 337 effect of using a fixed VF in estimations on a combination of  $\mu_s$  and  $g$  is rather small, as shown by  
 338 the blue lines in Fig. 6. Combining  $\mu_s$  and  $g$  provided the highest number of valid PSD estimates,  
 339 with the smallest difference in estimated distribution parameters with or without a fixed VF value.

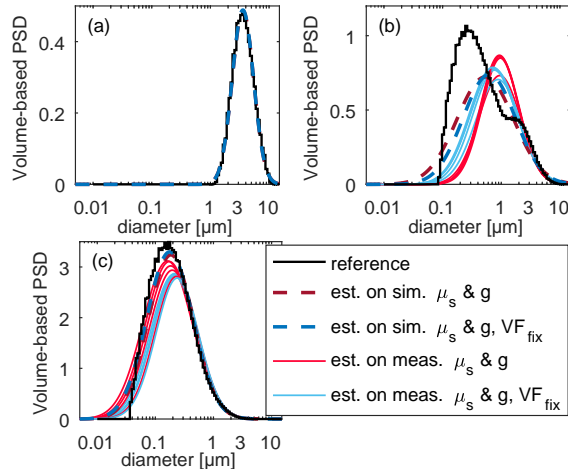


Figure 6: PSDs of milk fat globules estimated on a combination of  $\mu_s$  and  $g$  (simulated or experimental), with estimated or fixed VF. a) ‘raw’, b) ‘60 s’, c) ‘480 s’.

Therefore, if the concentration of scatterers is known, it is beneficial for the PSD estimation to fix the VF value and only estimate the distribution parameters. If VF is unknown, a simultaneous estimation of the PSD parameters and VF based on  $\mu_s$  and  $g$  spectra provides a good alternative.

### 3.7 Estimated milk fat volume fractions

Figure 7 provides an overview of all estimated VF values. Firstly, there is a clear trend in VF values estimated on simulated  $\mu_s$  spectra (red dots in Fig. 7a). As the PSD estimate for the raw milk sample was accurate, the accompanying VF is also relatively correct. Since the PSD estimation routine focuses on the large fat globules for bimodal and asymmetric PSDs, the VF is underestimated to compensate for the higher scattering level of larger particles. For higher degrees of homogenization, PSDs become again more monomodal and the estimated PSDs are more accurate, while the estimation of VF stabilizes or even improves. The same general trend is present in case of experimental  $\mu_s$  spectra (blue dots in Fig. 7a). Fig. 7b only presents a limited set of VF estimates, of which one is even close to being discarded, because the majority of the PSD estimates on  $\mu'_s$  were marked as invalid.

The relatively good PSD estimates based on  $\mu_s$  and  $g$  simultaneously were accompanied by relatively accurate VF estimates. Fig. 7c shows that the estimated VF values for the ‘30 s’, ‘60 s’ and ‘120 s’ samples based on experimental spectra are rather constant, with a small underestimation compared to the reference. The most intensively homogenized samples (‘480 s’ and ‘960 s’) on the other hand, have overestimated VF values, most likely caused by underestimation of the left distribution tail without an accompanying overestimation of the fraction large particles.

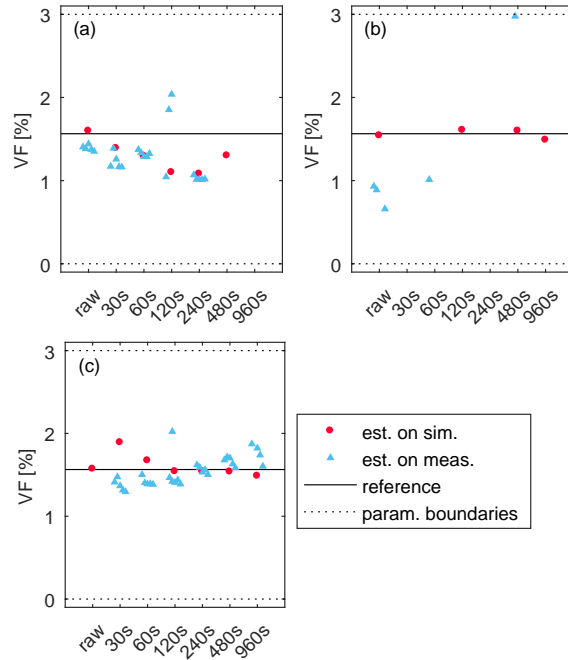


Figure 7: Volume fractions (VF) of milk fat estimated based on a)  $\mu_s$ , b)  $\mu'_s$ , and c) of  $\mu_s$  and  $g$  simultaneously.

### 3.8 General discussion

The presented PSD estimates were based on  $\mu_s$ ,  $g$  and  $\mu'_s$  spectra, with the most accurate results if the concentration of the fat globules was fixed, or when  $\mu_s$  and  $g$  were used simultaneously as input. From a practical point of view,  $\mu'_s$  would be most easy to implement as it requires solely DIS measurements or spatially resolved reflectance spectroscopy. Watté, Aernouts, Van Beers,

365 Postelmans, and Saeys (2016) et al. already reported the *in silico* optimization of a spatially  
366 resolved reflectance sensor design for determining the BOP of milk. However,  $\mu'_s$  was more noisy  
367 than  $\mu_s$  in this dataset and already resulted in larger errors in the estimated distribution width  
368 (Postelmans et al., 2019). Nevertheless, relatively accurate PSD estimates were obtained for  
369 homogenized samples if VF information was included. Stocker et al. (2017) used  $\mu'_s$  spectra to  
370 estimate bimodal PSDs of milk, but no reference measurements of the samples' PSDs are available  
371 to check the quality of these estimates. The PSD estimates for the 'raw' samples based on  $\mu_s$   
372 can be compared to those of Cabassi et al. (2013), who used corrected NIR absorption spectra  
373 to obtain PSD information. The D32 and span (D80-D20) of PSD estimates presented here were  
374 less correct if VF was estimated (RMSE of respectively 0.195 and 1.013 for our study compared  
375 to 0.110 and 0.47 for Cabassi et al.). However, the accuracy improved to a similar level if a fixed  
376 VF was used (RMSE<sub>D32</sub> of 0.131, RMSE<sub>D80-D20</sub> of 0.472).

377 The consistent underestimation of the fraction small fat globules, observed for all BOP types,  
378 is caused by the relatively limited scattering by submicron particles compared to the scattering  
379 by larger particles. On top of that, it might also be related to the baseline mismatch between  
380 experimental and simulated spectra (Fig. 2). Identifying and reducing the cause of the mismatch  
381 might help to decrease this phenomenon. In case of homogenized milk, this may imply an adap-  
382 tive casein correction that takes into account the full complexity of the homogenization process  
383 and the effects of a, most likely, increased particle refractive index on the scattering properties.  
384 Enhancing the sensitivity to the submicron size globules can also be obtained by including shorter  
385 wavelengths in the spectra, since the smaller the particle, the more its scattering peak shifts to  
386 smaller wavelengths. Michels, Foschum, and Kienle (2008) studied the BOP of different types of  
387 soy bean oil emulsions, and the start of the scattering peak in  $\mu_s$  appeared around 0.4  $\mu\text{m}$  wave-  
388 length for the Lipovenoes 10% and the ClinOleic 20% samples (largest particle diameter around  
389 0.55  $\mu\text{m}$ ).

390 Despite the issue of underestimating the left tail of the measured PSDs, the right tail is  
391 fitted quite well, even for the bimodal and asymmetric PSDs of homogenized samples. Therefore,  
392 it contains valuable information on the largest particles in the samples, even if the rest of the  
393 distribution is not estimated perfectly. Fig. S-5 provides an overview of the D90 and D32 of  
394 the estimated and reference PSDs. It shows that the estimated values for samples homogenized  
395 for 120 s or longer are consistent between the replicates and close to the reference values if a  
396 fixed VF was used. Furthermore, Table S-1 includes the mean difference between measured and  
397 reference values (D50, D90, D32, D43) listed per sample. The standard deviation on these values  
398 for milk samples homogenized for 120 s or longer are small, indicating a good reproducibility.  
399 Di Marzo et al. (2016) investigated if these PSD parameters could be predicted accurately to  
400 monitor the performance of the homogenizer inside a MIR analyser (Milkocan) and alert if it is  
401 not working properly (D90 > 1.7  $\mu\text{m}$ ) (Smith, Barbano, Lynch, & Fleming, 1995). Imposing an  
402 upper limit on particle size rather than specifications on the complete PSD would be applicable  
403 in milk homogenization since large fat globules have the largest impact on creaming properties.  
404 As the presented PSD estimation routine provides good results for the estimation of D90, it has  
405 large potential to be used for such purposes.

406 Although the monomodal distributions were fitted quite well and relatively accurate D90 and  
407 D32 values were obtained, there is a clear issue of bimodality. For the bimodal and asymmetric  
408 homogenized PSDs, e.g. samples '30 s' and '60 s', a combination of two lognormal distributions  
409 would be more suitable than a single one. This 'bimodal lognormal distribution' was already used  
410 by Stocker et al. (2017) to estimate the PSD of homogenized milk samples based on  $\mu_s$  and  $\mu'_s$ .  
411 Since the use of  $\mu_s$  and  $g$  simultaneously as input for PSD estimation gave the most promising  
412 results, a next step could be to estimate such a bimodal PSD based on these two scattering spectra.

## 413 4 Conclusion

414 The potential of estimating milk fat PSDs based on wavelength dependent light scattering prop-  
415 erties for monitoring the homogenization process of milk was investigated. Therefore, the bulk

416 optical properties of raw milk samples with an increasing degree of ultrasonic homogenization were  
417 experimentally determined by means of double integrating sphere and unscattered transmission  
418 measurements. The bulk scattering spectra of  $\mu_s$ ,  $g$ ,  $\mu'_s$  or  $\mu_s$  and  $g$  simultaneously were used as  
419 input for the estimation of lognormal PSD parameters and the volume fraction of milk fat globules.  
420 Estimated PSDs were compared to reference PSDs obtained with laser diffraction.

421 If the volume fraction and PSD parameters were estimated simultaneously, PSD estimates  
422 based on measured  $\mu_s$  spectra focused on the largest particles, especially in the mildly homogenized  
423 samples (bimodal distributions). For strongly homogenized samples, no valid estimates could be  
424 made due to the inability of the algorithm to distinguish between a small change in particle size  
425 and a change in VF. Furthermore, measured  $\mu'_s$  spectra produced practically no valid estimates  
426 since the optimizer stranded on one of the distribution parameter boundaries. Overall, estimates  
427 based on a combination of  $\mu_s$  and  $g$  proved to be most robust as valid estimates were produced  
428 for all samples except raw milk.

429 A second set of PSDs was estimated with the VF fixed at the reference value instead of  
430 estimating the VF. Including this information in the estimation routine drastically improved the  
431 number and accuracy of the PSD estimates, especially in case of  $\mu'_s$ . PSDs estimated on  $\mu_s$ ,  $g$   
432 or both still retained the tendency of underestimating the number of small particles, but not as  
433 severe as in case of a simultaneously estimated VF.

434 Overall, the single lognormal distribution was not able to fit bimodal PSDs and focussed on the  
435 largest fat globules. Nevertheless, the good fits for the right distribution tail provided relatively  
436 accurate information on the D90 of the samples, and on the D32 for samples homogenized for  
437 120 s or more. Therefore, the presented estimation routine could be a useful tool for monitoring  
438 specifications on the largest particles. Moreover, the good PSD estimates on a combination of  $\mu_s$   
439 and  $g$  spectra invite to estimate a weighted combination of two lognormal distributions based on  
440 these input spectra in order to improve the accuracy for bimodal PSDs.

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